PhD course in Probability, Home Assignment 1: Probabilities and Convergence

This home assignment consists of 7 exercises. The deadline is on March 4th 2015 23:59 pm. You can hand in the assignment later, but if you hand it in k days too late, then your score will be multiplied by 1 - 0.15k.

- 1. Consider a branching process with offspring distribution X such that $\mathbb{P}(X = k) = p^k(1-p)$ for k = 0, 1, ... (this is essentially a geometric distribution) where $p \in (0, 1)$ is the parameter. Let Z_n denote the number of individuals at level/generation n. In class, we proved that $\mathbb{P}(\lim Z_n \neq 0) > 0$ iff $\mathbb{E}[X] > 1$, which in our case translates into p > 1/2. We can call the induced graph the family tree of a Galton-Watson process with geometric offspring distribution, or just a geometric tree for short. We always start the process with one individual in generation 0, i.e. $Z_0 = 1$, and we call the corresponding vertex the root.
 - (a) It can be of interest to know the value of the survival probability $\theta(p) = \mathbb{P}(\lim Z_n \neq 0)$. Determine the function $\theta(p)$ explicitly.
 - (b) A binary tree is a tree in which every individual has exactly two descendants. We want to determine the probability that, starting at the root, the geometric tree contains a binary tree. Let $\theta_2(p)$ denote this probability and prove that $\theta_2(p)$ satisfies the equation

$$(1 - \theta_2(p))(1 - p(1 - \theta_2(p)))^2 - (1 - p)(1 - p + 2p\theta_2(p)) = 0.$$
(1)

Remark: Solving (1) gives the solution

$$\theta_2(p) = \begin{cases} \frac{3p-2+\sqrt{p(5p-4)}}{2p} & \text{if } 4/5 \le p \le 1\\ 0 & \text{otherwise.} \end{cases}$$

As can be seen from a plot (see Figure 1) and indeed from the expression itself, this has a real solution only when $4/5 \le p(\le 1)$. Therefore, we conclude that the geometric tree can contain a binary tree iff $4/5 \le p \le 1$. However, much more interesting is that *at* the critical value 4/5, the existence of a binary subtree is strictly positive!

Hint: It can be useful to condition on the number of individuals in the first generation.

2. Let $f : [0,1] \to \mathbb{R}$ be a continuous function (i.e. $f \in C([0,1])$), and let $X_n \sim \operatorname{Bin}(n, x)$, i.e. such that

$$\mathbb{P}(X_n = k) = \binom{n}{k} x^k (1 - x)^{n-k} \text{ for } k = 0, \dots, n.$$

Let $A = \{|X_n/n - x| > \delta\}$, and let $Z_n = f(x) - f(X_n/n)$. Using that $\mathbb{E}[Z_n] = \mathbb{E}[Z_n I(A)] + \mathbb{E}[Z_n I(A^c)]$, prove that

$$\lim_{n \to \infty} \sup_{0 \le x \le 1} \left| f(x) - \sum_{k=0}^{n} f(k/n) \binom{n}{k} x^k (1-x)^{n-k} \right| = 0.$$

Remark: This is Weierstrass' approximation theorem.



Figur 1: A plot of $\theta_2(p)$. Note the discontinuity.

- 3. In many applications, one is interested in the maxima of i.i.d. random variables. For example, let X_k denote the (random) height of the largest wave to hit the coastal area of the Netherlands during some fixed month. Then the distribution of $M_n := \max_{k \leq n} X_k$ will give some information about how high you should build the walls protecting the people living there. This is also relevant when pricing insurances. How bad will the worst storm be that occurs over a hundred year period? Prove the following:
 - (a) Thin tail distribution: Assume that X_k have distribution function F(x) such that

$$\lim_{x \to \infty} \frac{1 - F(x)}{e^{-x}} = 1.$$

Prove that $(M_n - \log n) \xrightarrow{w} X$ where X has distribution function $F_X(x) = \exp(-e^{-x})$ for every $x \in \mathbb{R}$.

(b) **Heavy tail distribution:** Assume that X_k have distribution function F(x) such that

$$\lim_{x \to \infty} \frac{1 - F(x)}{x^{-\alpha}} = 1.$$

Prove that $M_n/n^{1/\alpha} \xrightarrow{w} X$ where X has distribution function $F_X(x) = \exp(-x^{-\alpha})$ for every x > 0.

4. Let $(X_n)_{n\geq 1}$ be a sequence of random variables such that for every $\epsilon > 0$,

$$\lim_{n \to \infty} \mathbb{P}(|X_n - X| \ge \epsilon) = 0.$$

(a) Prove that

$$\sup_{n_1,n_2 \ge N} \mathbb{P}(|X_{n_1} - X_{n_2}| \ge \epsilon) \to 0 \text{ as } N \to \infty.$$

- (b) Prove that if $(X_n)_{n\geq 1}$ are independent, then X is almost surely constant.
- 5. In many applications (e.g financial mathematics), one is interested in some function of normal random variables. In particular taking the exponential. The purpose of this exercise is to explore some convergence properties of this. Thus,

let $\alpha, \beta \in \mathbb{R}$ and Z_1, Z_2, \ldots be an i.i.d. sequence of N(0, 1) random variables. Define

$$X_n = e^{\alpha S_n - \beta n}$$

where $S_n = \sum_{i=1}^n Z_i$.

Recall that:

- $X_n \xrightarrow{P} X$ if $\forall \epsilon > 0$, $\mathbb{P}(|X_n X| \ge \epsilon) \to 0$.
- $X_n \xrightarrow{r} X$ if $\mathbb{E}[|X_n X|^r] \to 0$, (where we assume that r > 0).
- (a) Prove that $X_n \xrightarrow{P} 0$ iff $\beta > 0$. (This should be done by direct methods, i.e. it does not suffice to do (c) and then use that a.s. convergence implies convergence in probability).
- (b) Prove that $X_n \xrightarrow{r} 0$ iff $r < 2\beta/\alpha^2$.
- (c) Prove that $X_n \stackrel{a.s.}{\to} 0$ iff $\beta > 0$.
- 6. Consider C[0,1], the set of continuous functions from [0,1] to \mathbb{R} , and let C[0,1] be equipped with the standard metric

$$\rho(f,g) = \sup_{t \in [0,1]} |f(t) - g(t)|.$$

Let

$$f_n(x) = nxI(x \in [0, 1/n]) + (2 - nx)I(x \in (1/n, 2/n])$$

which is simply the function that increases linearly from 0 to 1 for $x \in [0, 1/n]$, and then decreases linearly from 1 to 0 for $x \in (1/n, 2/n]$.

Let μ_n be the measure on C[0,1] which assigns unit mass to f_n , and let μ be the measure which assigns unit mass to the function $f(x) = 0 \forall x \in [0,1]$.

Prove that $\mu_n \xrightarrow{\psi} \mu$ but that the finite dimensional marginals of μ_n converges to those of μ .

Remark: This proves that while C_f is a separating class, it is not a convergence determining class.

7. Let S be a metric space with metric d. Assume that (X_n, Y_n) is a random variable taking values in $S \times S$ (so that X_n and Y_n are random variables taking values in S).

Define the random variable $Z_n(\omega) := d(X_n(\omega), Y_n(\omega))$, which is then a random variable taking values in \mathbb{R}^+ .

Prove that if $X_n \xrightarrow{w} X$ and $Z_n \xrightarrow{w} 0$, then $Y_n \xrightarrow{w} X$.