

# Probability: A Graduate Course

## Misprints and Corrections

October 17, 2006

### 1. Misprints

Page	Line/Problem	Should be
vi	L25	chance exists
22	P4	Suppose also that $\{A_k\}$ are disjoint.
23	P5c	$A_n, n \geq 1$ are non-increasing
23	P5d	$A_n, n \geq 1$ are non-decreasing
23	P5d	then $P(\bigcup_{n=1}^{\infty} A_n) \leq \alpha$ .
24	L5_	<i>matching</i>
26	L7_	$\mathbb{P}(\Omega)$
29	L2–7	These lines should be deleted.
29	L8	delete “Alternatively,”
30	L1_	non-decreasing functions in $D$ .
39	L5_	trials
40	L17	$f(x) = \frac{\alpha\beta^\alpha}{x^{\alpha+1}}$
42	L7_	$X \stackrel{d}{=} U$ .
56	L6	Since $0 \leq 2X_1 - X_n \nearrow 2X_1 - X$
72	L2_	the proof of The-
114	P5	$\sum_{n=-\infty}^{\infty} P(n < X \leq n + m) = m$
114	P6	and $a > 0$ ,
115	P8, L2	$\sup_n X_n < \infty$ a.s.
116	P18	$\iff \exists Z \in L^r$ , such that $ X_n  \leq Z$ for all $n$
120	L11_	$P(X > x) \leq$
123	L1_	$B_k = \{\max_{1 \leq j \leq k}  S_j  \leq x\}$
125	L13	$\sum_{k=1}^n \text{Var } X_k \leq x^2 + \dots$ . (. instead of , at the end)
127	L5_	independent summands there exists
128	L7	$E X ^r \leq c_r(E X + y ^r +  y ^r) < \infty$ .
129	L10	For $0 < p \leq r$
129	L11	$\ X\ _p \leq \ X\ _r$ .
129	L13	(and $q$ by $(1 - p/r)^{-1}$ )
133	L18,19	The median is not unique in general, and typically
133	(ii)	for some $r > 0$ ,
133	(iii)	$ \text{med}(X) - EX  \leq 2^{1/r} \ X - EX\ _r$ .
139	L12_	$= \frac{1}{2} P(\max_{1 \leq k \leq n} S_k - \text{med}(S_n - S_k) > x)$ ,
142	L10_	$ S_n - S_k  \geq \dots$

Page	Line/Problem	Should be
150	L7 <sub>-</sub> , 1 <sub>-</sub>	Khintchine
155	P9b	$Q_X(ah) = aQ_X(h)$ ?
158	L2	$\varphi_X(t) = E e^{itX} = \int_{-\infty}^{\infty} e^{itx} dF_X(x)$
158	L13	$ E e^{itX} $
159	L7	$\frac{pe^{it}}{1-qe^{it}}$
160	L4 <sub>-</sub> , 2 <sub>-</sub>	$H(a, b, x, T)$
160	L3 <sub>-</sub>	$H(a, b, X, T)$
161	L4	$H(a, b, x, T)$
161	L6	$H(a, b, X, T)$
164	L8	$E \exp\{it(X - a)\}$
165	L8 <sub>-</sub>	$\frac{1}{2} = \dots = \frac{1}{\pi} \int_0^{\infty} \left(\frac{\sin x}{x}\right)^2 dx$
172	L14	(2.4)
184	L5	$-\frac{1}{n} \quad -\frac{1}{n} \quad 1 - \frac{1}{n} \quad \dots \quad -\frac{1}{n}$
195	L8 <sub>-</sub>	$f(x) = \frac{1}{\sigma x \sqrt{2\pi}} \exp\{\dots\}$
195	L6 <sub>-</sub>	$\int_0^{\infty} \frac{C+2 \log x + \frac{(2 \log x - \mu)^2}{2\sigma^2}}{1+x^2} dx$
196	L2,4	delete $2\pi$
196	L4,5	$\sum_{k=1}^{\infty}$
196	L5	$\geq C \sum_{k=1}^{\infty} \frac{1}{\beta+k}$
198	P12	$\frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1-\cos y}{y^2} e^{- y } dy = \frac{1}{2} - \frac{\log 2}{\pi}$
198	P14	$\int_0^{\infty} \varphi(tu) e^{-u} du$ or $\frac{1}{2} \int_{-\infty}^{\infty} \varphi(tu) e^{- u } du$
198	P14	$\frac{1}{\pi} \int_{-\infty}^{\infty} \frac{\varphi(tu)}{1+u^2} du$
216	L14 <sub>-</sub>	let $Y_1, Y_2, \dots$ be
216	L8 <sub>-</sub>	these facts
229	L 6	converges to a limit $a \in \mathbb{R}$ if and only if
229	L 7	contains a subsequence converging to $a$ .
245	L 19	Theorem 10.2.
254	L9	$n \geq m$ .
254	L13	$= \frac{1}{2} P(\bigcup_{n=m}^{\infty} A_n^{(m)})$ .
258	L6 <sub>-</sub>	$\inf\{x : \dots\}$ (twice)
258	L2 <sub>-</sub>	$\dots = F_n(x)$
261	P9, L2	converges in distribution
261	P10	$P(X_n < a \text{ i.o. and } X_n > b \text{ i.o.})$
261	P11, L1	non-degenerate random
262	P13	$\sum_{n=1}^{\infty}  a_n  < \infty$
262	P16, L3	$\varphi(0) = 1$
263	P17	$\tilde{\varphi}(t) = \exp\left\{\int_{\mathbb{R}} \varphi(t, u) dG(u) - 1\right\}$
263	P21, L2	$n \geq 2$
263	P21, L3	Independence between $X_1, X_2, \dots$ not necessary
264	P22(a)	$\iff X_n^{(k)} \xrightarrow{P} X^{(k)}$
264	P22(b)	$\implies h(\mathbf{X}_n) \xrightarrow{P} h(\mathbf{X})$
274	L17	For the converse: Assume also that $\frac{\mu_n}{b_n} \rightarrow 0$ as $n \rightarrow \infty$ .

Page	Line/Problem	Should be
279	L10 <sub>-</sub>	$= \dots \leq \frac{2}{n} + \frac{4}{n} \sum_{j=1}^n \dots$
284	L12,14	$nE XI\{ X  \leq n \log_2 n\}$
286	L2 <sub>-</sub>	right-hand side
289	L10 <sub>-</sub>	symmetric and uniformly bounded
290	L11	bounded, symmetric
292	L11 <sub>-</sub>	Proposition 5.1.2
296	L2	$\limsup_{n \rightarrow \infty} \frac{ S_n }{n}$
299	L3	$\sum_{n=1}^{\infty} \frac{Y_n - E Y_n}{n^{1/r}}$
299	L7	$\rightarrow 0$
299	L9	$\xrightarrow{a.s.} 0$
300	L1	$\leq C n^{(1/2)-(1/2r)} \cdot E X ^r$
300	L6	$\leq \sum_{k=1}^n E X_k  I\{ X_k  > k^{1/r}\}$
303	L2 <sub>-</sub>	$Y_{N(t_{k_j})} \rightarrow Y$
305	L11 <sub>-</sub>	that, for any given $j = 0, 1, \dots, k-1$ ,
305	L10 <sub>-</sub>	$\frac{1}{n} \sum_{i=1}^n I\{X_i = j\} \xrightarrow{a.s.} \frac{1}{k}$
309	L11	Theorem 5.4.1
310	L10	one can prove
324	L3	Elements $a_{12}$ and $a_{21}$ of $\Lambda$ should be $\rho\sigma_x\sigma_y$ instead of $\rho$ .
325	P9, L1	non-degenerate random
325	P11c	$\sum_{k=2}^n$
327	P22, L2	$Y_n = \max\{\dots\}$
332	L2 <sub>-</sub>	$E \min\left\{\frac{2 tX_k }{s_n}, \frac{t^2 X_k^2}{2s_n^2}\right\}$
332	L1 <sub>-</sub>	$E \min\left\{\frac{t^2 X_k^2}{s_n^2}, \frac{ t^3 X_k ^3}{6s_n^3}\right\}$
335	L8	$X_n \xrightarrow{d} Y \iff E h(X_n) \rightarrow E h(Y)$
335	L6 <sub>-</sub>	$X_j$
338	L4,7	$\frac{X_k^2 t^2}{2s_n^2}$ <span style="float: right;">(<math>X_k^2</math> instead of <math>\sigma_k^2</math>)</span>
343	L8	$P(Z_k = 0) = 1 - \frac{1}{k^2}$
344	L2 <sub>-</sub>	$S_n = \sum_{j=1}^n X_{n,j}$
345	L18	If $\sum_{j=1}^n E X_{n,j}^2 I\{ X_{n,j}  > \varepsilon\} \rightarrow 0$ as $n \rightarrow \infty$
345	L19	$S_n \xrightarrow{d} N(0, 1)$
346	L3 <sub>-</sub>	$n_0 = [t]$
347	L14	$\frac{[n_0(1+\varepsilon^3)]-n_0}{\varepsilon^2 n_0}$
349	L10 <sub>-</sub> , 7 <sub>-</sub>	$n(\dots) = \frac{1}{2}(\dots)^2$
349	L9 <sub>-</sub>	Since $x^2$ is continuous
350	L9	$S_n = \sum_{k=1}^n X_k$
357	L6 <sub>-</sub>	Tri( $-T, T$ )-distributed
357	L1 <sub>-</sub>	$\dots = \int_{-\infty}^{\infty} e^{ixt} \frac{2 \sin^2(xT/2)}{\pi T x^2} dx = \dots$
361	L7 <sub>-</sub>	$r'_k$
361	L6 <sub>-</sub>	$r'_k + r''_k$
361	L5 <sub>-</sub> , 4 <sub>-</sub>	$\theta'_k$
362	L12	$\sup_x  F_{S_n/s_n}(x) - \Phi(x)  \leq \frac{16}{\pi} \frac{\beta_n^3}{s_n^3} \int_{-s_n^3/4\beta_n^3}^{s_n^3/4\beta_n^3} t^2 e^{-t^2/3} dt + \dots$

Page	Line/Problem	Should be
376	P1,P2	$\iff \sum_{k=1}^{\infty}$
377	P7, L2	$P(X_k = \frac{1}{\sqrt{k}}) = P(X_k = -\frac{1}{\sqrt{k}}) = 1/2$
377	P7, L4	$n \geq 2$
379	P17, L9	the amount of time (delete “relative”)
379	P18, L2	$(S_n)^p$ , suitably normalized,
380	P21	$\sqrt{n}(s_n^2 - \sigma^2)$
380	P23,L7	delete “suitably normalized”.
385	L12	In other words, $\frac{S_n}{\sqrt{2\sigma^2 \log \log n}}$
396	L13	$n_k = \min\{n : s_n^2 > \lambda^k\}$ (which reduces to $[\lambda^k] + 1 \dots$
399	L8_	be a strictly
400	L9_	be a strictly
436	L 15	$E X ^r = \infty$ for $r > \alpha$ . Both possibilities exist for $r = \alpha$ .
436	L19	$a_n = \inf\{\dots\}$
439	L9	$\int_1^{\infty} \frac{dy}{y}$
443	L1	infinitely divisible distribution
443	L1_	due to (i)
450	L7,8,9,10	$\sup_{k \in \mathbb{Z}^+}$
451	L3_	iff (one iff too many)
459	L9_	$d(X, Y) \leq P(X \neq Y)$
459	L8_	Let $X_1, X_2, \dots$ and $Y_1, Y_2, \dots$
464	P1	$X$ is symmetric stable
465	P3,L1	$X$ is symmetric stable
465	P3,L3	$XY$ is symmetric stable
465	P5	$\sum_{n=1}^{\infty}  a_n ^\alpha < \infty$ (twice)
465	P7c	Incorrect. Find the correct representation. (Use (a))
470	L19	with respect to $P$ ; $Q \ll P$ .
472	L3	$= \int_{\Lambda} (\dots) dP$ .
477	L3_	$\{A_n, n \geq 0\}$
491	L2	arbitrary. (One . too much)
495	L1_	$\int_{\Lambda} X_{\tau} dP = \int_{\Lambda} Z dP$ for all $\Lambda \in \mathcal{F}_{\tau}$ .
498	L6	$X_n^{(1)}(\omega), X_n^{(2)}(\omega),$ ( , is missing)
500	L8	as long as $n < \tau$
503	L3	delete $\leq E X_n - E X_0$
503	L9	$\lambda P(\Lambda) \leq \dots = E X_n - \int_{\Lambda} X_n dP - E X_0,$
504	L5	$\int_{\max\{0 \leq k \leq n : c_k X_k^+ > \lambda\}} (\dots) dP$
504	L5	$-c_n \int_{\{\max_{0 \leq k \leq n} c_k X_k^+ > \lambda\}} X_n^+ dP$
505	L5	$= 1 + \int_{\Omega} X \left( \int_1^Y \frac{1}{y} dy \right) dP$
508	L20	almost surely
519	L4	$X_n^+ \xrightarrow{a.s.} X_{\infty}^+$
531	L24	triviality
532	L1_	$\mu\tau$ (instead of $\mu(\tau)$ )
536	L15	and (delete “that”)

Page	Line/Problem	Should be
549	P7, L4	and that $E Y ^r < \infty$
550	L8-, 7-	$n \geq 0$
551	P16	$X_n = E(A_\infty   \mathcal{F}_n) - A_n$
552	P25	$X_n = e^{S_n - n/2}$
552	P26	$W_n = \min\{Y_n, a\}$
560	L1-	$-\sum_{n=1}^{\infty}$
564	L6-	assume also that $B_n \nearrow \infty$ as $n \rightarrow \infty$
566	L7-	non-zero exponent have an ultimately monotone equivalent.
572	L7-	The line should be deleted.

## 2. Corrections

- The results in Chapter 8, except for Theorem 1.1, are for independent, *identically distributed* random variables (this is obvious(?) from the proofs).
- The reference to Kolmogorov on page 424 is not correct. It was after a lecture (one of three in a series) given in 1919 by Lévy himself that someone told him ... (The passage was somewhat hastily taken from the obituary in *Ann. Probab.* **1** (1973).)
- Problem 1, page 464. For this problem we also need to know that non-negative stable distributions with index in  $(0, 1)$  possess a *Laplace transform*. In the present case this means that “we know” that

$$L_Y(s) = E e^{-sY} = \exp\{-s^\beta\} \quad \text{for } s > 0.$$

For more about Laplace transforms, see Feller’s Volume II, Chapter XIII, and for this kind of stable distributions, Section XIII.6.

- Problem 10, page 549: The second claim should be “the expected amount of money spent when the game is over”. The third claim should be deleted. (The slip is due to a confusion with the St. Petersburg game, Subsection 6.4.1, page 283-4).