# A PROBABILISTIC INTERPRETATION OF THE GAUSSIAN BINOMIAL COEFFICIENTS

TAKIS KONSTANTOPOULOS,\* Uppsala University

LINGLONG YUAN,\*\* Xi'an Jiaotong-Liverpool University

#### Abstract

We give a stand-alone simple proof of a probabilistic interpretation of the Gaussian binomial coefficients by conditioning a random walk to hit a given lattice point at a given time.

Keywords: Gaussian binomial coefficients, q-binomial, random walk

2010 Mathematics Subject Classification: Primary 60G50

Secondary 60G40;05E99

## Introduction

The Gaussian binomial coefficients (also known as q-binomial coefficients) [4] are generalizations of classical binomial coefficients and are usually defined as

$$\begin{bmatrix} n \\ m \end{bmatrix}_{q} = \frac{(1-q^{n})(1-q^{n-1})\cdots(1-q^{n-m+1})}{(1-q)(1-q^{2})\cdots(1-q^{m})}.$$

The term "generalization" is justified, e.g., by the fact that  $\lim_{q\to 1} {n \brack m}_q = {n \brack m}$ , which becomes obvious if we divide each term in the numerator and denominator of the last display by 1 - q and expand the ratio into a power series with finitely many terms. The Gaussian binomial coefficients turn out to be polynomial functions of the variable q and satisfy many analogs of the usual properties of binomial coefficients. We refer, e.g., to the textbook of Kac and Cheung [5].

Originally, they appeared in combinatorics, so it is not surprising that they are nowadays very important in random polymer models which have strong connections to algebraic combinatorics; see, for example, the recent work on the q-weighted version of the Robinson-Schensted algorithm introduced by O'Connell and Pei [6]. In the study of random graphs, Gaussian binomial coefficients

<sup>\*</sup> Postal address: Department of Mathematics, Uppsala University, SE-75106 Uppsala, Sweden

<sup>\*\*</sup> Postal address: Department of Mathematical Sciences, Xi'an Jiaotong-Liverpool University, 111 Ren'ai Road, Suzhou, P. R. China 215123

are present, for instance, in the distributions of the sizes of the transitive closure and transitive reduction of node 1 in a random acyclic digraph with n nodes, see [3] and [2]. Another application is in integer-valued random matrices; see, for example, [1] where the distribution of the m-rank of a random matrix is expressed in terms of these coefficients.

The purpose of this note is to give a short proof of a probabilistic interpretation of the Gaussian coefficients which, not surprisingly, is very similar to their combinatorial interpretation, given by Pólya [7], as counting the number of nondecreasing paths in a rectangle in the 2-dimensional integer lattice that leave a fixed area below them. The probabilistic proof given below (Theorem 1) is different than Pólya's. The note is stand-alone in that everything discussed is proved, including Heine's formula (see (4) below) that is needed at the end of the proof of Theorem 1. The probabilistic interpretation gives a natural meaning to several identities and properties satisfied by the coefficients (see end remarks).

### The statement and proof

Consider nondecreasing paths in the standard 2-dimensional integer lattice  $\mathbb{Z}^2$ , that is, finite or infinite sequences  $x_0, x_1, \ldots$  of elements of  $\mathbb{Z}^2$  such that  $\eta_i = x_i - x_{i-1}$  is either  $e_1$  or  $e_2$ , where  $e_1 = (1,0), e_2 = (0,1)$ , the standard unit vectors. Let r, s be nonnegative integers. By a random nondecreasing path from (0,0) to (r,s) we mean a finite nondecreasing path that starts at  $x_0 = (0,0)$ and ends at  $x_{r+s} = (r,s)$ , and that is chosen uniformly at random among the set of all such paths. Since there are  $\binom{r+s}{r}$  such paths, the increments sequence  $(\eta_1, \ldots, \eta_n)$  is assigned probability equal to  $\binom{r+s}{r}^{-1}$ .

**Theorem 1.** Consider a random nondecreasing path from (0,0) to (r,s). This path splits the rectangle  $[0,r] \times [0,s]$  into two regions. Let  $A_{r,s}$  be the area of the region under the path. Then

$$\mathbb{E}q^{A_{r,s}} = \begin{bmatrix} r+s \\ r \end{bmatrix}_q / \binom{r+s}{r}.$$

Proof. Toss a fair coin independently and let  $e_1$  represent heads and  $e_2$  tails. Denote by  $\xi_1, \xi_2, \ldots$ the successive outcomes, a random independent sequence with  $\mathbb{P}(\xi_t = e_i) = 1/2$ ,  $i = 1, 2, t \ge 1$ . Let  $X_0 = 0, X_t = \xi_1 + \cdots + \xi_t, t \ge 1$ . If it takes  $T_{r+1}$  coin tosses until the (r+1)-th head occurs for the first time then, conditional on the event that we have seen s tails up to  $T_{r+1}$ , the sequence  $(\xi_1, \ldots, \xi_{T_{r+1}-1})$ (of length  $T_{r+1} - 1 = s + (r+1) - 1 = s + r$ , under the conditioning) has uniform distribution. Thus, conditional on the same event, the path  $(X_0, X_1, \ldots, X_{T_{r+1}-1})$  is a random nondecreasing path from (0,0) to (r,s). Let  $N_i(t)$  be the number of heads/tails seen up to the t-th toss:

$$N_i(t) \coloneqq \sum_{k=1}^t \mathbf{1}\{\xi_k = e_i\}, \quad i = 1, 2,$$

and consider the stopping times

$$T_0 \coloneqq 0, \quad T_m \coloneqq \inf\{t \ge 1 : N_1(t) = m\}, \ m \ge 1.$$

Since  $T_1, T_2, \ldots$  is an increasing sequence of stopping times in i.i.d. Bernoulli trials, the random variables  $Z_i = N_2(T_{i+1}) - N_2(T_i)$ ,  $i = 0, 1, 2, \ldots$ , are i.i.d. geometric:  $\mathbb{P}(Z_i = j) = (1/2)^{j+1}$ ,  $j \ge 0$ , and so  $\mathbb{E}\theta^{Z_i} = \frac{1}{2}(1 - \theta/2)^{-1}$ ,  $|\theta| < 2$ . Simply putting it, the  $Z_i$  count the number of up-steps of the path between two successive right-steps and so

$$V = N_2(T_{r+1}) = \sum_{i=0}^r Z_i$$

is the total number of up-steps up to  $T_{r+1}$ . The distribution of V is

$$\mathbb{P}(V=s) = \left(\frac{1}{2}\right)^{r+s+1} \binom{r+s}{r}, \quad s \ge 0.$$

$$\tag{1}$$

The area  $A = A_{r,s}$  under the path  $(X_0, X_1, \dots, X_{T_{r+1}-1})$  is then

$$A = rZ_0 + (r-1)Z_1 + \dots + Z_{r-1}.$$

Letting  $q, \theta$  be variable with, say,  $|q|, |\theta| < 2$ , we have

$$\mathbb{E} q^{A} \theta^{V} = \mathbb{E} [(q^{r} \theta)^{Z_{0}} (q^{r-1} \theta)^{Z_{1}} \cdots (q\theta)^{Z_{r-1}} \theta^{Z_{r}}]$$
  
=  $(\frac{1}{2})^{r+1} \frac{1}{1 - \frac{q^{r} \theta}{2}} \frac{1}{1 - \frac{q^{r-1} \theta}{2}} \cdots \frac{1}{1 - \frac{q\theta}{2}} \frac{1}{1 - \frac{\theta}{2}} = (\frac{1}{2})^{r+1} \sum_{s=0}^{\infty} C_{r,s}(q) (\theta/2)^{s},$ (2)

where the  $C_{r,s}(q)$  are defined by the right-hand side as coefficients in the Taylor expansion in the variable  $\theta/2$ . On the other hand,

$$\mathbb{E} q^A \theta^V = \sum_{s=0}^{\infty} \theta^s \mathbb{P}(V=s) \mathbb{E}[q^A | V=s].$$
(3)

Equating coefficients in (2) and (3), also taking into account (1), gives

$$\mathbb{E}[q^A|V=s] = C_{r,s}(q) / \binom{r+s}{r}.$$

It remains to show that the  $C_{r,s}(q)$  are Gaussian binomial coefficients. To this end, we prove that if

$$F_r(x) \coloneqq \prod_{j=0}^r \frac{1}{1 - q^j x} = \sum_{s=0}^\infty C_{r,s}(q) x^s, \tag{4}$$

then the recurrence relation

$$C_{r,s}(q) = C_{r,s-1}(q) \frac{1 - q^{r+s}}{1 - q^s}, \quad s \ge 1,$$
(5)

holds. This follows quite easily from the observation that

$$(1 - q^{r+1}x)F_r(qx) = (1 - x)F_r(x)$$

Indeed, if, in this identity, we replace  $F_r(qx)$  and  $F_r(x)$  by their series, from the right-hand side of (4), and equate coefficients of similar powers, we obtain (5). Since, clearly,  $C_{r,0}(q) = F_r(0) = 1$ , we can iterate (5) to obtain

$$C_{r,s}(q) = \frac{1 - q^{r+s}}{1 - q^s} \frac{1 - q^{r+s-1}}{1 - q^{s-1}} \cdots \frac{1 - q^{r+1}}{1 - q} = \begin{bmatrix} r+s\\r \end{bmatrix}_q.$$

This completes the proof.

### Remarks

1. Since  $\begin{bmatrix} r+s \\ r \end{bmatrix}_q$  is proportional to  $\mathbb{E}q^{A_{r,s}}$  we have that  $\begin{bmatrix} r+s \\ r \end{bmatrix}_q$  is a polynomial in q.

2. Formula (4) with  $C_{r,s}(q)$  the Gaussian binomial coefficients is known as Heine's formula [5]. When q = 1 it corresponds to the Taylor series (Newton's formula)  $(1-x)^{-r} = \sum_{s\geq 0} {\binom{-r}{s}} (-x)^s = \sum_{s\geq 0} {\binom{r+s}{r}} x^s$ . 3. By symmetry, the area above the random nondecreasing path has the same distribution as the area below, i.e., the random variables  $A_{r,s}$  and  $rs - A_{r,s}$  have the same distribution. This is equivalent to the identity  $\begin{bmatrix} r+s\\r \end{bmatrix}_q = q^{rs} \begin{bmatrix} r+s\\r \end{bmatrix}_{1/q}$ .

4. By the definition of the random variable  $A_{r,s}$  as the area under a random nondecreasing path from (0,0) to (r,s) we see, by conditioning on the last edge of this path, that  $A_{r,s}$  is in distribution equal to  $A_{r,s-1}$  with probability s/(r+s) or to  $A_{r-1,s} + s$  with probability r/(r+s). Using then the result of Theorem 1, the well-known recursion  $\begin{bmatrix} r+s\\r \end{bmatrix}_q = \begin{bmatrix} r+s-1\\r \end{bmatrix}_q + q^s \begin{bmatrix} r+s-1\\r-1 \end{bmatrix}_q$  follows.

#### Acknowledgements

The authors thank the support of Swedish Research Council grant 2013-4688.

#### References

- R.P. Brent and B.D. McKay (1987). Determinants and ranks of random matrices over Z<sub>m</sub>. Discr. Math. 66, No. 1-2, 35-49.
- [2] D. Crippa and K. Simon (1995). q-distributions in random graphs: transitive closure and reduction. Tech. Report, ETH Zurich, Dept. of Computer Science.
- [3] D. Crippa and K. Simon (1997). q-distributions and Markov processes. Discr. Math. 170, No. 1, 81-98.
- [4] C.F. Gauß (1811). Summatio quarundam serierum singularium. In: Untersuchungen über höhere Arithmetik. German translation by H. Maser of Disquisitiones Arithmeticae and other papers on number theory, 2nd Ed. Chelsea, New York, 1965.
- [5] V. Kac and P. Cheung (2002). Quantum Calculus. Springer-Verlag, New York.
- [6] N. O'Connell and Y. Pei (2013). A q-weighted version of the Robinson-Schensted algorithm. *Electron. J. Probab.* 18, No. 95, 1-25.

[7] G. Pólya (1969). On the number of certain lattice polygons. J. Combin. Th. 6, 102-105.