

BOOTSTRAP PERCOLATION ON THE RANDOM GRAPH $G_{n,p}$

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Bootstrap percolation on the random graph $G_{n,p}$ is a process of spread of “activation” on a given realization of the graph with a given number of initially active nodes. At each step those vertices which have not been active but have at least $r \geq 2$ active neighbors become active as well.

We study the size A^* of the final active set. The parameters of the model are, besides r (fixed) and n (tending to ∞), the size $a = a(n)$ of the initially active set and the probability $p = p(n)$ of the edges in the graph. We show that the model exhibits a sharp phase transition: depending on the parameters of the model, the final size of activation with a high probability is either $n - o(n)$ or it is $o(n)$. We provide a complete description of the phase diagram on the space of the parameters of the model. In particular, we find the phase transition and compute the asymptotics (in probability) for A^* ; we also prove a central limit theorem for A^* in some ranges. Furthermore, we provide the asymptotics for the number of steps until the process stops.

1. Introduction. Bootstrap percolation on a graph G is defined as the spread of *activation* or *infection* according to the following rule, with a given threshold $r \geq 2$: We start with a set $\mathcal{A}(0) \subseteq V(G)$ of *active* vertices. Each inactive vertex that has at least r active neighbors becomes active. This is repeated until no more vertices become active, that is, when no inactive vertex has r or more active neighbors. Active vertices never become inactive, so the set of active vertices grows monotonously.

To avoid confusion, we will use the terminology that each active vertex *infects* all its neighbors, so that a vertex that is infected (at least) r times becomes active.

We are mainly interested in the final size A^* of the active set, and in particular whether eventually all vertices will be active or not. If they are, we say that the initial set $\mathcal{A}(0)$ *percolates* (completely). We will study a sequence of graphs of order $n \rightarrow \infty$; we then also say that (a sequence of) $\mathcal{A}(0)$ *almost percolates* if the number of vertices that remain inactive is $o(n)$, that is, if $A^* = n - o(n)$.

Bootstrap percolation on a lattice (which is a special example of a cellular automata) was introduced in 1979 by Chalupa, Leath and Reich [24] as a simplified

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model of some magnetic systems. Since then bootstrap percolation has been studied on various graphs, both deterministic and random. One can study either a random initial set or the deterministic problem of choosing an initial set that is optimal in some sense. A simple example of the latter is the classical folklore problem to find the minimal percolating set in a two-dimensional grid (i.e., a finite square $[n]^2$ in the square lattice); see Balogh and Pete [13] and Bollobás [18]. (These references also treat higher-dimensional grids $[n]^d$.) Another extremal problem is studied by Morris [39]. The problem with a random initial set was introduced by Chalupa, Leath and Reich [24] (lattices and regular infinite tree), and further studied on lattices by Schonmann [42]; it has, in particular, been studied on finite grids (in two dimensions or more), see Aizenman and Lebowitz [1], Balogh and Pete [13], Cerf and Cirillo [20], Cerf and Manzo [21], Holroyd [29], Balogh, Bollobás and Morris [9], Gravner, Holroyd and Morris [27]. In a recent paper, Balogh et al. [6] derived a sharp asymptotic for the critical density (i.e., the critical size of a random initial set) for bootstrap percolation on grids of any dimension, generalizing results of Balogh, Bollobás, and Morris [8]. Grids with a different edge set were studied by Holroyd, Liggett and Romik [30]. The study of bootstrap percolation on lattices is partly explained by its origin in statistical physics, and the bootstrap process is being successfully used in studies of the Ising model; see [22, 23, 26, 40]. Lately bootstrap percolation has also been studied on varieties of graphs different from lattices and grids; see, for example, Balogh and Bollobás [5] (hypercube); Balogh, Peres and Pete [12] (infinite trees); Balogh and Pittel [14], Janson [32] (random regular graphs); an extension where the threshold may vary between the vertices is studied by Amini [2]. An anisotropic bootstrap percolation was studied by Duminił-Copin and van Enter [25]. Further, a graph bootstrap percolation model introduced by Bollobás [17] already in 1968, where edges are infected instead of vertices, was analyzed recently by Balogh, Bollobás and Morris [10] and Balogh et al. [11].

In the present paper, we study bootstrap percolation on the Erdős–Rényi random graph $G_{n,p}$ with an initial set $\mathcal{A}(0)$ consisting of a given number a of vertices chosen at random. (By symmetry, we obtain the same results for any deterministic set of a vertices.) Recall that $G_{n,p}$ is the random graph on the set of vertices $V_n = \{1, \dots, n\}$ where all possible edges between pairs of different vertices are present independently and with the same probability p . As usual, we let $p = p(n)$ depend on n .

A problem equivalent to bootstrap percolation on $G_{n,p}$ in the case $p = \lambda/n$ was studied by Scalia-Tomba [41], although he used a different formulation as an epidemic. (Ball and Britton [3, 4] study a more general model with different degrees of severity of infection.) Otherwise, bootstrap percolation on $G_{n,p}$ was first studied by Vallier [47]; we here use a simple method (the same as [41]) that allows us to both simplify the proofs and improve the results. We will state the results for a general fixed $r \geq 2$ (the case $r = 1$ is much different; see Remark 5.9);

the reader may for simplicity consider the case $r = 2$ only, since there are no essential differences for higher r .

We will see that there is a threshold phenomenon: typically, either the final size A^* is small (at most twice the initial size a), or it is large [sometimes exactly n , but if p is so small that there are vertices of degree less than r , these can never become active except initially so eventually at most $n - o(n)$ will become infected]. We can study the threshold in two ways: in the first version, we keep n and p fixed and vary a . In the second version, we fix n and a and vary p . We will state some results for both versions and some for the former version only; these too can easily be translated to the second version. We will also study dynamical versions, where we add new external infections or activations or new edges until we reach the threshold; see Section 4.

Apart from the final size A^* , we will also study the time τ the bootstrap process takes until completion. We count the time in *generations*: generation 0 is $\mathcal{A}(0)$, generation 1 is the set of other vertices that have at least r neighbors in generation 0, and so on. The process stops as soon as there is an empty generation, and we let τ be the number of (nonempty) generations. Thus, if we let \mathcal{G}_k be the set of vertices activated in generation k , then

$$(1.1) \quad \tau := \max\{k \geq 0 : \mathcal{G}_k \neq \emptyset\} = \min\{k \geq 1 : \mathcal{G}_k = \emptyset\} - 1.$$

REMARK 1.1. Bootstrap percolation does not seem to be a good model for usual infectious diseases; see, however, Ball and Britton [3]. It might be a better model for the spread of rumors or other ideas or beliefs; cf. the well-known rule, “What I tell you three times is true” in Carroll [19].

Bootstrap percolation can be also viewed as a simplified model for propagation of activity in a neural network. Although related neuronal models are too involved for a rigorous analysis (see, e.g., [36, 44, 46]) they inspired study of bootstrap percolation on $G_{n,p}$ by Vallier [47]. There is a further discussion on the application of bootstrap percolation on $G_{n,p}$ to neuromodelling in [45].

REMARK 1.2. Instead of $G_{n,p}$, one might consider the random graph $G(n, m)$, with a given number $m = m(n)$ of edges. It is easy to obtain a result for $G(n, m)$ from our results for $G_{n,p}$, using monotonicity, but we usually leave this to the reader. [In the dynamical model in Section 4.3, we consider $G(n, m)$, however.]

REMARK 1.3. An alternative to starting with an initial active set of fixed size a is to let each vertex be initially activated with probability $\alpha = \alpha(n) > 0$, with different vertices activated independently. Note that this is the same as taking the initial size a random with $a \in \text{Bin}(n, \alpha)$. For most results the resulting random variation in a is negligible, and we obtain the same results as for $a = n\alpha$, but for the Gaussian limit in Theorems 3.6(iii) and 4.5, the asymptotic variances are changed by constant factors. We leave the details to the reader.

Some open problems arise from our study. In [9], Balogh, Bollobás and Morris determine the critical probability for bootstrap percolation on grids when the dimension $d = d(n) \rightarrow \infty$. A similar idea translated to the $G(n, p)$ graph would be to study what happens when $r = r(n) \rightarrow \infty$. This problem is not treated here although our methods might be useful also for such problems. The problem of majority percolation where a vertex becomes activated if at least half of its neighbors are active [$r(v) = d(v)/2$] has been studied on the hypercube by Balogh, Bollobás and Morris [7]. On the d -dimensional grid $d(v)/2 = d$ but on the $G(n, p)$ graph, this problem is completely different and still open. (We thank the referee for these suggestions.)

The method is described in Section 2. The main results are stated in Section 3, with further results in Sections 4 and 5. Proofs are given in Sections 6–11.

1.1. *Notation.* All unspecified limits are as $n \rightarrow \infty$. We use \xrightarrow{d} for convergence in distribution and \xrightarrow{P} for convergence in probability of random variables; we further use O_p and o_p in the standard sense (see, e.g., [33] and [34]), and we use w.h.p. (with high probability) for events with probability tending to 1 as $n \rightarrow \infty$. Note that, for example, “ $= o(1)$ w.h.p.” is equivalent to “ $= o_p(1)$ ” and to “ $\xrightarrow{P} 0$,” and that “ $\sim a_n$ w.h.p.” is equivalent to “ $= (1 + o_p(1))a_n$,” see [33]. A statement of the type “when \mathcal{P} , then w.h.p. \mathcal{Q} ” (or similar wording), where \mathcal{P} and \mathcal{Q} are two events, means that $\mathbb{P}(\mathcal{P} \text{ and (not } \mathcal{Q})) \rightarrow 0$, that is, that w.h.p. “(not \mathcal{P}) or \mathcal{Q} ” holds. (See, e.g., Theorem 3.2 and Proposition 10.10.) If $\mathbb{P}(\mathcal{P})$ is bounded away from 0, this is equivalent to “conditioned on \mathcal{P} , \mathcal{Q} holds w.h.p.”

If X_n is a sequence of random variables, and μ_n and σ_n^2 are sequences of real numbers, with $\sigma_n^2 > 0$, we say that $X_n \in \text{AsN}(\mu_n, \sigma_n^2)$ if $(X_n - \mu_n)/\sigma_n \xrightarrow{d} N(0, 1)$.

Occasionally we use the subsubsequence principle ([34], page 12), which says that to prove a limit result (e.g., for real numbers, or for random variables in probability or in distribution), it is sufficient to show that every subsequence has a subsubsequence where the result holds. We may thus, without loss of generality, select convenient subsequences in a proof, for example, such that another given sequence either converges or tends to ∞ .

2. A useful reformulation. In order to analyze the bootstrap percolation process on $G_{n,p}$, we change the time scale; we forget the generations and consider at each time step the infections from one vertex only. Choose $u_1 \in \mathcal{A}(0)$ and give each of its neighbors a *mark*; we then say that u_1 is *used*, and let $\mathcal{Z}(1) := \{u_1\}$ be the set of used vertices at time 1. We continue recursively: at time t , choose a vertex $u_t \in \mathcal{A}(t-1) \setminus \mathcal{Z}(t-1)$. We give each neighbor of u_t a new mark. Let $\Delta\mathcal{A}(t)$ be the set of inactive vertices with r marks; these now become active, and we let $\mathcal{A}(t) = \mathcal{A}(t-1) \cup \Delta\mathcal{A}(t)$ be the set of active vertices at time t . We finally

set $\mathcal{Z}(t) = \mathcal{Z}(t - 1) \cup \{u_t\} = \{u_s : s \leq t\}$, the set of used vertices. [We start with $\mathcal{Z}(0) = \emptyset$, and note that necessarily $\Delta\mathcal{A}(t) = \emptyset$ for $t < r$.]

The process stops when $\mathcal{A}(t) \setminus \mathcal{Z}(t) = \emptyset$, that is, when all active vertices are used. We denote this time by T ,

$$(2.1) \quad T := \min\{t \geq 0 : \mathcal{A}(t) \setminus \mathcal{Z}(t) = \emptyset\}.$$

Clearly, $T \leq n$; in particular, T is finite. The final active set is $\mathcal{A}(T)$; it is clear that this is the same set as the one produced by the bootstrap percolation process defined in the **Introduction**; only the time development differs. Hence we may as well study the version just described. [This is true for any choice of the vertices u_t . For definiteness, we may assume that we keep the unused, active vertices in a queue and choose u_t as the first vertex in the queue, and that the vertices in $\Delta\mathcal{A}(t)$ are added at the end of the queue in order of their labels. Thus u_t will always be one of the oldest unused, active vertices, which will enable us to recover the generations; see further Section 10. In Section 4, we consider other ways of choosing u_t .] This reformulation was used already by Scalia-Tomba [41] (for a more general model). It is related to the (continuous-time) construction by Sellke [43] for an epidemic process.

Let $A(t) := |\mathcal{A}(t)|$, the number of active vertices at time t . Since $|\mathcal{Z}(t)| = t$ and $\mathcal{Z}(t) \subseteq \mathcal{A}(t)$ for $t = 0, \dots, T$, we also have

$$(2.2) \quad T = \min\{t \geq 0 : A(t) = t\} = \min\{t \geq 0 : A(t) \leq t\}.$$

Moreover, since the final active set is $\mathcal{A}(T) = \mathcal{Z}(T)$, its size A^* is

$$(2.3) \quad A^* := A(T) = |\mathcal{A}(T)| = |\mathcal{Z}(T)| = T.$$

Hence, the set $\mathcal{A}(0)$ percolates if and only if $T = n$, and $\mathcal{A}(0)$ almost percolates if and only if $T = n - o(n)$.

We analyze this process by the standard method of revealing the edges of the graph $G_{n,p}$ only on a need-to-know basis. We thus begin by choosing u_1 as above and then reveal its neighbors; we then find u_2 and reveal its neighbors, and so on. Let, for $i \notin \mathcal{Z}(s)$, $I_i(s)$ be the indicator that there is an edge between the vertices u_s and i . This is also the indicator that i gets a mark at time s , so if $M_i(t)$ is the number of marks i has at time t , then

$$(2.4) \quad M_i(t) = \sum_{s=1}^t I_i(s),$$

at least until i is activated (and what happens later does not matter). Note that if $i \notin \mathcal{A}(0)$, then, for every $t \leq T$, $i \in \mathcal{A}(t)$ if and only if $M_i(t) \geq r$. The crucial feature of this description of the process, which makes the analysis simple, is that the random variables $I_i(s)$ are i.i.d. $\text{Be}(p)$.

We have defined $I_i(s)$ only for $s \leq T$ and $i \notin \mathcal{Z}(s)$, but it is convenient to add further (redundant) variables so that $I_i(s)$ are defined, and i.i.d., for all $i \in V_n$

and all $s \geq 1$. One way to do this formally is to reverse the procedure above. We start with i.i.d. $I_i(s) \in \text{Be}(p)$, for $i \in V_n$ and $s \geq 1$, and a set $\mathcal{A}(0) \subseteq V_n$. We let $\mathcal{Z}(0) := \emptyset$ and start with an empty graph on V_n . We then, as above, for $t = 1, \dots, n$ select $u_t \in \mathcal{A}(t - 1) \setminus \mathcal{Z}(t - 1)$ if this set is nonempty; otherwise we select $u_t \in V_n \setminus \mathcal{Z}(t - 1)$ (taking, e.g., the smallest such vertex). We define $M_i(t)$ by (2.4) for all $i \in V_n$ and $t \geq 0$, and update $\mathcal{A}(t) := \mathcal{A}(0) \cup \{i : M_i(t) \geq r\}$ and $\mathcal{Z}(t) := \mathcal{Z}(t - 1) \cup \{u_t\} = \{u_s : s \leq t\}$. Furthermore, add an edge $u_t i$ to the graph for each vertex $i \notin \mathcal{Z}(t)$ such that $I_i(t) = 1$. Finally, define T by (2.1) or (2.2).

It is easy to see that this constructs a random graph $G_{n,p}$ and that $\mathcal{A}(t)$, $t \leq T$, is as above for this graph, so the final active set of the bootstrap percolation on the graph is $\mathcal{A}(T)$.

Define also, for $i \in V_n \setminus \mathcal{A}(0)$,

$$(2.5) \quad Y_i := \min\{t : M_i(t) \geq r\}.$$

If $Y_i \leq T$, then Y_i is the time vertex i becomes active, but if $Y_i > T$, then i never becomes active. Thus, for $t \leq T$,

$$(2.6) \quad \mathcal{A}(t) = \mathcal{A}(0) \cup \{i \notin \mathcal{A}(0) : Y_i \leq t\}.$$

By (2.4), each $M_i(t)$ has a binomial distribution $\text{Bin}(t, p)$. Further, by (2.4) and (2.5), each Y_i has a negative binomial distribution $\text{NegBin}(r, p)$,

$$(2.7) \quad \mathbb{P}(Y_i = k) = \mathbb{P}(M_i(k - 1) = r - 1, I_i(k) = 1) = \binom{k - 1}{r - 1} p^r (1 - p)^{k-r};$$

moreover, these random variables Y_i are i.i.d.

We let, for $t = 0, 1, 2, \dots$,

$$(2.8) \quad S(t) := |\{i \notin \mathcal{A}(0) : Y_i \leq t\}| = \sum_{i \notin \mathcal{A}(0)} \mathbf{1}\{Y_i \leq t\},$$

so, by (2.6), and our notation $A(0) = a$,

$$(2.9) \quad A(t) = A(0) + S(t) = S(t) + a.$$

By (2.9), (2.2) and (2.3), it suffices to study the stochastic process $S(t)$. Note that $S(t)$ is a sum of $n - a$ i.i.d. processes $\mathbf{1}\{t \geq Y_i\}$, each of which is 0/1-valued and jumps from 0 to 1 at time Y_i , where Y_i has the distribution $\text{NegBin}(r, p)$ in (2.7). We write $S(t) = S_{n-a}(t)$ when we want to emphasize the number of summands in $S(t)$; more generally we define $S_m(t) := \sum_{i=1}^m \mathbf{1}\{Y_i \leq t\}$ for any $m \leq n$ [assuming for consistency that $\mathcal{A}(0) = \{n - a + 1, \dots, n\}$].

The fact that $S(t)$, and thus $A(t)$, is a sum of i.i.d. processes makes the analysis easy; in particular, for any given t ,

$$(2.10) \quad S(t) \in \text{Bin}(n - a, \pi(t)),$$

where

$$(2.11) \quad \pi(t) := \mathbb{P}(Y_1 \leq t) = \mathbb{P}(M_1(t) \geq r) = \mathbb{P}(\text{Bin}(t, p) \geq r).$$

In particular, we have

$$(2.12) \quad \mathbb{E} S(t) = (n - a)\pi(t),$$

$$(2.13) \quad \text{Var } S(t) = (n - a)\pi(t)(1 - \pi(t)) \leq \mathbb{E} S(t) \leq n\pi(t).$$

To avoid rounding to integers sometimes below, we define $S(t) := S(\lfloor t \rfloor)$ and $\pi(t) := \pi(\lfloor t \rfloor)$ for all real $t \geq 0$. We also sometimes (when it is obviously harmless) ignore rounding to simplify notation.

3. Main results.

3.1. *Limits in probability.* For given r, n , and p we define, for reasons that will be seen later,

$$(3.1) \quad t_c := \left(\frac{(r - 1)!}{np^r} \right)^{1/(r-1)},$$

$$(3.2) \quad a_c := \left(1 - \frac{1}{r} \right) t_c,$$

$$(3.3) \quad b_c := n \frac{(pn)^{r-1}}{(r - 1)!} e^{-pn}.$$

In particular, for $r = 2$, $t_c := 1/(np^2)$ and $a_c := 1/(2np^2)$. For future use, note also that (3.1) can be written

$$(3.4) \quad n \frac{(pt_c)^r}{r!} = \frac{t_c}{r}.$$

Our standard assumptions $p \ll n^{-1/r}$ and $p \gg n^{-1}$ imply that

$$(3.5) \quad \begin{aligned} t_c &\rightarrow \infty, & pt_c &\rightarrow 0, & t_c/n &\rightarrow 0, \\ a_c &\rightarrow \infty, & a_c/n &\rightarrow 0, & b_c/n &\rightarrow 0, & pb_c &\rightarrow 0, \end{aligned}$$

and further

$$(3.6) \quad \mathbb{E} S_n(t_c) = n\pi(t_c) \sim n \frac{(pt_c)^r}{r!} = \frac{t_c}{r},$$

$$(3.7) \quad \begin{aligned} n - \mathbb{E} S_n(n) &= n(1 - \pi(n)) = n\mathbb{P}(\text{Bin}(n, p) \leq r - 1) \\ &\sim n\mathbb{P}(\text{Bin}(n, p) = r - 1) \\ &\sim b'_c := n \frac{(pn)^{r-1}}{(r - 1)!} (1 - p)^n. \end{aligned}$$

If $p \ll n^{-1/2}$, then $(1 - p)^n \sim e^{-np}$ and (3.7) yields

$$(3.8) \quad n - \mathbb{E} S_n(n) = n(1 - \pi(n)) \sim b'_c \sim b_c;$$

if p is larger [$p = \Omega(n^{-1/2})$, i.e., $n^{-1/2} = O(p)$], this is not quite true, but in this case both b'_c and b_c decrease to 0 very fast; in all cases

$$(3.9) \quad n - \mathbb{E} S_n(n) = n(1 - \pi(n)) = b_c + o(b_c + 1).$$

Recall that our main interest is in $S(t) = S_{n-a}(t)$ rather than $S_n(t)$; see (2.10); for $S(t)$ we obviously have similar results, with additional error terms depending on a ; see (2.12) and, for example, (8.2).

Note further that by (3.3), for any $\beta \in (-\infty, \infty)$,

$$np - (\log n + (r - 1) \log(np)) \rightarrow \begin{cases} -\infty, \\ \beta, \\ \infty, \end{cases} \iff b_c \rightarrow \begin{cases} \infty, \\ (r - 1)!^{-1} e^{-\beta}, \\ 0, \end{cases}$$

which by simple calculations yields, provided $p \geq n^{-1}$,

$$(3.10) \quad np - (\log n + (r - 1) \log \log n) \rightarrow \begin{cases} -\infty, \\ \beta, \\ \infty, \end{cases} \\ \iff b_c \rightarrow \begin{cases} \infty, \\ (r - 1)!^{-1} e^{-\beta}, \\ 0. \end{cases}$$

Our first result, to be refined later, shows that the threshold for almost percolation is $a = a_c$. The proof of the theorems in this section are given later (Sections 8–10). Let us recall that A^* is the final size of the active set, and that $A^* = T = A(T) = a + S_{n-a}(T)$.

THEOREM 3.1. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$.*

(i) *If $a/a_c \rightarrow \alpha < 1$, then $A^* = (\varphi(\alpha) + o_p(1))t_c$, where $\varphi(\alpha)$ is the unique root in $[0, 1]$ of*

$$(3.11) \quad r\varphi(\alpha) - \varphi(\alpha)^r = (r - 1)\alpha.$$

[For $r = 2$, $\varphi(\alpha) = 1 - \sqrt{1 - \alpha}$.]

Further, $A^*/a \xrightarrow{P} \varphi_1(\alpha) := \frac{r}{r-1} \varphi(\alpha)/\alpha$, with $\varphi_1(0) := 1$.

(ii) *If $a/a_c \geq 1 + \delta$, for some $\delta > 0$, then $A^* = n - o_p(n)$; in other words, we have w.h.p. almost percolation. More precisely, $A^* = n - O_p(b_c)$.*

(iii) *In case (ii), if further $a \leq n/2$, say, we further have complete percolation, that is, $A^* = n$ w.h.p., if and only if $b_c \rightarrow 0$, that is, if and only if $np - (\log n + (r - 1) \log \log n) \rightarrow \infty$.*

It is easily verified that φ_1 is a continuous, strictly increasing function $[0, 1] \rightarrow [1, r/(r - 1)]$. In particular, in the subcritical case (i), we thus have w.h.p. $A^* < (r/(r - 1))a \leq 2a$, so the activation will not spread to many more than the originally active nodes.

In the supercritical case (ii), we have the following more detailed result.

THEOREM 3.2. *Suppose that $r \geq 2$, $n^{-1} \ll p \ll n^{-1/r}$ and $a = o(n)$, and that $A^* = n - o_p(n)$ as, for example, in Theorem 3.1(ii). Then:*

(i) *If $np - (\log n + (r - 1) \log \log n) \rightarrow -\infty$, so $b_c \rightarrow \infty$ by (3.10), then $A^* = n - b_c(1 + o_p(1))$. In particular, w.h.p. we do not have complete percolation.*

(ii) *If $np - (\log n + (r - 1) \log \log n) \rightarrow \infty$, so $b_c \rightarrow 0$ by (3.10), then w.h.p. $A^* = n$, so we have complete percolation.*

(iii) *If $np - (\log n + (r - 1) \log \log n) \rightarrow \beta \in (-\infty, \infty)$, so $b_c \rightarrow b > 0$ by (3.10), then $n - A^* \xrightarrow{d} \text{Po}(b)$; in particular, $\mathbb{P}(A^* = n) \rightarrow \exp(-b) \in (0, 1)$.*

More generally, even if we do not have almost percolation w.h.p., the result holds w.h.p. provided $A^ \geq 3t_c$.*

By the last statement we mean that $\mathbb{P}(\text{the result fails and } A^* \geq 3t_c) \rightarrow 0$. In particular, it holds w.h.p. conditioned on $A^* \geq 3t_c$, provided we have $\liminf \mathbb{P}(A^* \geq 3t_c) > 0$.

REMARK 3.3. Let \mathcal{B} be the set of vertices in $G_{n,p}$ with degrees less than r . These are never activated unless they happen to be in the initially active set $\mathcal{A}(0)$, and for each of the vertices, this has probability $a/n \rightarrow 0$ if $a = o(n)$; hence trivially $A^* \leq n - |\mathcal{B}|(1 - o_p(1))$. We have [cf. (3.7) and (3.9)]

$$\mathbb{E} |\mathcal{B}| = n\mathbb{P}(\text{Bin}(n - 1, p) \leq r - 1) \sim b_c + o(b_c + 1)$$

with concentration of $|\mathcal{B}|$ around its mean if $b_c \rightarrow \infty$ and a limiting Poisson distribution if $b_c \rightarrow b < \infty$; see [34], Sections 6.2 and 6.3, and [15]. Comparing this with Theorem 3.2 we see that in the supercritical case, and with $a = o(n)$, the final inactive set $V_n \setminus \mathcal{A}(T)$ differs from \mathcal{B} by $o_p(|\mathcal{B}|)$ vertices only, and in the case $b_c = O(1)$ [combining cases (ii) and (iii) in Theorem 3.2], w.h.p. $V_n \setminus \mathcal{A}(T) = \mathcal{B}$. In other words, when we get a large active set, the vertices that remain inactive are mainly the ones with degrees less than r , and if further $b_c = O(1)$, they are w.h.p. exactly the vertices with degrees less than r .

We can, as discussed earlier, also consider thresholds for p for a given a .

THEOREM 3.4. *Suppose that $r \geq 2$ and that $a \rightarrow \infty$ with $a = o(n)$. Then the threshold for p for almost percolation is*

$$(3.12) \quad p_c := \left(\frac{(r - 1)^{r-1} (r - 1)!}{r^{r-1}} \right)^{1/r} (na^{r-1})^{-1/r}$$

in the sense that if, for some $\delta > 0$, $p \leq (1 - \delta)p_c$, then $A^ \leq 2a = o(n)$ w.h.p., while if $p \geq (1 + \delta)p_c$, then $A^* = n - o(n)$ w.h.p. In the latter case, further $A^* = n$ w.h.p. if and only if $p = (\log n + (r - 1) \log \log n + \omega(n))/n$ for some $\omega(n) \rightarrow \infty$.*

Note that $n^{-1} \ll p_c \ll n^{-1/r}$. Equation (3.12) is the inverse to (3.2) in the sense that the functions $a \mapsto p_c$ and $p \mapsto a_c$ that they define are the inverses of each other. For $r = 2$, (3.12) simplifies to $p_c = (2na)^{-1/2}$.

REMARK 3.5. Note that the thresholds for complete and almost percolation are different only for large a . Indeed, for such a case the threshold p_c for almost percolation can be so small that the graph $G(n, p_c)$ may not be even connected. Then, besides p_c , we have the second threshold for the complete percolation; for example, if $a = n/\log n$ and $r = 2$, there are two thresholds: $p_c = \Theta(\sqrt{\log n}/n)$ for almost percolation, and $\Theta(\log(n)/n)$ for complete percolation. If a is small enough so that $G(n, p_c)$ is dense enough (e.g., if $a \leq 0.49n/\log^2 n$ when $r = 2$), these two thresholds coincide.

3.2. *Gaussian limits.* To study the threshold at a_c more precisely, we approximate $\pi(t)$ in (2.11) by the corresponding Poisson probability,

$$(3.13) \quad \tilde{\pi}(t) := \mathbb{P}(\text{Po}(tp) \geq r) = \psi(tp) := \sum_{j=r}^{\infty} \frac{(pt)^j}{j!} e^{-pt}.$$

Note that ψ is a differentiable, increasing function on $(0, \infty)$, and that

$$(3.14) \quad \frac{d}{dt} \tilde{\pi}(t) = p\psi'(pt) = p \frac{(pt)^{r-1}}{(r-1)!} e^{-pt} = \frac{p^r t^{r-1}}{(r-1)!} e^{-pt}.$$

By a standard estimate for Poisson approximation of a binomial distribution (see, e.g., [15], Theorem 2.M),

$$(3.15) \quad |\pi(t) - \tilde{\pi}(t)| \leq d_{\text{TV}}(\text{Bin}(t, p), \text{Po}(tp)) < p,$$

where d_{TV} denotes the total variation distance. A sharper estimate for small t will be given in Lemma 9.4.

We define, for given n and p ,

$$(3.16) \quad a_c^* := - \min_{t \leq 3t_c} \frac{n\tilde{\pi}(t) - t}{1 - \tilde{\pi}(t)},$$

and let $t_c^* \in [0, 3t_c]$ be the point where the minimum is attained. Under our standard assumptions $n^{-1} \ll p \ll n^{-1/r}$, for $t \leq 3t_c$, when $pt \rightarrow 0$ by (3.5), we have, by (3.13) and (3.4),

$$(3.17) \quad n\tilde{\pi}(t) \sim n \frac{(pt)^r}{r!} = \left(\frac{t}{t_c}\right)^r t_c$$

and thus $\tilde{\pi}(t) \rightarrow 0$ and $1 - \tilde{\pi}(t) \sim 1$; it follows easily that $a_c^* \sim a_c$ and $t_c^* \sim t_c$. More precise estimates are given in Lemma 9.5, where it also is shown that t_c^* is unique (for large n , at least). Furthermore, by Lemma 9.1 below, for large n , the

minimum in (3.16) could as well be taken over $t \leq n/2$, say, since $n\tilde{\pi}(t) - t \geq 0$ for $t \in [3t_c, n/2]$.

The following theorem shows that the precise threshold for a is $a_c^* \pm O(\sqrt{a_c})$, with a width of the threshold of the order $\sqrt{a_c} \sim \sqrt{a_c^*}$. Φ denotes the standard normal distribution function. Note that Theorem 3.2 applies, provided $a = o(n)$, and provides more detailed information on A^* when A^* is large [i.e., in (ii) and in (iii) conditioned on, say, $A^* \geq 3t_c$].

THEOREM 3.6. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$.*

(i) *If $(a - a_c^*)/\sqrt{a_c} \rightarrow -\infty$, then for every $\varepsilon > 0$, w.h.p. $A^* \leq t_c^* \leq t_c(1 + \varepsilon)$. If further $a/a_c^* \rightarrow 1$, then $A^* = (1 + o_p(1))t_c$.*

(ii) *If $(a - a_c^*)/\sqrt{a_c} \rightarrow +\infty$, then $A^* = n - O_p(b_c)$.*

(iii) *If $(a - a_c^*)/\sqrt{a_c} \rightarrow y \in (-\infty, \infty)$, then for every $\varepsilon > 0$ and every $b^* \gg b_c$ with $b^* = o(n)$,*

$$\mathbb{P}(A^* > n - b^*) \rightarrow \Phi((r - 1)^{1/2}y),$$

$$\mathbb{P}(A^* \in [(1 - \varepsilon)t_c, (1 + \varepsilon)t_c]) \rightarrow 1 - \Phi((r - 1)^{1/2}y).$$

For the corresponding result when we keep a fixed and change p , we define, for given n and a ,

$$(3.18) \quad \gamma(p) := \inf_{t \leq n/2} \{(n - a)\tilde{\pi}(t) - t\}.$$

Since $\tilde{\pi}$ is an increasing function of p , $\gamma(p)$ is increasing, with $\gamma(0) = -n/2$ and, provided, for example, $a = o(n)$, $\gamma(1) = o(1)$ [attained at $t = o(1)$]. Given $a = a(n) \rightarrow \infty$ with $a = o(n)$, there is thus (for large n) a unique p_c^* such that

$$(3.19) \quad \gamma(p_c^*) = -a.$$

We will see in Lemma 9.2 that $p_c^* \sim p_c$. It is easily verified that, for large n at least, $a_c^* = a \iff \gamma(p) = -a$, and thus $p \mapsto a_c^*$ and $a \mapsto p_c^*$ are the inverses of each other.

THEOREM 3.7. *Suppose $r \geq 2$ and $a \rightarrow \infty$ with $a = o(n)$.*

(i) *If $(p/p_c^* - 1)a^{1/2} \rightarrow -\infty$, then $A^* \leq ((r/(r - 1) + o_p(1))a$. If further $p/p_c^* \rightarrow 1$, then $A^* = ((r/(r - 1) + o_p(1))a$.*

(ii) *If $(p/p_c^* - 1)a^{1/2} \rightarrow +\infty$, then $A^* = n - o_p(n)$; if further $np - (\log n + (r - 1) \log \log n) \rightarrow \infty$, then $A^* = n$ w.h.p.*

(iii) *If $(p/p_c^* - 1)a^{1/2} \rightarrow \lambda \in (-\infty, \infty)$, then for every $\varepsilon > 0$,*

$$(3.20) \quad \mathbb{P}(A^* > (1 - \varepsilon)n) \rightarrow \Phi(r(r - 1)^{-1/2}\lambda),$$

$$(3.21) \quad \mathbb{P}\left(A^* \in \left[\left(\frac{r}{r - 1} - \varepsilon\right)a, \left(\frac{r}{r - 1} + \varepsilon\right)a\right]\right) \rightarrow 1 - \Phi(r(r - 1)^{-1/2}\lambda).$$

If further $np - (\log n + (r - 1) \log \log n) \rightarrow \infty$, then (3.20) can be replaced by $\mathbb{P}(A^ = n) \rightarrow \Phi(r(r - 1)^{-1/2}\lambda)$.*

In the subcritical cases in Theorems 3.1(i) and 3.6(i), we also obtain a Gaussian limit for the size of the final active set.

THEOREM 3.8. *Suppose $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Let t_* be the smallest positive root of*

$$(3.22) \quad (n - a)\tilde{\pi}(t_*) + a - t_* = 0.$$

(i) *If $a/a_c \rightarrow \alpha \in (0, 1)$, then $t_* \sim \varphi(\alpha)t_c$ with $\varphi(\alpha) \in (0, 1)$ given by (3.11), and $A^* \in \text{AsN}(t_*, \varphi_2(\alpha)t_c)$, where $\varphi_2(\alpha) := \varphi(\alpha)^r(1 - \varphi(\alpha)^{r-1})^{-2}/r$.*

(ii) *If $a/a_c \rightarrow 1$ and also $(a - a_c^*)/\sqrt{a_c} \rightarrow -\infty$, then $t_* \sim t_c$, more precisely*

$$(3.23) \quad t_* = t_c^* - (1 + o(1))\sqrt{\frac{2t_c}{r-1}(a_c^* - a)}$$

and

$$A^* \in \text{AsN}\left(t_*, \frac{t_c}{2(r-1)^2(1 - a/a_c^*)}\right).$$

REMARK 3.9. It follows from the proof that in both cases, for large n at least, t_* is the unique root of (3.22) in $[0, t_c^*]$. In (i), also $t_* < t_c$, so t_* is the unique root in $[0, t_c]$. In (ii), this is not always true. By Lemma 9.5, still for large n , $t_c^* > t_c$ and $t_c^* - t_c \sim pt_c^2/(r-1)$. If, for example, $r = 2$ and $p = \log n/n$, then $t_c^* - t_c \sim n/\log^3 n$, while $a = a_c^* - \sqrt{n}$ yields $t_c^* - t_* \sim \sqrt{2}n^{3/4}/\log n \ll t_c^* - t_c$.

3.3. The number of generations. In the supercritical case $a - a_c^* \gg \sqrt{a_c}$, when $\mathcal{A}(0)$ w.h.p. almost percolates by Theorem 3.6, we have the following asymptotic formula for the number of generations until the bootstrap percolation process stops.

THEOREM 3.10. *Suppose that $r \geq 2$, $n^{-1} \ll p \ll n^{-1/r}$ and $a = o(n)$. Assume that $a - a_c^* \gg \sqrt{a_c}$ [so that $\mathcal{A}(0)$ w.h.p. almost percolates]. Then, w.h.p.,*

$$(3.24) \quad \tau \sim \frac{\pi\sqrt{2}}{\sqrt{r-1}}\left(\frac{t_c}{a - a_c^*}\right)^{1/2} + \frac{1}{\log r}\left(\log \log(np) - \log_+ \log \frac{a}{a_c}\right) + \frac{\log n}{np} + O_p(1).$$

This theorem is an immediate consequence of Propositions 10.1, 10.4, 10.7 and 10.10 in Section 10. Moreover, these propositions show that the three terms [excepting the error term $O_p(1)$] in the formula (3.24) are the numbers of generations required for three distinct phases of the evolution: the beginning including (possibly) a bottleneck when the size is about t_c ; a period of doubly exponential growth; and a final phase where the last vertices are activated. Note that each of the three terms may be the largest one.

EXAMPLE 3.11. Let $p = n^{-\alpha}$, with $1/r < \alpha < 1$, and suppose $a = O(a_c)$. Then the third term in (3.24) is $O(1)$ and can be ignored while the second term is $\log \log n / \log r + O(1)$. If we are safely supercritical, say $a = 2a_c$, then the first term too is $O(1)$ and the result is $\tau \sim \log \log n / \log r$ w.h.p., dominated by the second term.

If instead the process is only barely supercritical, with $a = a_c^* + a_c^\beta$ say, with $1/2 < \beta < 1$, then the first term in (3.24) is Cn^γ with $C > 0$ and the exponent $\gamma = \frac{1-\beta}{2} \cdot \frac{r\alpha-1}{r-1}$, which dominates the other terms. Note that the exponent here can be any positive number in $(0, 1/4)$ (with $\gamma \approx 1/4$ if $\alpha \approx 1$ and $\beta \approx 1/2$, so the graph is very sparse and the initial set is minimal).

Finally, if $p = \log \log n / n$, say, so the graph is very sparse, and $a = 2a_c$, then again the first term in (3.24) is $O(1)$, the second is $O(\log \log \log \log n)$, while the third is $\log n / \log \log n$, which thus dominates the sum.

Note that the second term is $O(\log \log n)$, and the third is $o(\log n)$ [and in many cases $O(1)$ so it can be ignored], while the first term may be as large as $n^{1/4-o(1)}$ [although it too in many cases is $O(1)$].

REMARK 3.12. In the subcritical case, one could presumably obtain similar results for the number of generations until the process stops, but we have not pursued this topic here.

4. Dynamical models. We usually assume, as above, that a and p are given, but we can also consider dynamical models where one of them grows with time.

4.1. *Adding external activations.* In the first dynamical model, we let n and p be given, and consider a realization of $G_{n,p}$. We start with all vertices inactive (and completely uninfected). We then activate the vertices (from the outside) one by one, in random order. After each external activation, the bootstrap percolation mechanism works as before, activating all vertices that have at least r active neighbors until no such vertices remain; this is done instantaneously (or very rapidly) so that this is completed before the next external activation. Let A_0 be the number of externally activated vertices the first time that the active set \mathcal{A} is “big” in some sense. For example, for definiteness, we may define “big” as $|\mathcal{A}| > n/2$. [It follows from Theorem 3.1 that any threshold $|\mathcal{A}| > cn$ for a constant $c \in (0, 1)$ will give the same asymptotic results, as well as thresholds tending to 0 or n sufficiently slowly. If $np - (\log n + (r - 1) \log \log n) \rightarrow \infty$, we may also choose the condition $|\mathcal{A}| = n$, that is, complete percolation $\mathcal{A} = V_n$.] Then A_0 is a random variable (depending both on the realization of $G_{n,p}$ and on the order of external activations). In this formulation, the threshold result in Theorem 3.1 may be stated as follows.

THEOREM 4.1. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then $A_0/a_c \xrightarrow{P} 1$.*

PROOF. The active set after a external activations is the same as the final active set $\mathcal{A}(T)$ in the static model considered in the rest of this paper with these vertices chosen to be active initially. Hence, for any given a , $A_0 \leq a$ if and only if bootstrap percolation with a initially active yields a big final active set. In particular, if $\delta > 0$, then Theorem 3.1(i) implies that $\mathbb{P}(A_0 \leq (1 - \delta)a_c) \rightarrow 0$, while Theorem 3.1(ii) and (iii) imply that $\mathbb{P}(A_0 \leq (1 + \delta)a_c) \rightarrow 1$. \square

More precisely, Theorem 3.6 yields a Gaussian limit.

THEOREM 4.2. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then $A_0 \in \text{AsN}(a_c^*, a_c/(r - 1))$.*

PROOF. Let $x \in (-\infty, \infty)$. Then, arguing as in the proof of Theorem 4.1 but now using Theorem 3.6(iii) (with $y = x/\sqrt{r - 1}$), we find

$$\mathbb{P}\left(\frac{A_0 - a_c^*}{\sqrt{a_c/(r - 1)}} \leq x\right) = \mathbb{P}(A_0 \leq a_c^* + x\sqrt{a_c/(r - 1)}) \rightarrow \Phi(x). \quad \square$$

We have here for simplicity assumed that the external activations are done by sampling without replacement, but otherwise independently of whether the vertices already are (internally) activated. A natural variation is to only activate vertices that are inactive. Let A'_0 be the number of externally activated vertices when the active set becomes big in this version. Since a new activation of an already active vertex does not matter at all, A'_0 equals in the version above the number of externally active vertices among the first A_0 that are not already internally activated. Thus $A_0 - A'_0$ is the number of external activations that hit an already active vertex. It is easily verified that this is $o_p(a_c)$, and thus Theorem 4.1 holds for A'_0 as well; we omit the details. It seems likely that it is possible to derive a version of the Gaussian limit in Theorem 4.2 for A'_0 too, but that would require a more careful estimate of $A_0 - A'_0$ (and in particular its variance), which we have not done, so we leave this possibility as an open problem.

REMARK 4.3. One way to think about this dynamical model, where we add new active vertices successively and may think of these as being initially active, is to see it as a sequence of bootstrap percolation processes, one for each $a = 0, 1, \dots, n$; the processes live on the same graph $G_{n,p}$ but have different numbers of initially active vertices, and they are coupled in a natural way. In order to really have the same realization of $G_{n,p}$ for different a , we have to be careful in the choice of the order in which we explore the vertex neighborhoods, that is, the choice of u_t . [Recall that $G_{n,p}$ is constructed from the indicators $I_i(s)$ and the sequence (u_t) ; see Section 2.] We can achieve this by first making a list L of all vertices in the (random) order in which they are externally activated. We then at each time t choose u_t as an unused internally activated vertex (e.g., the most

recent one) if there is any such vertex, and otherwise as the next unused vertex in the list L .

This model makes it possible to pose now questions about the bootstrap percolation. For example, we may consider the critical process starting with exactly A_0 initially active vertices (i.e., the first process that grows beyond the bottleneck and becomes big) and ask for the number of generations until the process dies out. Alternatively, we may consider the process starting with exactly $A_0 - 1$ initially active vertices (i.e., the last process that does not become big) and ask for its final size.

Such questions will not be treated in the present paper, but we mention that it is easily seen that the final size with $A_0 - 1$ initially active vertices is $t_c(1 + o_p(1))$ so that the final size jumps from about t_c to about n with the addition of a single additional initial vertex. Furthermore, we conjecture that, under suitable conditions, the number of generations for the process with A_0 initially active vertices is of order $a_c^{1/3}$ (which is much larger than the number of generations for any fixed a ; see Section 3.3).

4.2. *Adding external infections.* An alternative to external activations is external infections, where we again start with all vertices inactive and uninfected, and infect vertices one by one from the outside, choosing the infected vertices at random (independently and with replacement); as before, r infections (external or internal) are needed for activation, and active vertices infect their neighbors. Let J_0 be the number of external infections when the active set first becomes “big” (as in Section 4.1). (Thus, J_0 is a random variable.)

In the original model, each initially active vertex infects about np other vertices so the total number of initial infections is about npa ; it is thus easy to guess that $J_0 \approx npa_c$. Indeed, this is the case as is shown by the next theorem. We cannot (as far as we know) directly derive this from our previous results, since the dependencies between infections in the two versions are slightly different, but it follows by a minor variation of our method; see Section 8. We believe that the result could be sharpened to a Gaussian limit as in Theorem 4.2, but we leave this to the reader.

THEOREM 4.4. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then $J_0/(np \times a_c) \xrightarrow{P} 1$.*

In particular, for $r = 2$, we thus have $J_0 \stackrel{P}{\sim} npa_c = 1/(2p)$.

4.3. *Adding edges.* In the second dynamical model, n and a are given; we start with n vertices of which a are active, but no edges. We then add the edges of the complete graph K_n one by one, in random order. As in the previous dynamical model, bootstrap percolation takes place instantaneously after each new edge is added.

It is convenient to use the standard method of adding the edges at random times (as in, e.g., [31]). Thus, each edge e in K_n is added at a time U_e , where U_e are independent and uniformly distributed on $[0, 1]$. Then, at a time $u \in [0, 1]$, the resulting graph is $G_{n,p}$ with $p = u$. (We use u to denote this time variable, in order not to confuse it with the time t used to describe the bootstrap percolation process.)

Let the random variable M be the number of edges required to obtain a big active set \mathcal{A} , where “big” is defined as in Section 4.1.

THEOREM 4.5. *Suppose $r \geq 2$ and $a \rightarrow \infty$ with $a = o(n)$. Then*

$$M = \binom{n}{2} p_c(1 + o_p(1)) = \frac{1}{2}n^2 p_c(1 + o_p(1)).$$

More precisely,

$$M \in \text{AsN}\left(\binom{n}{2} p_c^*, \frac{r-1}{4r^2} \frac{(n^2 p_c)^2}{a}\right).$$

The proof is given in Section 9.

REMARK 4.6 (Coupling different p). The proof of Theorem 4.5 is based on using our earlier results for a single p . We might also want to study the bootstrap percolation process for all p at once [or equivalently, in $G(n, m)$ for all m at once], that is, with a coupling of the models for different p , for given n and a . As in Remark 4.3, this requires a careful choice of the order in which the vertices are inspected. We can achieve this by modifying the formulation in Section 2 as follows:

When we have chosen a vertex u_t , we reveal the times U_{tj} that the edges from it appear; this tells us the neighborhood of u_t at any time u . We begin by choosing u_1, \dots, u_a as the initially active vertices. We then, after each choice of u_t , $t \geq a$, calculate for each of the remaining $n - t$ vertices the time when it acquires the r th edge to $\{u_1, \dots, u_t\}$, and let u_{t+1} be the vertex such that this time is minimal. Then, fixing any time $u = p$, the chosen vertices u_t will all be active until the first time that no unused active vertices remain, and the process stops. In this manner, we have found a choice of u_1, u_2, \dots that satisfies the description in Section 2 for all $p \in [0, 1]$ simultaneously.

As in Remark 4.3, we can use this model to study, for example, the last “small” or the first “big” bootstrap percolation process, when we add edges one by one with a given set of initially active vertices. Again, we will not consider such questions in the present paper.

5. Boundary cases. We have above assumed $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. In this section we treat the cases when these assumptions do not hold. Proofs are given in Section 11.

We begin with the sparse case $p(n) \sim c/n$ when t_c and a_c defined by (3.1) and (3.2) are of order n . (The exact values are no longer relevant, since they are based on approximations no longer valid.) This suggests that the interesting case is when $a \asymp n$, that is, when a positive fraction of all vertices are initially active. Indeed, Theorem 5.2 shows that, w.h.p., if we start with a positive fraction of the graph, then the activation spreads to a larger part of the graph but does not reach almost all vertices; if, on the contrary, the size of the original set of activated vertices is negligible with respect to the size of the graph, then the activation does not spread to a positive fraction of the graph. Provided c is large enough, there is, as found by Scalia-Tomba [41], also in this case a dichotomy, or “phase transition,” similar to Theorem 3.1, with a sudden jump from a “small” to a “large” final active set, although in this case all sets are of order n so the jump is not as dramatic as for larger p .

Define, for $x \geq 0, c \geq 0$ and $\theta \in [0, 1]$,

$$\begin{aligned}
 (5.1) \quad f(x, c, \theta) &:= (1 - \theta)\mathbb{P}(\text{Po}(cx) \geq r) + \theta - x \\
 &= (1 - \theta) \sum_{j=r}^{\infty} \frac{(cx)^j}{j!} e^{-cx} - x + \theta
 \end{aligned}$$

$$(5.2) \quad = 1 - x - (1 - \theta)\mathbb{P}(\text{Po}(cx) \leq r - 1)$$

$$(5.3) \quad = 1 - x - (1 - \theta) \sum_{j=0}^{r-1} \frac{(cx)^j}{j!} e^{-cx},$$

and let $x_0(\theta)$ be the smallest root $x \geq 0$ of

$$(5.4) \quad f(x, c, \theta) = 0;$$

similarly, let $x_1(\theta)$ be the largest root in $[0, 1]$ of this equation.

Since $f(0, c, \theta) = \theta \geq 0$ and $f(1, c, \theta) = -(1 - \theta)\mathbb{P}(\text{Po}(c) \leq r - 1) \leq 0$, there is always at least one root in $[0, 1]$, and $0 \leq x_0(\theta) \leq x_1(\theta) \leq 1$; further $0 < x_0(\theta) \leq x_1(\theta) < 1$ when $0 < \theta < 1$ while $x_0(0) = 0$ and $x_0(1) = x_1(1) = 1$. We also define

$$(5.5) \quad c_c = c_c(r) := r + \frac{\mathbb{P}(\text{Po}(r - 1) \leq r - 2)}{\mathbb{P}(\text{Po}(r - 1) = r - 1)} = r + \frac{\sum_{j=0}^{r-2} (r - 1)^j / j!}{(r - 1)^{r-1} / (r - 1)!}.$$

Thus $c_c(2) = 3, c_c(3) = 9/2, c_c(4) = 53/9$.

LEMMA 5.1. (i) *If $0 \leq c \leq c_c$, then (5.4) has a unique root $x = x_0(\theta) \in [0, 1]$ for every $\theta \in [0, 1]$, and $x_0(\theta)$ is a continuous strictly increasing function of θ .*

(ii) *If $c > c_c$, then there exists $\theta_c^- = \theta_c^-(c)$ and $\theta_c = \theta_c(c)$ with $0 \leq \theta_c^- < \theta_c < 1$ such that (5.4) has three roots in $[0, 1]$ when $\theta \in (\theta_c^-, \theta_c)$, but a unique root*

when $\theta \in [0, \theta_c^-)$ or $\theta \in (\theta_c, 1]$; if $\theta = \theta_c^- > 0$ or $\theta = \theta_c$, there are two roots, one of them double. The smallest root $x_0(\theta)$ is strictly increasing and continuous on $[0, 1]$ except at θ_c where it has a jump from $x_0(\theta_c)$ to $x_1(\theta_c) > x_0(\theta_c)$, where $x_1(\theta_c) = x_0(\theta_c+) := \lim_{\theta \searrow \theta_c} x_0(\theta)$ is the other root for $\theta = \theta_c$. Furthermore, if $\theta = \theta_c$, then $f(x, c, \theta) \geq 0$ for $x \in [0, x_1(\theta)]$, and $x_0(\theta)$ is a double root.

THEOREM 5.2. *Suppose that $r \geq 2$, $p \sim c/n$ and $a \sim \theta n$ for some constants $c \geq 0$ and $\theta \geq 0$.*

- (i) *If $\theta = 0$, that is, if $a = o(n)$, then $A^*/a \xrightarrow{P} 1$.*
- (ii) *If $c = 0$, that is, if $p = o(1/n)$, then $A^*/a \xrightarrow{P} 1$.*
- (iii) *If $0 \leq c \leq c_c$, then $A^*/n \xrightarrow{P} x_0(\theta)$, where $x_0(\theta)$ is the unique nonnegative root of (5.4).*
- (iv) *If $c > c_c$ and $\theta \neq \theta_c(c)$ given by Lemma 5.1, then $A^*/n \xrightarrow{P} x_0(\theta)$, where $x_0(\theta)$ is the smallest nonnegative root of (5.4).*

There is thus a jump in the final size at $a = \theta_c n$. Remark 11.1 shows how to find θ_c .

REMARK 5.3. $\theta_c(c)$ and $\theta_c^-(c)$ are decreasing functions of c . [$\theta_c(c)$ is strictly decreasing, while $\theta_c^-(c)$ is constant 0 for large c .] Hence their largest value is, by the calculation in Remark 11.1,

$$\begin{aligned} \theta_c^* &= \theta_c^*(r) := \theta_c(c_c) = \theta_c^-(c_c) \\ &= 1 - \frac{1}{r\mathbb{P}(\text{Po}(r-1) = r-1) + \mathbb{P}(\text{Po}(r-1) \leq r-2)}. \end{aligned}$$

Thus $\theta_c^*(2) = 1 - e/3$, $\theta_c^*(3) = 1 - e^2/9$, $\theta_c^*(4) = 1 - 2e^3/53$.

The threshold for $\theta_c^- = 0$ can be calculated too. For $r = 2$, $\theta_c^-(c) = 0$ for $c \geq e^y/y$, where $e^y = 1 + y + y^2$; numerically, this is $c \geq 3.35091\dots$

REMARK 5.4. We have here considered a given $p \sim c/n$ and varied $a \sim \theta n$. If we instead, as in Theorem 3.4, take a given $a \sim \theta n$ for a fixed θ and vary $c = pn$, we have a similar phenomenon. Lemma 5.1 and Theorem 5.2 apply for every combination of θ and c , and by considering the set of $(c, \theta) \in \mathbb{R}_+^2$ such that (5.4) has two or three roots, it follows from Remark 5.3 that if $\theta \geq \theta_c^*$, then $A^*/n \xrightarrow{P} x_0(c)$, where $x_0(c)$ is the unique root of (5.4) and thus a continuous function of c , while if $\theta < \theta_c^*$, then there is a range of c where (5.4) has three roots, and one value of c where the limit value $x_0(c)$ jumps from a “small” to a “large” value. Thus there is, again, a kind of phase transition.

The following theorem shows that if we for simplicity take $p = c/n$, then the precise threshold for a in Theorem 5.2(iv) is $\theta_c n \pm O(\sqrt{n})$, with a width of the threshold of the order \sqrt{n} .

THEOREM 5.5. *Suppose that $r \geq 2$ and $p = \frac{c}{n}$ with $c > c_c$ fixed. Let $\theta_c = \theta_c(c)$, $x_0 = x_0(\theta_c)$ and $x_1 = x_1(\theta_c)$ be as in Lemma 5.1; thus x_0 and x_1 are the two roots in $[0, 1]$ of $f(x, c, \theta_c) = 0$, with $x_0 < x_1$.*

(i) *If $(a - \theta_c n)/\sqrt{n} \rightarrow -\infty$, then for any $\varepsilon > 0$, w.h.p. $A^* \leq (1 + \varepsilon)x_0 n$. If further $a \sim \theta_c n$, then $A^* = (1 + o_p(1))x_0 n$.*

(ii) *If $(a - \theta_c n)/\sqrt{n} \rightarrow +\infty$, then for any $\varepsilon > 0$, w.h.p. $A^* \geq (1 - \varepsilon)x_1 n$. If further $a \sim \theta_c n$, then $A^* = (1 + o_p(1))x_1 n$.*

(iii) *If $(a - \theta_c n)/\sqrt{n} \rightarrow y \in (-\infty, \infty)$, then there exists a sequence $\varepsilon_n \rightarrow 0$ such that*

$$\mathbb{P}(A^* \in [(1 - \varepsilon_n)x_1 n, (1 + \varepsilon_n)x_1 n]) \rightarrow \Phi(y/\sigma),$$

$$\mathbb{P}(A^* \in [(1 - \varepsilon_n)x_0 n, (1 + \varepsilon_n)x_0 n]) \rightarrow 1 - \Phi(y/\sigma),$$

where $\sigma^2 = (1 - \theta_c)\psi(cx_0)/(1 - \psi(cx_0)) > 0$.

At the other, dense, endpoint of our range we have $p(n) \sim cn^{-1/r}$. Then t_c and a_c in (3.1) and (3.2) are of order constant. (Again the exact values are irrelevant.) This suggests, and the following theorem makes more precise, that the process will either die out or grow very quickly, with the outcome determined by the first few steps, and that the activation can spread from a set of constant size to the entire graph with a positive probability, which, however, is bounded away from 1.

THEOREM 5.6. *Suppose $r \geq 2$ and $p \sim cn^{-1/r}$ for a constant $c > 0$.*

(i) *If a is fixed with $a \geq r$, then*

$$(5.6) \quad \mathbb{P}(A^* = n) \rightarrow \zeta(a, c)$$

for some $\zeta(a, c) \in (0, 1)$. Furthermore, there exist numbers $\zeta(a, c, k) > 0$ for $k \geq a$ such that $\mathbb{P}(A^* = k) \rightarrow \zeta(a, c, k)$ for each fixed $k \geq a$, and $\sum_{k=a}^{\infty} \zeta(a, c, k) + \zeta(a, c) = 1$.

(ii) *If $a \rightarrow \infty$, then $\mathbb{P}(A^* = n) \rightarrow 1$, that is, $A^* = n$ w.h.p.*

REMARK 5.7. The limiting probabilities in Theorem 5.6 can be expressed as hitting probabilities of an inhomogeneous random walk. Let $\xi_k \in \text{Po}(\binom{k-1}{r-1}c^r)$, $k \geq 1$, be independent, and let $\tilde{S}_k := \sum_{j=1}^k (\xi_j - 1)$ and $\tilde{T} := \min\{k : a + \tilde{S}_k = 0\} \in \mathbb{N} \cup \{\infty\}$. Then

$$(5.7) \quad \zeta(a, c) = \mathbb{P}(\tilde{T} = \infty) = \mathbb{P}(a + \tilde{S}_k \geq 1 \text{ for all } k \geq 1)$$

and $\zeta(a, c, k) = \mathbb{P}(\tilde{T} = k)$. Consequently, Theorem 5.6(i) can also be written as

$$d_{\text{TV}}(A^*, \min(\tilde{T}, n)) \rightarrow 0,$$

where d_{TV} is the total variation distance.

If the probability of connections p is even larger, $p \gg n^{-1/r}$, then the initial set percolates as long as $a \geq r$.

THEOREM 5.8. *Let $r \geq 2$. If $p \gg n^{-1/r}$ and $a \geq r$, then $A^* = n$ w.h.p.*

REMARK 5.9. The case $r = 1$ is different. In this case, infection is equivalent to activation, and spreads to every vertex connected to an active vertex. Thus the final active set is the union of the components of the graph that contain at least one initially active vertex. It is well known that this is equivalent to the Reed–Frost model for epidemics, where each infected person infects everyone else with probability p , all infections being independent. (This equivalence is easily seen by the argument in Section 2.) The Reed–Frost model has been much studied; see, for example, von Bahr and Martin-Löf [48], Martin-Löf [37, 38]. We state some known result for comparison with our results for $r \geq 2$; proofs can be found in [48, 37], where also further details are given (including central limit theorems as in Section 3.2), or by modifying the proofs of the results above. Many results follow also easily from known results on the component structure of $G_{n,p}$.

If $p = \log n/n + \omega(n)/n$, with $\omega(n) \rightarrow \infty$, then w.h.p. $G_{n,p}$ is connected and thus $A^* = n$ as soon as $a \neq 0$. More generally, if $p \gg n^{-1}$ and $a \geq 1$, then w.h.p. $A^* = n - o(n)$; cf. Theorem 5.8.

The case $p = c/n$ is perhaps more interesting. There are many (w.h.p. $\geq c'n$) isolated vertices, so we cannot have percolation or almost percolation unless $a/n \rightarrow 1$. If $c > 1$, there is a single giant component of size $\rho n + o_p(n)$, with $\rho = \rho(c) > 0$, and thus, if $a \geq 1$ is fixed, then there is a dichotomy, with either $A^* = o(n)$ or $A^* = \rho n + o(n)$ w.h.p., with probabilities converging to the positive $(1 - \rho)^a$ and $1 - (1 - \rho)^a$, respectively; cf. Theorem 5.6.

If $c \leq 1$ and a is fixed, then, by the same argument, A^* converges to the total size of a Galton–Watson process with $\text{Po}(c)$ offspring distribution and a initial individuals (a Borel–Tanner distribution). Thus A^*/a is stochastically bounded but does not converge in probability to a constant; cf. Theorem 5.2(i).

If $c < 1$ and $a \rightarrow \infty$ but $a = o(n)$, then $A^*/a \xrightarrow{P} 1/(1 - c)$; cf. Theorem 5.2(i).

If $p \sim c/n$ with any $c > 0$ and $a \sim \theta n$ with $\theta > 0$, then $A^*/n \xrightarrow{P} x_0(\theta)$, where x_0 is the unique positive root of (5.4), where now $f(x, c, \theta) = 1 - x - (1 - \theta)e^{-cx}$ by (5.3) with $r = 1$. In other words, Theorem 5.2(iii) holds for $r = 1$, too, with $c_c = \infty$, and there is no threshold. [For $\theta = 0$ there is the well-known threshold at $c = 1$, but note that also in this case, x_0 is continuous in both c and θ and there is no jump as in Theorem 5.2(iv).]

6. Overview of the proofs. By (2.3), (2.2) and (2.9), for $u = 1, 2, \dots$

$$\begin{aligned}
 (6.1) \quad A^* \geq u &\iff T \geq u \iff \min_{t < u} (A(t) - t) > 0 \\
 &\iff a + \min_{t < u} (S(t) - t) > 0.
 \end{aligned}$$

Hence, $A^* = T$ is completely determined by the stochastic process $\min_{t < u} (S(t) - t)$, $u \geq 1$. In particular, $\mathcal{A}(0)$ percolates if and only if $a > -\min_{t < n} (S(t) - t) = \max_{t < n} (t - S(t))$.

Note that (6.1) is an exact representation of A^* ; we have not yet made any approximations. To obtain asymptotic results, we introduce some simple approximations. We give an informal overview of the argument here; details will follow in later sections.

First, $S(t) \approx \mathbb{E} S(t)$ by the law of large numbers. A simple calculation will show that $f(t) := \mathbb{E} S(t) - t$ starts at 0 for $t = 0$, then decreases to a minimum at $t \approx t_c$ given in (3.1), and then increases until $\mathbb{E} S(t) \approx n$ and thus $f(t) \approx n - t$; then $f(t) \approx n - t$ holds until $t = n$, so $f(t)$ decreases again in this range to a final value $f(n) = \mathbb{E} S(n) - n \approx 0$.

There are thus two candidates for the minimum point of $S(t) - t$: either $t \approx t_c$ or $t \approx n$. What happens at $t \approx n$ makes the difference between almost percolation and complete percolation; we will study this too in detail later, but for the moment we ignore it and concentrate on whether we have almost percolation or not, and we see that, roughly, this is determined by whether $a > -(S(t_c) - t_c)$ or not, which can be approximated by $a > -(\mathbb{E} S(t_c) - t_c)$. A simple calculation yields $\mathbb{E} S(t_c) - t_c \approx -a_c$, which establishes the threshold at a_c .

This argument also gives the following picture of the course of the activation $\mathcal{A}(t)$ in the critical case $a \approx a_c$. (We leave the modifications in the subcritical and supercritical cases to the reader.) We start with $A(0) = a$. At first, there are very few new vertices that reach the threshold of r infections, and the number $A(t) - t$ of unused vertices goes down, and approaches 0 as t approaches t_c . However, the rate of activation of new vertices is increasing, because a pool of vertices with $r - 1$ infections has accumulated, and as $t \rightarrow t_c$, new vertices are activated at about the same rate as they are used. There are now two possibilities: either the activation dies out at this point, with a total size about $t_c = r/(r - 1)a_c$, or it survives this bottleneck, and it then rapidly grows after time t_c until almost all vertices are active. In the latter case there are again two possibilities: either all remaining vertices are finally active (complete percolation), or a few are not.

7. Approximation of $S(t)$ by its mean. For notational convenience, we assume that $V_n \setminus \mathcal{A}(0) = \{1, \dots, n - a\}$. Note first that $(n - a)^{-1} S(t) = (n - a)^{-1} \sum_{i=1}^{n-a} \mathbf{1}\{Y_i \leq t\}$ is the empirical distribution function of $\{Y_i\}_1^{n-a}$. By the law of large numbers for the binomial distribution (2.10), for every $t = t(n)$, $S(t) = (n - a)\pi(t) + o_p(n)$. Moreover, by the Glivenko–Cantelli theorem ([35], Proposition 4.24), the following holds uniformly for all t :

LEMMA 7.1. $\sup_{t \geq 0} |S(t) - \mathbb{E} S(t)| = o_p(n)$.

PROOF. If $n - a \geq \sqrt{n}$, say, this is a weaker version of [35], Proposition 4.24. For smaller $n - a$, the result is trivial, since $0 \leq S(t) \leq n - a$. \square

For small t , the uniform error bound in Lemma 7.1 is not good enough. [It can be improved to $O_p(n^{1/2})$, see Lemma 7.3, but this too is too large for our purposes.] For each t , (2.13) gives a bound $O_p((n\pi(t))^{1/2})$. We extend this to a uniform bound for a range of t by a martingale argument. We begin by introducing a pair of well-known martingales for empirical distribution functions. (See [31], Lemma 2.1, for a continuous time version.)

LEMMA 7.2. *The stochastic process*

$$(7.1) \quad \frac{S(t) - \mathbb{E} S(t)}{1 - \pi(t)}, \quad t = 0, 1, \dots,$$

is a martingale, and

$$(7.2) \quad \frac{S(t) - \mathbb{E} S(t)}{\pi(t)}, \quad t = r, r + 1, \dots,$$

is a reverse martingale.

PROOF. Since $S(t)$ is a sum of $n - a$ i.i.d. processes $\mathbf{1}\{Y_i \leq t\}$, it suffices to treat each of these separately, that is, for the first part to show that, for each i ,

$$X(t) = X_i(t) := \frac{\mathbf{1}\{Y_i \leq t\} - \mathbb{P}(Y_i \leq t)}{1 - \mathbb{P}(Y_i \leq t)} = 1 - \frac{\mathbf{1}\{Y_i > t\}}{\mathbb{P}(Y_i > t)}$$

is a martingale. This is elementary: if $Y_i \leq t$, then $X(t) = X(t + 1) = 1$. If $Y_i > t$, then $X(t) = -\pi(t)/(1 - \pi(t))$ either jumps to $X(t + 1) = 1$ or decreases to $X(t + 1) = -\pi(t + 1)/(1 - \pi(t + 1))$, and the conditional probabilities of these events are $(\pi(t + 1) - \pi(t))/(1 - \pi(t))$ and $(1 - \pi(t + 1))/(1 - \pi(t))$, respectively, so a simple calculation yields $\mathbb{E}(X(t + 1) | Y_i > t) = -\pi(t)/(1 - \pi(t))$. [Alternatively, this follows from the case $X(t) = 1$ and the fact that $\mathbb{E} X(t + 1) = \mathbb{E} X(t)$.] Hence, $\mathbb{E}(X(t + 1) | X(1), \dots, X(t)) = X(t)$.

For the second part, we similarly find that $\tilde{X}(t) := \mathbf{1}\{Y_i \leq t\}/\pi(t)$ is a reverse martingale, that is, that $\mathbb{E}(\tilde{X}(t) | \tilde{X}(t + 1), \dots) = \tilde{X}(t + 1)$. \square

LEMMA 7.3. *For any t_0 ,*

$$(7.3) \quad \mathbb{E} \left(\sup_{t \leq t_0} |S(t) - \mathbb{E} S(t)| \right)^2 \leq 16n\pi(t_0),$$

$$(7.4) \quad \mathbb{E} \left(\sup_{t \geq t_0} |S(t) - \mathbb{E} S(t)| \right)^2 \leq 16n(1 - \pi(t_0)).$$

PROOF. Assume first $\pi(t_0) \leq 1/2$. Let $\xi(t) := (S(t) - \mathbb{E} S(t))/(1 - \pi(t))$. This is a martingale by Lemma 7.2, and Doob's inequality ([35], Proposition 7.16)

yields, using (2.13),

$$\begin{aligned}
 \mathbb{E}\left(\sup_{t \leq t_0} |S(t) - \mathbb{E} S(t)|\right)^2 &\leq \mathbb{E} \sup_{t \leq t_0} |\xi(t)|^2 \\
 (7.5) \qquad \qquad \qquad &\leq 4 \mathbb{E} |\xi(t_0)|^2 = 4 \frac{\text{Var } S(t_0)}{(1 - \pi(t_0))^2} \\
 &\leq 8n\pi(t_0),
 \end{aligned}$$

which proves (7.3) in this case. Similarly, if $\pi(t_0) \geq 1/2$, then we obtain, using the reverse martingale (7.2),

$$(7.6) \qquad \mathbb{E}\left(\sup_{t \geq t_0} |S(t) - \mathbb{E} S(t)|\right)^2 \leq 4 \frac{\text{Var } S(t_0)}{\pi(t_0)^2} \leq 8n(1 - \pi(t_0)).$$

Now, let t_1 be the largest integer such that $\pi(t_1) \leq 1/2$. We can apply (7.5) with $t_0 = t_1$ and (7.6) with $t_0 = t_1 + 1$, and thus

$$\begin{aligned}
 \mathbb{E}\left(\sup_{t \geq 0} |S(t) - \mathbb{E} S(t)|\right)^2 &\leq \mathbb{E}\left(\sup_{t \leq t_1} |S(t) - \mathbb{E} S(t)|\right)^2 \\
 &\quad + \mathbb{E}\left(\sup_{t \geq t_1+1} |S(t) - \mathbb{E} S(t)|\right)^2 \\
 &\leq 8n.
 \end{aligned}$$

This immediately implies (7.3) for $\pi(t_0) > 1/2$ and (7.4) for $\pi(t_0) < 1/2$. \square

8. Approximation of $\mathbb{E} S(t)$ and proofs of Theorems 3.1–3.4. For (real) $t > 0$ and $pt \leq 1$, say, by (2.11),

$$\begin{aligned}
 \pi(t) &= \sum_{j=r}^{\lfloor t \rfloor} \binom{\lfloor t \rfloor}{j} p^j (1-p)^{\lfloor t \rfloor - j} = \binom{\lfloor t \rfloor}{r} p^r (1 + O(pt)) \\
 (8.1) \qquad &= \frac{t^r p^r}{r!} (1 + O(pt + t^{-1}))
 \end{aligned}$$

[cf. (3.17)], and thus, by (2.12),

$$(8.2) \qquad \mathbb{E} S(t) - n \frac{t^r p^r}{r!} = O(nt^r p^r (pt + t^{-1} + a/n)).$$

It thus makes sense to approximate $f(t) := \mathbb{E} S(t) - t$ by $\bar{f}(t) := n(tp)^r / r! - t$. An elementary calculation shows that \bar{f} has, on $[0, \infty)$ a unique, global minimum at t_c given by (3.1), and that the minimum value is $\bar{f}(t_c) = -a_c$. We obtain, for example, the following estimate.

LEMMA 8.1. *Suppose that $r \geq 2$, $n^{-1} \ll p \ll n^{-1/r}$ and $a = o(n)$. Then*

$$\sup_{0 \leq x \leq 10r} \left| S(xt_c) - \frac{1}{r} x^r t_c \right| = o_p(t_c).$$

PROOF. First, (8.2) and (3.4) yield, recalling (3.5), uniformly for $x \leq 10r$,

$$\mathbb{E} S(xt_c) = nx^r \frac{t_c^r p^r}{r!} (1 + o(1/x)) = x^r \frac{t_c}{r} (1 + o(1/x)) = \frac{x^r}{r} t_c + o(t_c).$$

Further, Lemma 7.3 yields by (8.1) and (3.4),

$$\sup_{0 \leq x \leq 10r} |S(xt_c) - \mathbb{E} S(xt_c)|^2 = O_p(n\pi(10rt_c)) = O_p(nt_c^r p^r) = O_p(t_c) = o_p(t_c^2),$$

and the result follows. \square

We shall use Lemma 8.1 to prove now that in the subcritical case ($a \sim \alpha a_c$ with $\alpha < 1$) there exists $t < t_c$ such that w.h.p. $A(t) \leq t$ and then determine the precise value of $A(T) = T$.

PROOF OF THEOREM 3.1(i). The assumption on a may be written by (3.2),

$$(8.3) \quad a = (\alpha + o(1))a_c = (\alpha(1 - r^{-1}) + o(1))t_c.$$

Hence, (2.9) and Lemma 8.1, taking $x = 1$, yield

$$\begin{aligned} A(t_c) - t_c &= S(t_c) + a - t_c = t_c/r + o_p(t_c) + a - t_c \\ &= t_c(r^{-1} + \alpha(1 - r^{-1}) - 1 + o_p(1)). \end{aligned}$$

Since $\alpha(1 - r^{-1}) < 1 - r^{-1}$, w.h.p. $A(t_c) - t_c < 0$, and thus, by (2.2), $T < t_c$.

We apply Lemma 8.1 again, now taking $x = T/t_c$, and see that $S(T) = (T/t_c)^r t_c/r + o_p(t_c)$. Since $S(T) = A(T) - a = T - a$, we find, using (8.3), that

$$T - \alpha(1 - r^{-1})t_c = S(T) + o(t_c) = \left(\frac{T}{t_c}\right)^r \frac{t_c}{r} + o_p(t_c)$$

and thus

$$(8.4) \quad r \frac{T}{t_c} - (r - 1)\alpha = \left(\frac{T}{t_c}\right)^r + o_p(1).$$

Since the function $h(x) := rx - x^r$ is strictly increasing from 0 to $r - 1$ on $[0, 1]$, (8.4) implies (using the fact just shown that $T/t_c < 1$ w.h.p.) that $T/t_c \xrightarrow{P} y$, where y is the unique root in $[0, 1]$ of $h(y) = (r - 1)\alpha$, that is, $y = \varphi(\alpha)$ given by (3.11).

This proves the first assertion, and if $\alpha > 0$, the second follows. If $\alpha = 0$, then $a = o(t_c)$, and (8.2) implies, for every fixed $\lambda > 0$, $\mathbb{E} S(\lambda a) = O(na^r p^r) = o(ant_c^{r-1} p^r) = o(a)$. Hence, for every fixed $\lambda > 1$, $A(\lambda a) = S(\lambda a) + a = a + o_p(a)$, so w.h.p. $A(\lambda a) < \lambda a$, and thus $a \leq T < \lambda a$. Consequently, when $\alpha = 0$, $T/a \xrightarrow{P} 1$. \square

We turn to the proof of the supercritical case in Theorem 3.1. The following lemma shows that if the process of activation can escape the bottleneck at t_c , then

the process continues until (almost) percolation. The idea is to split the time interval $[3t_c, n]$ into different intervals. Then in the proof of Theorem 3.1(ii) and (iii), it remains to show that if a is supercritical, then $A(t) > t$ for $t < 3t_c$.

Let $b^* := b_c \omega(n)$, where $\omega(n) \rightarrow \infty$ slowly but is otherwise arbitrary.

LEMMA 8.2. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then, for any a , w.h.p. $A(t) > t$ for all $t \in [3t_c, n - b^*]$.*

PROOF. By (2.9), $A(t) = S_{n-a}(t) + a \geq S_n(t)$, so it suffices to show that $S_n(t) > t$ (or equivalently, to take $a = 0$). We separate the proof into a number of different cases for different ranges of t . We assume at some places, without further mention, that n is large enough.

Case 1: $t \in [3t_c, 8rt_c]$. By Lemma 8.1, w.h.p. for all such t ,

$$S_n(t) \geq \frac{1}{r} \left(\frac{t}{t_c}\right)^r t_c - t_c \geq \frac{3^{r-1}}{r} t - t_c \geq \frac{3}{2} t - t_c > t.$$

Case 2: $t \in [8rt_c, p^{-1}]$. Let $t_j := 2^j rt_c$, $j \geq 1$, and let $J := \min\{j \geq 1 : pt_j \geq 1\}$. For $t_c \leq t \leq p^{-1}$, using (3.1),

$$\begin{aligned} \pi(t) &\geq \binom{t}{r} p^r (1-p)^{t-r} = \frac{t^r}{r!} p^r e^{-tp} (1 + o(1)) \\ &\geq \frac{1}{3} \frac{t^r p^r}{r!} = \frac{1}{3} t \left(\frac{t}{t_c}\right)^{r-1} \frac{t_c^{r-1} p^r}{r!} = \frac{t}{3rn} \left(\frac{t}{t_c}\right)^{r-1}. \end{aligned}$$

Hence, for $3 \leq j \leq J - 1$, $\mathbb{E} S_n(t_j) = n\pi(t_j) \geq \frac{2^j}{3} t_j \geq \frac{8}{3} t_j$, and thus, using Chebyshev's inequality and (2.13),

$$\mathbb{P}(S_n(t_j) \leq 2t_j) \leq \mathbb{P}\left(S_n(t_j) \leq \frac{3}{4} \mathbb{E} S_n(t_j)\right) \leq \frac{\text{Var } S_n(t_j)}{((1/4) \mathbb{E} S_n(t_j))^2} \leq \frac{16}{n\pi(t_j)} \leq \frac{6}{t_j}.$$

Hence,

$$\begin{aligned} \mathbb{P}(S_n(t) \leq t \text{ for some } t \in [8rt_c, t_J]) &\leq \sum_{j=3}^{J-1} \mathbb{P}(S_n(t_j) \leq 2t_j) \\ &\leq \sum_{j=3}^{J-1} \frac{6}{t_j} < \frac{12}{t_3} < \frac{2}{rt_c} = o(1). \end{aligned}$$

Case 3: $t \in [p^{-1}, c_1 n]$ for a suitable small $c_1 > 0$. Let $t'_1 := \lceil p^{-1} \rceil$. Then

$$\pi(t'_1) = \mathbb{P}(\text{Bin}(t'_1, p) \geq r) = \mathbb{P}(\text{Po}(t'_1 p) \geq r) + O(p) \geq 2c_1$$

for some small c_1 . Hence w.h.p. $S_n(t'_1) > c_1 n$ and consequently $S_n(t) \geq S_n(t'_1) > c_1 n \geq t$.

Case 4: $t \in [c_1 n, n - p^{-1}]$. Let $t'_2 := \lfloor c_1 n \rfloor$ and $t'_3 := n - p^{-1}$. Then

$$\begin{aligned} 1 - \pi(t'_2) &= \mathbb{P}(\text{Bin}(t'_2, p) < r) \\ &= O((t'_2 p)^{r-1} e^{-t'_2 p}) = O((np)^{r-1} e^{-c_1 np}) \\ &= o((np)^{-1}). \end{aligned}$$

Thus, $\mathbb{E}(n - S_n(t'_2)) = n(1 - \pi(t'_2)) = o(p^{-1})$, and w.h.p., $n - S_n(t'_2) < p^{-1}$, that is, $S_n(t'_2) > n - p^{-1} = t'_3$.

Case 5: $t \in [n - p^{-1}, n - b^*]$. We have $t'_3 := n - p^{-1}$. Then

$$\begin{aligned} 1 - \pi(t'_3) &= \mathbb{P}(\text{Bin}(\lfloor t'_3 \rfloor, p) < r) \\ &= O((t'_3 p)^{r-1} e^{-t'_3 p}) = O((np)^{r-1} e^{-np}) \\ &= O(b_c/n). \end{aligned}$$

Hence, $\mathbb{E}(n - S_n(t'_3)) = n(1 - \pi(t'_3)) = O(b_c) = o(b^*)$, and thus w.h.p. $n - S_n(t'_3) < b^*$, that is, $S_n(t'_3) > n - b^*$. \square

REMARK 8.3. The proof shows that once we reach at least $1/p$ active vertices, the active set will w.h.p. grow to at least $n - b^*$ in at most 3 generations. (Hence, the size then is $n - O_p(b_c)$; see [33], Lemma 3.)

LEMMA 8.4.

$$\min_{x \geq 0} \left(\frac{x^r}{r} - x \right) = \frac{1}{r} - 1,$$

attained at $x = 1$ only.

PROOF. Elementary calculus. \square

PROOF OF THEOREM 3.1(ii) AND (iii). For $0 \leq t \leq 3t_c$, we may assume $a \leq 3t_c$ since otherwise $A(t) > t$ trivially. In this case, Lemmas 8.1 and 8.4 (with $x = t/t_c$) show that w.h.p., uniformly in $t \leq 3t_c$,

$$\begin{aligned} A(t) = a + S(t) &\geq (1 + \delta)(1 - r^{-1})t_c + \frac{1}{r} \left(\frac{t}{t_c} \right)^r t_c - o(t_c) \\ &\geq \delta(1 - r^{-1})t_c + \frac{t}{t_c} t_c - o(t_c) > t. \end{aligned}$$

This and Lemma 8.2 show that w.h.p. $A(t) > t$ for all $t \leq n - b^*$, and thus $A^* > n - b^*$.

Hence $n - A^* < b^* = b_c \omega(n)$ w.h.p., for any choice of $\omega(n) \rightarrow \infty$, which is equivalent to $n - A^* = O_p(b_c)$; see, for example, [33], Lemma 3. This proves (ii).

If $b_c \rightarrow 0$, we may choose $b^* = 1$; then w.h.p. $n - A^* < 1$, so $A^* = n$. Conversely, if $b_c \not\rightarrow 0$, then, at least for a subsequence, there exists with probability at least $c > 0$ a vertex with degree $\leq r - 1$, and with probability $1 - a/n$, this vertex will never be activated so $A^* < n$; see Remark 3.3. This proves (iii). \square

PROOF OF THEOREM 3.2. Choose $b^* := npb_c \gg b_c$. By (3.3), $b^*p = (np)^{r+1}e^{-np}/(r - 1)! \rightarrow 0$. Hence, $(n - b^*)p = np + o(1) \rightarrow \infty$ and

$$\begin{aligned} 1 - \pi(n - b^*) &= \mathbb{P}(\text{Bin}(\lfloor n - b^* \rfloor, p) \leq r - 1) \\ &\sim \frac{(n - b^*)^{r-1} p^{r-1}}{(r - 1)!} (1 - p)^{n-b^*} \\ &\sim 1 - \pi(n). \end{aligned}$$

Consequently [see (3.7)],

$$\begin{aligned} \mathbb{E}(A(n) - A(n - b^*)) &= \mathbb{E}(S(n) - S(n - b^*)) \\ (8.5) \qquad \qquad \qquad &\leq n(\pi(n) - \pi(n - b^*)) \\ &= o(n(1 - \pi(n))) = o(b'_c). \end{aligned}$$

By assumption and Lemma 8.2, w.h.p. $T > n - b^*$, and thus $A(n - b^*) \leq A(T) \leq A(n)$. Hence (8.5) implies

$$(8.6) \qquad \qquad \qquad A^* = T = A(T) = A(n) + o_p(b'_c).$$

Further,

$$(8.7) \qquad \qquad \qquad n - A(n) = n - a - S(n) \in \text{Bin}(n - a, 1 - \pi(n))$$

with mean $(n - a)(1 - \pi(n)) \sim b'_c$; see (3.7).

If $b_c \rightarrow \infty$, then $b'_c \sim b_c$; thus (8.7) implies $n - A(n) = b_c + o_p(b_c)$, and (8.6) yields (i).

In $b_c \rightarrow b < \infty$, then $b'_c = b_c + o(1) \rightarrow b$; thus (8.6) yields $A^* = A(n) + o_p(1)$, and hence (since the variables are integer valued) $A^* = A(n)$ w.h.p. Further, in this case (8.7) implies $n - A(n) \xrightarrow{d} \text{Po}(b)$, and (ii) and (iii) follow. \square

PROOF OF THEOREM 3.4. An easy consequence of Theorem 3.1. \square

We end this section with a proof of Theorem 4.4, where we start with a number of external infections (but no initially active vertices). As said in Section 4.2, we do this by a minor variation of our method. We include this proof to show the flexibility of the method, but we omit parts that are identical or almost identical to the proofs above.

PROOF OF THEOREM 4.4. In order to preserve independence between vertices, we consider the model with a Poisson number $W \in \text{Po}(\mu)$ of external infections (independent of everything else). Then each vertex i receives $W_i \in \text{Po}(\mu/n)$

external infections, and these random variables are independent. The analysis in Section 2 becomes slightly modified: the number of infections (marks) at time t now is $M_i^\mu(t) := W_i + M_i(t)$, so Y_i is replaced by $Y_i^\mu := \min\{t : M_i^\mu(t) \geq r\}$ and $S(t)$ is replaced by $S^\mu(t) := \sum_{i=1}^n \mathbf{1}\{Y_i^\mu \leq t\}$. We now have $A(t) = S^\mu(t)$, so $A^* = T = \min\{t \geq 0 : S^\mu(t) = t\}$.

We take $\mu = ynpa_c$ for a fixed $y > 0$ and claim that if $y < 1$, then w.h.p. $A^* < t_c/r$ and thus $J_0 > W$, while if $y > 1$, then w.h.p. $A^* = n - o_p(n)$ and thus $J_0 \leq W$. The result then follows by taking $y = 1 \pm \varepsilon/2$ for small $\varepsilon > 0$.

To prove these claims, we first note that $\mathbb{E} S^\mu(t) = n\mathbb{P}(M_i^\mu(t) \geq r)$ with, for such μ and $t = O(t_c)$,

$$\begin{aligned} \mathbb{P}(M_i^\mu(t) \geq r) &= \mathbb{P}(W_i + M_i(t) \geq r) \\ (8.8) \quad &= \sum_{j=0}^{r-1} \mathbb{P}(W_i = j) \mathbb{P}(M_i(t) \geq r - j) + \mathbb{P}(W_i \geq r) \\ &\sim \sum_{j=0}^r \frac{(\mu/n)^j}{j!} \cdot \frac{(tp)^{r-j}}{(r-j)!} = \frac{(tp + \mu/n)^r}{r!} = \frac{p^r (t + ya_c)^r}{r!}. \end{aligned}$$

We obtain as in Lemma 8.1, using versions of Lemmas 7.2 and 7.3 for $S^\mu(t)$,

$$(8.9) \quad \sup_{0 \leq x \leq 10r} \left| S^\mu(xt_c) - \frac{1}{r} (x + ya_c/t_c)^r t_c \right| = o_p(t_c).$$

Recall that $a_c/t_c = 1 - 1/r$ by (3.2). If $y < 1$, then (8.9) with $x = 1/r$ implies that w.h.p.

$$A\left(\frac{t_c}{r}\right) = S^\mu\left(\frac{t_c}{r}\right) < \frac{1}{r} \left(\frac{1}{r} + \left(1 - \frac{1}{r}\right)\right)^r t_c = \frac{t_c}{r}$$

and thus $A^* = T < t_c/r$ as claimed.

Conversely, if $y > 1$, then Lemma 8.4 shows that

$$\frac{(x + ya_c/t_c)^r}{r} \geq x + y \frac{a_c}{t_c} + \left(\frac{1}{r} - 1\right) = x + (y - 1) \left(1 - \frac{1}{r}\right).$$

Hence, (8.9) shows that w.h.p. $A(xt_c) = S^\mu(xt_c) > xt_c$ for $x \leq 10r$, and thus $A^* = T > 10rt_c$. Further, since $S^\mu(t) \geq S_n(t)$, Lemma 8.2 implies that w.h.p. $A(t) > t$ for all $t \in [3t_c, n - b^*]$, and thus $A^* \geq n - b^*$ w.h.p., which proves the second claim and completes the proof. \square

Note that (8.8) and (8.9) show that $A(t) = S^\mu(t)$ is, to the first order, $\mathbb{E} S_n(t)$ shifted horizontally by $\mu/(np) = ya_c$, while in our standard model $A(t)$ is S_t shifted vertically by a . Since we study the hitting time of the linear barrier $A(t) = t$, these are essentially equivalent.

9. Proofs of Theorems 3.6–3.8 and 4.5. We begin with an estimate of $\tilde{\pi}(t)$ defined in (3.13).

LEMMA 9.1. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then, for large n , $n\tilde{\pi}(t) \geq 1.4t$ for $t \in [3t_c, n/2]$.*

PROOF. Assume not. Then we can find, for a subsequence $n = n_k \rightarrow \infty$, $t = t_k \in [3t_c, n/2]$ such that $n\tilde{\pi}(t) < 1.4t$. Selecting a subsequence, we may further assume that $pt \rightarrow z \in [0, \infty]$. We consider three cases separately.

(i) $z = 0$, that is, $pt \rightarrow 0$. Then, from (3.13) and (3.4),

$$n\tilde{\pi}(t) \sim n \frac{(pt)^r}{r!} = \frac{t^r}{rt_c^{r-1}} \geq \frac{3^{r-1}}{r} t \geq \frac{3}{2} t.$$

(ii) $pt \rightarrow z \in (0, \infty)$. Then $n\tilde{\pi}(t) \sim n\psi(z)$ with $\psi(z) > 0$, and $t = O(1/p) = o(n) \ll n\tilde{\pi}(t)$.

(iii) $z = \infty$, that is, $pt \rightarrow \infty$. Then $n\tilde{\pi}(t) \sim n \geq 2t$.

In all cases we have for large n a contradiction to $n\tilde{\pi}(t) < 1.4t$. \square

LEMMA 9.2. *Suppose $r \geq 2$ and $a \rightarrow \infty$ with $a = o(n)$. Then p_c and p_c^* defined by (3.12) and (3.19) satisfy $p_c^* \sim p_c$. In particular, $n^{-1} \ll p_c^* \ll n^{-1/r}$.*

PROOF. Let $p = yp_c$ for some fixed $y > 0$, and define t_c and a_c by (3.1) and (3.2). Then $t_c = y^{-(r-1)/r} (r/(r-1))a$ and $a_c = ay^{-(r-1)/r}$. Further, $n^{-1} \ll p \ll n^{-1/r}$ and, by (3.5), $pt_c \rightarrow 0$ and $t_c = o(n)$. Hence, if $x = O(1)$, and $t = xt_c$, then $pt = o(1)$, $t = o(n)$ and, uniformly in bounded x , by (3.13) and (3.4),

$$\tilde{\pi}(t) = \frac{(pt)^r}{r!} + O((pt)^{r+1}) = \frac{x^r t_c}{nr} (1 + o(1)).$$

Hence, uniformly in $x \leq 3$,

$$(9.1) \quad (n-a)\tilde{\pi}(t) - t = \left(\frac{x^r}{r} + o(1) - x \right) t_c.$$

By Lemma 9.1, for large n , $(n-a)\tilde{\pi}(t) - t \geq 0$ for $t \in [3t_c, n/2]$, and thus, by (9.1) and Lemma 8.4,

$$(9.2) \quad \begin{aligned} \gamma(p) &= \inf_{t \leq 3t_c} \{(n-a)\tilde{\pi}(t) - t\} = \left(\inf_{x \leq 3} \left(\frac{x^r}{r} - x \right) + o(1) \right) t_c \\ &= \left(\frac{1}{r} - 1 + o(1) \right) t_c = -(1 + o(1)) a_c \\ &= -(y^{-(r-1)/r} + o(1)) a. \end{aligned}$$

Hence, if $y = 1 - \delta < 1$, then $y^{-(r-1)/r} > 1$ and thus, for large n , $\gamma(p) < -a$ so $p_c^* > p = (1 - \delta)p_c$. Conversely, if $y = 1 + \delta > 1$, then (9.2) yields, for large n , $\gamma(p) > -a$ so $p_c^* < p = (1 + \delta)p_c$.

Consequently, $p_c^*/p_c \rightarrow 1$. \square

We also need more precise estimates of $S(t)$. The following Gaussian process limit is fundamental. $D[0, B]$ denotes the space of right-continuous functions on $[0, B]$, with the Skorohod topology; see, for example, [16] (for $B = 1$; the general case is similar by a change of variables) or [35], Chapter 16.

LEMMA 9.3. *Suppose $r \geq 2$ and $a \rightarrow \infty$ with $a = o(n)$. Then*

$$(9.3) \quad Z(x) := \frac{S(xt_c) - \mathbb{E} S(xt_c)}{\sqrt{t_c}} \xrightarrow{d} W(x^r/r)$$

in $D[0, B]$ for any fixed B , where W is a standard Brownian motion.

The conclusion, convergence in $D[0, B]$ for every fixed B , can also be expressed as convergence in $D[0, \infty)$.

PROOF OF LEMMA 9.3. This is a result on convergence of empirical distribution functions (of $\{Y_i\}$); cf. [16], Theorem 16.4; we get here a Brownian motion instead of a Brownian bridge as in [16] because we consider for each B only a small initial part of the distribution of Y_i .

For every fixed $x > 0$, by (8.1) and (3.5), $\pi(xt_c) \sim (xt_c p)^r / r! \rightarrow 0$, and thus by (2.13) and (3.4)

$$\text{Var } S(xt_c) \sim n\pi(xt_c) \sim \frac{np^r x^r t_c^r}{r!} = \frac{x^r t_c}{r} \rightarrow \infty.$$

Hence (2.10) and the central limit theorem yield $Z(x) \xrightarrow{d} N(0, x^r/r)$ for every $x > 0$, which proves $Z(x) \xrightarrow{d} W(x^r/r)$ for each fixed x .

This is easily extended to finite-dimensional convergence: Suppose that $0 < x_1 < \dots < x_\ell$ are fixed, and let $I_{ij} := \mathbf{1}\{Y_i \in (x_{j-1}t_c, x_j t_c]\}$, with $x_0 = 0$. Thus, $S(x_j t_c) - S(x_{j-1} t_c) = \sum_{i=1}^{n-a} I_{ij}$. Then, for $1 \leq j \leq \ell$ and $k \neq j$,

$$\mathbb{E} I_{ij} = \pi(x_j t_c) - \pi(x_{j-1} t_c),$$

$$\text{Var } I_{ij} = \mathbb{E} I_{ij}(1 - \mathbb{E} I_{ij}) \sim \mathbb{E} I_{ij}$$

$$= \pi(x_j t_c) - \pi(x_{j-1} t_c) \sim \left(\frac{x_j^r}{r} - \frac{x_{j-1}^r}{r} \right) \frac{t_c}{n},$$

$$\text{Cov}(I_{ij}, I_{ik}) = -\mathbb{E} I_{ij} \mathbb{E} I_{ik} = O(\pi(x_\ell t_c)^2) = O((t_c/n)^2) = o(t_c/n).$$

Note that $(I_{ij})_{j=1}^\ell$, $i = 1, 2, \dots, n$, are i.i.d. random vectors. The multi-dimensional central limit theorem with (e.g.) the Lindeberg condition (which follows from

the one-dimensional version in, for example, [28], Theorem 7.2.4, or [35], Theorem 5.12, by the Cramér–Wold device) thus shows that $(Z(x_j) - Z(x_{j-1}))_{j=1}^\ell \xrightarrow{d} (V_j)_{j=1}^\ell$ with V_j jointly normal with $\mathbb{E} V_j = 0$, $\text{Var} V_j = x_j^r/r - x_{j-1}^r/r$ and $\text{Cov}(V_j, V_k) = 0$ for $j \neq k$. Hence, $(V_j)_{j=1}^\ell \stackrel{d}{=} (W(x_j^r/r) - W(x_{j-1}^r/r))_{j=1}^\ell$, and thus $(Z(x_j))_{j=1}^\ell \xrightarrow{d} (W(x_j^r/r))_{j=1}^\ell$.

To show (9.3), it thus remains to show tightness of $Z(x)$. We use [16], Theorem 15.6, with $\gamma = 2$ and $\alpha = 1$ (an alternative would be to instead use Aldous’s tightness criterion ([35], Theorem 16.11)); it thus suffices to prove that, for every x_1, x_2, x_3 with $0 \leq x_1 \leq x_2 \leq x_3 \leq B$, and some constant C depending on B but not on n or x_1, x_2, x_3 ,

$$(9.4) \quad \mathbb{E}\{|Z(x_2) - Z(x_1)|^2 | Z(x_3) - Z(x_2)|^2\} \leq C(x_3 - x_1)^2.$$

With the notation above and $I'_{ij} := I_{ij} - \mathbb{E} I_{ij}$, the left-hand side of (9.4) can be written

$$t_c^{-2} \mathbb{E} \sum_{i,j,k,l=1}^{n-a} I'_{i2} I'_{j2} I'_{k3} I'_{l3} = t_c^{-2} \sum_{i,j,k,l=1}^{n-a} \mathbb{E}(I'_{i2} I'_{j2} I'_{k3} I'_{l3}).$$

By independence, the only nonzero terms are those where i, j, k, l either coincide in two pairs, or all four indices coincide, and it follows easily that (for any i)

$$(9.5) \quad \mathbb{E}\{|Z(x_2) - Z(x_1)|^2 | Z(x_3) - Z(x_2)|^2\} \leq 3t_c^{-2}(n - a)^2 \mathbb{E} I_{i2} \mathbb{E} I_{i3}.$$

Further, since each Y_i is integer-valued, the left-hand side of (9.4) vanishes unless there is at least one integer in each of the intervals $(x_1 t_c, x_2 t_c]$ and $(x_2 t_c, x_3 t_c]$, which implies that $x_3 t_c - x_1 t_c > 1$, so we only have to consider this case. It follows from (2.7) that for $m \leq x_3 t_c \leq B t_c$,

$$\mathbb{P}(Y_i = m) \leq \frac{m^{r-1}}{(r-1)!} p^r \leq \frac{B^{r-1} t_c^{r-1} p^r}{(r-1)!} = \frac{B^{r-1}}{n}$$

and thus, assuming $x_3 t_c - x_1 t_c > 1$,

$$\begin{aligned} \mathbb{E} I_{i2} + \mathbb{E} I_{i3} &\leq (\lfloor x_3 t_c \rfloor - \lfloor x_1 t_c \rfloor) \frac{B^{r-1}}{n} \leq (x_3 t_c - x_1 t_c + 1) \frac{B^{r-1}}{n} \\ &\leq 2(x_3 t_c - x_1 t_c) \frac{B^{r-1}}{n}. \end{aligned}$$

Consequently, (9.5) yields,

$$\begin{aligned} \mathbb{E}\{|Z(x_2) - Z(x_1)|^2 | Z(x_3) - Z(x_2)|^2\} &\leq 3 \frac{n^2}{t_c^2} \left(2(x_3 t_c - x_1 t_c) \frac{B^{r-1}}{n} \right)^2 \\ &= 12(x_3 - x_1)^2 B^{2(r-1)}, \end{aligned}$$

which proves (9.4) with $C = 12B^{2(r-1)}$. The proof is complete. \square

We also need a more careful estimate of $\pi(t)$ than above, and we use the corresponding Poisson probability $\tilde{\pi}(t)$ defined in (3.13).

LEMMA 9.4. *Assume $n^{-1} \ll p \ll n^{-1/r}$. Uniformly for $t \geq 1$, $\pi(t) = \tilde{\pi}(t)(1 + O(t^{-1}))$. In particular, uniformly for $t \leq 3t_c$,*

$$\pi(t) = \tilde{\pi}(t) + O((pt)^r/t) = \tilde{\pi}(t) + O(\tilde{\pi}(t_c)/t_c) = \tilde{\pi}(t) + O(n^{-1}).$$

PROOF. Assume first $pt \leq 1$. By (2.11),

$$\begin{aligned} \pi(t) &= \sum_{j=r}^t \mathbb{P}(\text{Bin}(t, p) = j) = \sum_{j=r}^t \frac{t^j}{j!} \left(1 + O\left(\frac{j^2}{t}\right)\right) p^j (1-p)^{t+O(j)} \\ &= \sum_{j=r}^{\infty} \frac{(pt)^j}{j!} e^{-pt+O(tp^2)} (1 + O(j^2/t + jp)) \\ &= \tilde{\pi}(t)(1 + O(tp^2 + t^{-1} + p)) = \tilde{\pi}(t)(1 + O(t^{-1})). \end{aligned}$$

For $pt > 1$, $\tilde{\pi}(t)$ is bounded below, and the result follows from (3.15).

If $t \leq 3t_c$, then $t = O(t_c) = o(1/p)$ by (3.5), and thus, using (3.13) and (3.4),

$$\pi(t) - \tilde{\pi}(t) = O(\tilde{\pi}(t)/t) = O((pt)^r/t) = O((pt_c)^r/t_c) = O(1/n). \quad \square$$

PROOF OF THEOREM 3.6. It suffices to consider a such that $a \sim a_c = (1 - r^{-1})t_c$. It then follows by (2.9) and Lemma 8.1 that, uniformly for $x \leq 10r$,

$$\begin{aligned} (9.6) \quad A(xt_c) - xt_c &= a + S(xt_c) - xt_c = a_c + \frac{1}{r}x^r t_c - xt_c + o_p(t_c) \\ &= \left(1 - r^{-1} + \frac{1}{r}x^r - x\right)t_c + o_p(t_c). \end{aligned}$$

By Lemma 8.4, the coefficient $1 - r^{-1} + x^r/r - x$ equals 0 at $x = 1$ but is strictly positive for all other $x \geq 0$. It follows that for every $\delta > 0$, w.h.p. $A(xt_c) - xt_c > 0$ for all $x \in [0, 1 - \delta] \cup [1 + \delta, 10r]$. By a simple standard argument, there thus exists a sequence $\delta_n \rightarrow 0$, where we may further assume that $\delta_n > |t_c^*/t_c - 1|$, such that w.h.p. $A(xt_c) - xt_c > 0$ for all $x \in [0, 1 - \delta_n] \cup [1 + \delta_n, 10r]$.

Hence, w.h.p. either $T \in [(1 - \delta_n)t_c, (1 + \delta_n)t_c]$, or $A(t) > t$ for all $t \leq 10rt_c$; in the latter case, for any $b^* \gg b_c$, w.h.p. $A(t) > t$ for all $t \leq n - b^*$ by Lemma 8.2, so $T \geq n - b^*$; hence $T = n - O_p(b_c)$ and, more precisely, provided $a = o(n)$, Theorem 3.2 applies.

We thus only have to investigate the interval $[(1 - \delta_n)t_c, (1 + \delta_n)t_c]$ more closely. By the Skorohod coupling theorem ([35], Theorem 4.30), we may assume that the processes for different n are coupled such that the limit (9.3) holds a.s., and not just

in distribution. Since convergence in $D[0, B]$ to a continuous function is equivalent to uniform convergence, this means that (a.s.) $Z(x) \rightarrow W(x^r/r)$ uniformly for $x \leq B$; in particular, uniformly for $x \in [1 - \delta_n, 1 + \delta_n]$,

$$(9.7) \quad \begin{aligned} S(xt_c) &= (n - a)\pi(xt_c) + t_c^{1/2}Z(x) \\ &= (n - a)\pi(xt_c) + t_c^{1/2}(W(1/r) + o(1)). \end{aligned}$$

Let $\xi := W(1/r) \in N(0, 1/r)$. Then, by (9.7) and Lemma 9.4, uniformly for $x \in [1 - \delta_n, 1 + \delta_n]$,

$$(9.8) \quad \begin{aligned} S(xt_c) &= (n - a)\tilde{\pi}(xt_c) + O(1) + t_c^{1/2}(\xi + o(1)) \\ &= (n - a)\tilde{\pi}(xt_c) + t_c^{1/2}(\xi + o(1)) \end{aligned}$$

and thus, refining (9.6),

$$(9.9) \quad \begin{aligned} A(xt_c) - xt_c &= a + S(xt_c) - xt_c \\ &= a + (n - a)\tilde{\pi}(xt_c) - xt_c + t_c^{1/2}\xi + o_p(t_c^{1/2}). \end{aligned}$$

Hence, recalling (3.16) and that the minimum there is attained at $t_c^* \in [(1 - \delta_n)t_c, (1 + \delta_n)t_c]$,

$$(9.10) \quad \begin{aligned} &\min_{t \in [(1 - \delta_n)t_c, (1 + \delta_n)t_c]} \frac{A(t) - \lfloor t \rfloor}{1 - \tilde{\pi}(t)} \\ &= a + \min_{t \in [(1 - \delta_n)t_c, (1 + \delta_n)t_c]} \frac{n\tilde{\pi}(t) - t}{1 - \tilde{\pi}(t)} + t_c^{1/2}\xi + o_p(t_c^{1/2}) \\ &= a - a_c^* + t_c^{1/2}\xi + o_p(t_c^{1/2}). \end{aligned}$$

We have shown that w.h.p. $A^* = T \leq (1 + \delta_n)t_c$ if and only if this minimum is ≤ 0 , and otherwise $T = n - O_p(b_c)$, and the results follow; for (i) we also observe that (9.9) and (9.10) imply that w.h.p. $A(t_c^*) - t_c^* < 0$ and thus $T < t_c^*$. For example, in (iii) we have

$$\begin{aligned} a - a_c^* + t_c^{1/2}\xi + o_p(t_c^{1/2}) &= ya_c^{1/2} + t_c^{1/2}\xi + o_p(t_c^{1/2}) \\ &= ((r - 1)^{1/2}y + r^{1/2}\xi + o_p(1))(t_c/r)^{1/2}, \end{aligned}$$

and the probability that this is positive tends to

$$\mathbb{P}((r - 1)^{1/2}y + r^{1/2}\xi > 0) = \Phi((r - 1)^{1/2}y),$$

since $r^{1/2}\xi \in N(0, 1)$. \square

PROOF OF THEOREM 3.7. It suffices to consider $p \sim p_c^* \sim p_c$, which implies that $a_c = a_c(p) \sim a_c(p_c) = a$. Hence the arguments in the proof of Theorem 3.6 apply. In particular, again it suffices to consider $t \in J = J_n := [(1 - \delta_n)t_c, (1 +$

$\delta_n)t_c]$, where now $t_c = t_c(p_c) = (r/(r - 1))a$. The infimum in (3.18) is attained for some $t = t_c^{**}$, where by Lemma 9.1 $t_c^{**} \leq 3t_c$ for large n , and an argument as in (9.6) shows that $t_c^{**} \sim t_c$. We may assume that δ_n is chosen such that $t_c^{**} \in J$. Then, by (9.9),

$$(9.11) \quad \min_{t \in J} \{A(t) - \lfloor t \rfloor\} = a + \min_{t \in J} \{(n - a)\tilde{\pi}(t) - t\} + t_c^{1/2}\xi + o_p(t_c^{1/2}),$$

where, by (3.18) and the comments just made (for large n),

$$(9.12) \quad a + \min_{t \in J} \{(n - a)\tilde{\pi}(t) - t\} = a + \gamma(p) = -\gamma(p_c^*) + \gamma(p).$$

Further, writing (3.18) as $\gamma(p) := \min_t \{F(tp) - t\}$, with $F(x) = (n - a)\psi(x)$, we have at the minimum point $t := t_c^{**}$ the derivative $pF'(pt_c^{**}) - 1 = 0$. Hence, uniformly for $|p_1 - p_c^*| \leq \varepsilon p_c^*$ and $|t - t_c| \leq \varepsilon t_c$, for any $\varepsilon = \varepsilon_n \rightarrow 0$, using (3.14), $F'(p_1t) = (1 + o(1))F'(pt_c^{**}) = (1 + o(1))/p_c^*$ and thus by the mean-value theorem, for some p_1 between p and p_c^* ,

$$F(tp) - F(tp_c^*) = t(p - p_c^*)F'(tp_1) = t_c \frac{p - p_c^*}{p_c^*} (1 + o(1)).$$

Since the minimum in (3.18) may be taken over such t only, for suitable ε_n , this yields

$$\gamma(p) - \gamma(p_c^*) = t_c \frac{p - p_c^*}{p_c^*} (1 + o(1)).$$

Consequently, (9.11) and (9.12) yield

$$\min_{t \in J} \{A(t) - \lfloor t \rfloor\} = t_c \frac{p - p_c^*}{p_c^*} (1 + o(1)) + t_c^{1/2}\xi + o_p(t_c^{1/2}).$$

Hence,

$$\begin{aligned} \mathbb{P}\left(\min_{t \in J} \{A(t) - \lfloor t \rfloor\} > 0\right) &= \mathbb{P}\left(t_c^{1/2} \frac{p - p_c^*}{p_c^*} + \xi > 0\right) + o(1) \\ &= \mathbb{P}\left(-r^{1/2}\xi < (rt_c)^{1/2} \frac{p - p_c^*}{p_c^*}\right) + o(1), \end{aligned}$$

where $r^{1/2}\xi \in N(0, 1)$ and $t_c = \frac{r}{r-1}a$, and the different parts of the theorem follow. □

LEMMA 9.5. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then, for large n at least, the minimum point t_c^* in (3.16) is unique, and $t_c^* \sim t_c$, $a_c^* \sim a_c$; more precisely,*

$$(9.13) \quad t_c^* = \left(1 + \frac{pt_c}{r-1} + o(pt_c)\right)t_c,$$

$$(9.14) \quad a_c^* = \left(1 - \frac{1}{r} + \frac{pt_c}{r+1} + o(pt_c)\right)t_c = \left(1 + \frac{rpt_c}{r^2-1} + o(pt_c)\right)a_c.$$

PROOF. Let

$$(9.15) \quad g(t) := \frac{n\tilde{\pi}(t) - t}{1 - \tilde{\pi}(t)} = \frac{n - t}{1 - \tilde{\pi}(t)} - n.$$

Then

$$g'(t) = \frac{-(1 - \tilde{\pi}(t)) + (n - t)\tilde{\pi}'(t)}{(1 - \tilde{\pi}(t))^2} = \frac{(n - t)\tilde{\pi}'(t) + \tilde{\pi}(t) - 1}{(1 - \tilde{\pi}(t))^2}.$$

Let

$$(9.16) \quad h(t) := (1 - \tilde{\pi}(t))^2 g'(t) = (n - t)\tilde{\pi}'(t) + \tilde{\pi}(t) - 1.$$

Then $h'(t) = (n - t)\tilde{\pi}''(t) > 0$ for $t < (r - 1)/p$, and in particular (for large n) for $t \leq 3t_c$; see (3.5). Further, $h(0) = -1$ and by (3.14) and (3.1), for large n ,

$$h(3t_c) = 3^{r-1} - 1 + o(1) > 0;$$

hence, there is a unique $t_c^* \in [0, 3t_c]$ such that $h(t_c^*) = 0$, or equivalently $g'(t_c^*) = 0$. Further, $g''(t_c^*) = h'(t_c^*)/(1 - \tilde{\pi}(t_c^*))^2 > 0$ so t_c^* is the unique minimum point of $g(t)$ in $[0, 3t_c]$, as we defined t_c^* after (3.16).

Let $x = t_c^*/t_c \in [0, 3]$. Then, by (3.14), (3.1) and (3.17),

$$0 = h(t_c^*) = \left(1 - \frac{t_c^*}{n}\right) x^{r-1} e^{-pt_c^*} + O\left(\frac{t_c^*}{n}\right) - 1 = x^{r-1} e^{-pt_c^*} - 1 + O\left(\frac{t_c}{n}\right).$$

Hence, recalling $n^{-1} \ll p$ and $pt_c \rightarrow 0$,

$$(9.17) \quad x = e^{pt_c^*/(r-1)} \left(1 + O\left(\frac{t_c}{n}\right)\right) = 1 + \frac{pt_c^*}{r-1} + o(pt_c).$$

In particular, $x = 1 + o(1)$, so $t_c^* \sim t_c$, and (9.13) follows from (9.17). Finally, substituting (9.13) in (3.16) yields, using (9.15) together with $\tilde{\pi}(t_c^*) = O(t_c/n) = o(pt_c)$ by (3.17) and $n^{-1} \ll p$, and also (3.13) and (3.17),

$$\begin{aligned} \frac{a_c^*}{t_c} &= -\frac{g(t_c^*)}{t_c} = (1 + o(pt_c)) \frac{t_c^* - n\tilde{\pi}(t_c^*)}{t_c} = (1 + o(pt_c)) \left(x - \frac{n\tilde{\pi}(t_c^*)}{t_c}\right) \\ &= (1 + o(pt_c)) \left(x - \frac{n(pt_c^*)^r}{r!t_c} e^{-pt_c^*} \left(1 + \frac{pt_c^*}{r+1} + o(pt_c^*)\right)\right) \\ &= (1 + o(pt_c)) \left(x - \frac{x^r}{r} e^{-pt_c} \left(1 + \frac{pt_c}{r+1} + o(pt_c)\right)\right) \\ &= x - \frac{x^r}{r} \left(1 - pt_c + \frac{pt_c}{r+1}\right) + o(pt_c) \\ &= x - \frac{x^r}{r} + \frac{1}{r} \left(pt_c - \frac{pt_c}{r+1}\right) + o(pt_c), \end{aligned}$$

and (9.14) follows by (9.17) and $x - x^r/r = 1 - 1/r + O(x - 1)^2$. \square

PROOF OF THEOREM 3.8. In case (i), that is, when $\alpha < 1$, by Theorem 3.1(i),

$$(9.18) \quad T = A^* = (\varphi(\alpha) + o_p(1))t_c.$$

By Theorem 3.6(i), this holds as well in case (ii), that is, when $\alpha = 1$ and, correspondingly, $\varphi(\alpha) = \varphi(1) = 1$. Thus, for any $\alpha \leq 1$ there exist $\delta_n \rightarrow 0$ such that w.h.p. $T \in I_n := [(\varphi(\alpha) - \delta_n)t_c, (\varphi(\alpha) + \delta_n)t_c]$. As in the proof of Theorem 3.6, we may, by the Skorohod coupling theorem ([35], Theorem 4.30) assume that the limit in (9.3) holds a.s., uniformly in $x \leq B$. For $t \in I_n$, $t/t_c \rightarrow \varphi(\alpha)$, and (9.3) then implies that, uniformly for $t \in I_n$,

$$(9.19) \quad \begin{aligned} S(t) &= \mathbb{E} S(t) + t_c^{1/2} Z(t/t_c) \\ &= (n-a)\pi(t) + t_c^{1/2} W(\varphi(\alpha)^r/r) + o_p(t_c^{1/2}). \end{aligned}$$

Let $\xi := W(\varphi(\alpha)^r/r) \in N(0, \varphi(\alpha)^r/r)$. Then, by (9.19) and Lemma 9.4, for $t \in I_n$,

$$S(t) = (n-a)\tilde{\pi}(t) + t_c^{1/2}(\xi + o_p(1)).$$

Since w.h.p. $T \in I_n$, we may here substitute $t = T$, and obtain

$$(9.20) \quad \begin{aligned} 0 &= A(T) - T = a + S(T) - T \\ &= a + (n-a)\tilde{\pi}(T) - T + t_c^{1/2}(\xi + o_p(1)). \end{aligned}$$

Define the function $\tilde{g}(t)$ by

$$(9.21) \quad \tilde{g}(t) := a + (n-a)\tilde{\pi}(t) - t;$$

thus (3.22) is $\tilde{g}(t_*) = 0$. Then we have shown in (9.20),

$$(9.22) \quad \tilde{g}(T) = -t_c^{1/2}(\xi + o_p(1)).$$

The function \tilde{g} is continuous on $[0, \infty)$ with $\tilde{g}(0) = a > 0$. Consider the two cases separately.

(i): When $\alpha < 1$ we have, by (9.21) and (3.17), $\tilde{g}(t_c) = a + (1 + o(1))t_c/r - t_c = a - (1 + o(1))a_c < 0$ (for large n), since $a \sim \alpha a_c$. Further, on $[0, t_c]$, using (3.14) and (3.1),

$$(9.23) \quad \begin{aligned} \tilde{g}'(t) &= (n-a)\tilde{\pi}'(t) - 1 = \frac{n-a}{n} \left(\frac{t}{t_c}\right)^{r-1} e^{-pt} - 1 \\ &= \left(\frac{t}{t_c}\right)^{r-1} - 1 + o(1); \end{aligned}$$

this is negative for $t < (1 - \varepsilon)t_c$ for any $\varepsilon > 0$ and large n , and it follows that (for large n , at least), \tilde{g} has a unique root t_* in $[0, t_c]$. It follows from (3.17) and (3.11) that $t_*/t_c \rightarrow \varphi(\alpha)$.

Since also $T/t_c \xrightarrow{P} \varphi(\alpha)$, (9.23) implies that $\tilde{g}'(t) = -(1 - \varphi(\alpha)^{r-1}) + o_p(1)$ for all t between t_* and T , and thus the mean value theorem yields

$$\tilde{g}(T) = \tilde{g}(T) - \tilde{g}(t_*) = (T - t_*)(-(1 - \varphi(\alpha)^{r-1}) + o_p(1)),$$

which together with (9.22) yields, recalling $\varphi(\alpha) < 1$,

$$T - t_* = -((1 - \varphi(\alpha)^{r-1})^{-1} + o_p(1))\tilde{g}(T) = ((1 - \varphi(\alpha)^{r-1})^{-1}\xi + o_p(1))t_c^{1/2}.$$

The result in (i) follows.

(ii): Let $g(t) := \tilde{g}(t)/(1 - \tilde{\pi}(t)) - a$ and $h(t)$ be as in the proof of Lemma 9.5, (9.15) and (9.16). We know that $g(t_c^*) = -a_c^*$ and $g'(t_c^*) = 0$. Further, for $t \sim t_c$, we have by (3.5), (3.13), (3.17) and (3.14),

$$\begin{aligned} \tilde{\pi}(t) &\sim \frac{(tp)^r}{r!} \sim \frac{t_c}{rn} = o(1), \\ \tilde{\pi}'(t) &\sim \frac{r}{t}\tilde{\pi}(t) \sim \frac{1}{n}, \\ \tilde{\pi}''(t) &\sim \frac{r-1}{t}\tilde{\pi}'(t) \sim \frac{r-1}{nt_c}. \end{aligned}$$

Hence, by (9.16), $h(t) = o(1)$, $h'(t) = (n-t)\pi''(t) \sim (r-1)/t_c$ and

$$\begin{aligned} (9.24) \quad g''(t) &= \frac{h'(t)}{(1 - \tilde{\pi}(t))^2} + 2\frac{h(t)\tilde{\pi}'(t)}{(1 - \tilde{\pi}(t))^3} \\ &= \frac{r-1}{t_c}(1 + o(1)) + o\left(\frac{1}{n}\right) = \frac{r-1}{t_c}(1 + o(1)). \end{aligned}$$

Consequently, a Taylor expansion yields, for $t \sim t_c \sim t_c^*$,

$$(9.25) \quad g(t) = -a_c^* + \frac{r-1}{2t_c}(t - t_c^*)^2(1 + o(1)).$$

We have $\tilde{g}(t_*) = 0$ and thus $g(t_*) = -a$. Further, (3.17) and (3.11) again yield $t_*/t_c \rightarrow \varphi(1) = 1$. Hence, (9.25) yields (3.23).

Since Theorem 3.6 yields $T = t_c(1 + o_p(1))$ and $T < t_c^*$ w.h.p., (9.22) yields $g(T) = \tilde{g}(T)/(1 - \tilde{\pi}(T)) - a = -a - t_c^{1/2}(\xi + o_p(1))$; thus, similarly, (9.25) yields, using $a_c^* - a \gg t_c^{1/2}$,

$$\begin{aligned} t_c^* - T &= (1 + o_p(1))\sqrt{\frac{2t_c}{r-1}(a_c^* - a - t_c^{1/2}(\xi + o_p(1)))} \\ &= (1 + o_p(1))\sqrt{\frac{2t_c}{r-1}(a_c^* - a)} = (1 + o_p(1))(t_c^* - t_*). \end{aligned}$$

Hence, w.h.p., every t between T and t_* satisfies $t_c^* - t = (1 + o(1))(t_c^* - t_*)$, and then by (9.24),

$$g'(t) = (1 + o(1)) \frac{r-1}{t_c} (t - t_c^*) = -(1 + o(1)) \frac{r-1}{t_c} (t_c^* - t_*).$$

Finally, the mean value theorem yields, similarly to case (i),

$$T - t_* = \frac{g(T) - g(t_*)}{-(1 + o(1))(r-1)/t_c(t_c^* - t_*)} = \frac{t_c^{1/2}(\xi + o_p(1))}{(2(r-1)t_c^{-1}(a_c^* - a))^{1/2}},$$

and the result in (ii) follows, since $\xi \in N(0, 1/r)$ and $a_c^* \sim a_c = \frac{r-1}{r}t_c$. \square

PROOF OF THEOREM 4.5. We use the version described in Section 4.3 where edges are added at random times. Let \widehat{U} be the time the active set becomes big, that is, the time the M th edge is added. For any given p , then $\widehat{U} \leq p$ if and only if at time p , the active set is big, which is the same as saying that there is a big active set in $G_{n,p}$. Fix $x \in (-\infty, \infty)$ and choose $p = p_c^* + (r-1)^{1/2}r^{-1}xa^{-1/2}p_c$. Then Theorem 3.7 [with $\lambda = (r-1)^{1/2}r^{-1}x$] yields $\mathbb{P}(\widehat{U} \leq p) \rightarrow \Phi(x)$. In other words, $(\widehat{U} - p_c^*)/((r-1)^{1/2}r^{-1}a^{-1/2}p_c) \xrightarrow{d} N(0, 1)$, or

$$(9.26) \quad \widehat{U} \in \text{AsN}(p_c^*, (r-1)r^{-2}a^{-1}p_c^2).$$

Let $N(u)$ be the number of edges at time u . Then $N(0) = 0$ and, in analogy with Lemma 7.2, $(N(u) - \binom{n}{2}u)/(1-u)$ is a martingale on $[0, 1)$. Thus, Doob's inequality yields, as in the proof of Lemma 7.3, for any $u_0 \in [0, 1]$, $\mathbb{E}(\sup_{u \leq u_0} |N(u) - \binom{n}{2}u|^2) \leq 16\binom{n}{2}u_0 = O(n^2u_0)$; cf. [31], Lemma 3.2. Hence,

$$\sup_{u \leq u_0} \left| N(u) - \binom{n}{2}u \right| = O_p(nu_0^{1/2}).$$

Choosing $u_0 = 2p_c$, we thus obtain, since $\widehat{U} \leq 2p_c$ w.h.p.,

$$(9.27) \quad M = N(\widehat{U}) = \binom{n}{2}\widehat{U} + O_p(np_c^{1/2}).$$

We have by (3.12), for some constant $c = c(r)$,

$$\frac{np_c^{1/2}}{n^2p_c/a^{1/2}} = \frac{a^{1/2}}{np_c^{1/2}} = \frac{a^{1/2}c(na^{r-1})^{1/(2r)}}{n} = c\left(\frac{a}{n}\right)^{1-1/2r} = o(1).$$

Consequently, the error term in (9.27) is $o_p(n^2p_c/a^{1/2})$, and the result follows from (9.27) and (9.26). \square

10. The number of generations. Let $T_0 := 0$ and define inductively

$$(10.1) \quad T_{j+1} := A(T_j), \quad j \geq 0.$$

Thus $A(T_0) = A(0) = |\mathcal{A}(0)| = |\mathcal{G}_0|$, the size of generation 0 (the initially active vertices). Further, by our choice of u_t as one of the oldest unused, active vertices, $\mathcal{Z}(T_1) = \mathcal{Z}(A(0)) = \mathcal{A}(0) = \mathcal{G}_0$ and $\mathcal{Z}(T_2) = \mathcal{Z}(A(T_1)) = \mathcal{A}(T_1) = \mathcal{G}_0 \cup \mathcal{G}_1$; in general, by induction, all vertices in generation k (and earlier) have been found and declared active at time T_k , and they have been used at time $T_{k+1} = A(T_k)$. In other words,

$$\bigcup_{j=0}^k \mathcal{G}_j = \mathcal{A}(T_k) = \mathcal{Z}(T_{k+1}), \quad k \geq 0.$$

In particular, the size of generation k equals

$$|\mathcal{G}_k| = |\mathcal{Z}(T_{k+1}) \setminus \mathcal{Z}(T_k)| = T_{k+1} - T_k, \quad k \geq 0,$$

and the number of generations τ defined by (1.1) is

$$\tau = \max\{k \geq 0 : T_{k+1} > T_k\} = \min\{k \geq 1 : T_{k+1} = T_k\} - 1.$$

We begin by considering the supercritical case. We then consider the spread of activation in the bootstrap percolation process in three different stages in each of the following subsections. We first consider the bottleneck when the size is close to t_c ; we know that this is where the activation will stop in the critical case, and in the slightly supercritical case, the activation will grow slowly here, and this will dominate the total time. Then follows a period of doubly exponential growth, and finally, when there are only $o(n)$ vertices remaining, it may take some time to sweep up the last of them. Recall that Example 3.11 shows that each of the three phases may dominate the two others.

We define, for any $m \leq n$,

$$(10.2) \quad \tau(m) := \inf\{j : T_j \geq m\}$$

with the interpretation that $\tau(m) = \infty$ if this set of j is empty, that is, if $m > A^* = T$.

10.1. *The bottleneck.* We consider first $\tau(3t_c)$, that is, the number of generations required to achieve at least $3t_c$ active vertices. [The constant 3 is chosen for convenience; any constant > 1 would give the same result within $O(1)$ w.h.p.] In the really supercritical case, this is achieved quickly.

PROPOSITION 10.1. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Assume $a \geq (1 + \delta)a_c$ for some $\delta > 0$. Then, w.h.p. $\tau(3t_c) = O(1)$.*

PROOF. Lemmas 8.1 and 8.4 imply that uniformly for $0 \leq t \leq 3t_c$, with $x = t/t_c$,

$$A(t) - t = S(t) - t + a = \left(\frac{1}{r}x^r - x\right)t_c + a + o_p(t_c) \geq -a_c + (1 + \delta)a_c + o_p(a_c)$$

and thus w.h.p.

$$A(t) - t \geq \frac{\delta}{2}a_c \geq \frac{\delta}{4}t_c.$$

Hence, in this range, w.h.p. each generation has size at least $(\delta/4)t_c$, and the numbers of generations $\tau(3t_c)$ required to reach $3t_c$ is thus w.h.p. bounded by $12/\delta$. □

In the slightly supercritical case when $a \sim a_c$, this part may be a real bottleneck, however. We will approximate $A(t)$ by deterministic functions and begin with a definition: given a function $F : [0, \infty) \rightarrow [0, \infty)$, define the iterates $T_{j+1}^F := F(T_j^F)$ with $T_0^F := 0$. Thus $T_j = T_j^A$.

LEMMA 10.2. *If $A \leq F$, then $T_j \leq T_j^F$ for every j . If $A \geq F$, then $T_j \geq T_j^F$ for every j .*

PROOF. By induction. Assume, for example, $A \leq F$ and $T_j \leq T_j^F$. Then, since A is (weakly) increasing,

$$T_{j+1} = A(T_j) \leq A(T_j^F) \leq F(T_j^F) = T_{j+1}^F. \quad \square$$

We next prove a deterministic lemma.

LEMMA 10.3. *Let $a, b, t_0 > 0$, and let $F(t) := t + a + b(t - t_0)^2$. Assume $a \leq t_0$ and $bt_0 \leq 1$. Let N be the smallest integer such that $T_N^F > 2t_0$. Then*

$$N = (1 + O(bt_0)) \int_{-t_0}^{t_0} \frac{1}{a + bx^2} dx + O(1).$$

PROOF. Assume that $t \in [0, 2t_0]$ and let $\Delta := F(t) - t$. The assumptions on a and b imply $0 < \Delta \leq a + bt_0^2 \leq 2t_0$. For $s \in [t, t + \Delta]$ we have $|F'(s) - 1| = |2b(s - t_0)| \leq 6bt_0$, and thus, by the mean-value theorem, $|F(s) - s - (F(t) - t)| \leq 6bt_0\Delta = 6bt_0(F(t) - t)$. Thus, uniformly for such s , $F(s) - s = (F(t) - t)(1 + O(bt_0))$ and thus $(F(t) - t)^{-1} = (F(s) - s)^{-1}(1 + O(bt_0))$. Consequently,

$$\begin{aligned} 1 &= \int_t^{t+\Delta} \frac{1}{F(t) - t} ds = (1 + O(bt_0)) \int_t^{t+\Delta} \frac{1}{F(s) - s} ds \\ &= (1 + O(bt_0)) \int_t^{t+\Delta} \frac{1}{a + b(s - t_0)^2} ds. \end{aligned}$$

If $t = T_j^F$, then $t + \Delta = F(t) = T_{j+1}^F$. Summing for $j = 0, \dots, N - 1$ we thus obtain

$$N = (1 + O(bt_0)) \int_0^{T_N^F} \frac{1}{a + b(s - t_0)^2} ds \geq (1 + O(bt_0)) \int_0^{2t_0} \frac{1}{a + b(s - t_0)^2} ds,$$

and similarly, omitting $j = N - 1$,

$$N - 1 \leq (1 + O(bt_0)) \int_0^{2t_0} \frac{1}{a + b(s - t_0)^2} ds.$$

The result follows, using the change of variable $s = x + t_0$. \square

PROPOSITION 10.4. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Assume $a/a_c \rightarrow 1$ and $a - a_c^* \gg \sqrt{a_c}$. Then,*

$$\tau(3t_c) = \frac{\pi\sqrt{2} + o_p(1)}{\sqrt{r-1}} \left(\frac{t_c}{a - a_c^*} \right)^{1/2}.$$

PROOF. By (8.1) and (3.4), $n\pi(3t_c) = O(t_c)$. Hence, by Lemmas 7.3 and 9.4, for $t \leq 3t_c$,

$$(10.3) \quad S(t) = \mathbb{E} S(t) + O_p(t_c^{1/2}) = (n - a)\tilde{\pi}(t) + O_p(t_c^{1/2}).$$

Let $H(t) := a + (n - a)\tilde{\pi}(t) - t$ and define $h := \inf_{t \leq 3t_c} H(t)$. Let the infimum be attained at t_* ; it follows from (3.17) and Lemma 8.4 that $t_* \sim t_c$; cf. (9.2). We have $H(t_*) = h$, $H'(t_*) = 0$ and, uniformly for $t \leq 3t_c$, using (3.14), (3.5) and (3.4),

$$\begin{aligned} H''(t) &= (n - a)\tilde{\pi}''(t) = (n - a)p^r \frac{t^{r-1}}{(r - 1)!} \left(\frac{r - 1}{t} - p \right) e^{-pt} \\ &= np^r \frac{t^{r-2}}{(r - 1)!} (r - 1 + o(1)) = \left(\frac{t}{t_c} \right)^{r-2} \frac{r - 1 + o(1)}{t_c} \\ &= \frac{r - 1}{t_c} \left(1 + o(1) + O\left(\frac{|t - t_c|}{t_c} \right) \right). \end{aligned}$$

Hence, by a Taylor expansion, for $0 \leq t \leq 3t_c$,

$$(10.4) \quad H(t) = h + \frac{r - 1}{2t_c} (t - t_*)^2 \left(1 + o(1) + O\left(\frac{|t - t_c|}{t_c} \right) \right).$$

Notice that in the last two formulas, the term $o(1)$ tends to 0 as $n \rightarrow \infty$, uniformly in $t \leq 3t_c$, and $O(\dots)$ is uniform in n ; these uniformities allow us to combine the two terms in a meaningful way.

On the interval $[0, 3t_c]$, $\tilde{\pi}(t) = o(1)$ by (3.17) and (3.5), and thus by (3.16)

$$\begin{aligned}
 (10.5) \quad h &\sim \inf_{t \leq 3t_c} \frac{H(t)}{1 - \tilde{\pi}(t)} = \inf_{t \leq 3t_c} \frac{a + (n - a)\tilde{\pi}(t) - t}{1 - \tilde{\pi}(t)} \\
 &= a + \inf_{t \leq 3t_c} \frac{n\tilde{\pi}(t) - t}{1 - \tilde{\pi}(t)} = a - a_c^*.
 \end{aligned}$$

In particular, by our assumption, $h \gg a_c^{1/2}$. Consequently, by (10.3) and (10.4), for any fixed small $\varepsilon > 0$ and $|t - t_c| \leq 2\varepsilon t_c$, w.h.p.

$$\begin{aligned}
 (10.6) \quad A(t) - t &= a + S(t) - t = H(t) + o_p(h) \\
 &= (1 + O(\varepsilon)) \left(h + \frac{r - 1}{2t_c} (t - t_*)^2 \right).
 \end{aligned}$$

Let $t_1 := (1 - \varepsilon)t_* \sim (1 - \varepsilon)t_c$ and $t_2 := (1 + \varepsilon)t_* \sim (1 + \varepsilon)t_c$. For $0 \leq t \leq t_1$ and $t_2 \leq t \leq 3t_c$, Lemmas 8.1 and 8.4 imply that w.h.p. $A(t) - t \geq ct_c$, for some constant $c = c(\varepsilon) > 0$. The numbers of generations required to cover the intervals $[0, t_1]$ and $[t_2, 3t_c]$ are thus $O(1/c(\varepsilon))$, so $\tau(3t_c) = \tau'_\varepsilon + O(1/c(\varepsilon))$, where τ'_ε is the number of generations needed to increase the size from at least t_1 to at least t_2 . To find τ'_ε , we may redefine T_n by starting with $T_0 := t_1$ and iterate as in (10.1) until we reach t_2 . (Note that since A is increasing, if we start with a larger T_0 , then every T_n will be larger. Hence, to start with exactly t_1 can only affect τ'_ε by at most 1.) By (10.6) and Lemma 10.2, we may on the interval $[t_1, t_2]$ w.h.p. obtain upper and lower bounds from $F_\pm(t) = t + (1 \pm C\varepsilon)(h + b(t - t_*)^2)$, where $b := (r - 1)/(2t_c) > 0$ and C is some constant. Let $t_0 := t_* - t_1 = \varepsilon t_* > 0$. We have $a_c^* \sim a_c$ and by assumption $a \sim a_c$, so by (10.5), $h = o(a_c) = o(t_c)$ and thus $h < t_0/2$ for large n . Furthermore, $bt_0 = O(\varepsilon t_*/t_c) = O(\varepsilon)$. If ε is small enough, we thus have $bt_0 \leq 1/2$ and, by a translation $t \mapsto t - t_1$, Lemma 10.3 applies to both F_+ and F_- and yields, w.h.p., using (10.5),

$$\begin{aligned}
 \tau'_\varepsilon &= (1 + O(\varepsilon)) \int_{-\varepsilon t_*}^{\varepsilon t_*} \frac{dx}{h + bx^2} + O(1) \\
 &= (1 + O(\varepsilon)) \int_{-\infty}^{\infty} \frac{dx}{h + bx^2} + O\left(\frac{1}{b\varepsilon t_*}\right) + O(1) \\
 &= (1 + O(\varepsilon)) \int_{-\infty}^{\infty} \frac{dx}{h + bx^2} + O\left(\frac{1}{\varepsilon}\right) \\
 &= (1 + O(\varepsilon)) \frac{\pi}{(hb)^{1/2}} + O\left(\frac{1}{\varepsilon}\right) \\
 &= (1 + O(\varepsilon)) \left(\frac{2t_c}{(r - 1)(a - a_c^*)} \right)^{1/2} \pi + O(1/\varepsilon).
 \end{aligned}$$

Since $t_c/(a - a_c^*) \rightarrow \infty$, it follows that for every $\varepsilon > 0$, w.h.p., with $c'(\varepsilon) := \min(c(\varepsilon), \varepsilon) > 0$,

$$\begin{aligned} \tau(3t_c) &= \tau'_\varepsilon + O(1/c(\varepsilon)) = (1 + O(\varepsilon)) \frac{\pi\sqrt{2}}{\sqrt{r-1}} \left(\frac{t_c}{a - a_c^*}\right)^{1/2} + O(1/c'(\varepsilon)) \\ &= \frac{\pi\sqrt{2} + O(\varepsilon)}{\sqrt{r-1}} \left(\frac{t_c}{a - a_c^*}\right)^{1/2}. \end{aligned}$$

The result follows since $\varepsilon > 0$ is arbitrary. \square

REMARK 10.5. In the critical case $(a - a_c^*)/\sqrt{a_c} \rightarrow y \in (-\infty, \infty)$, we can use a minor variation of the same argument, now using Lemma 9.3, where $h \sim a - a_c$ above is replaced by the random

$$h' = a - a_c + t_c^{1/2}W(1/r) + o_p(t_c^{1/2}) = r^{-1/2}t_c^{1/2}(y\sqrt{r-1} + \xi + o_p(1)),$$

where $\xi \sim N(0, 1)$. We have $\tau(3t_c) < \infty \iff A^* \geq 3t_c \iff h' > 0$; this is w.h.p. equivalent to $y\sqrt{r-1} + \xi > 0$. [This thus happens with probability $\Phi((r-1)^{1/2}y) + o(1)$, as stated in Theorem 3.6(iii).] The argument above then shows that conditioned on $\tau(3t_c) < \infty$ (i.e., on $A^* \geq 3t_c$),

$$\tau(3t_c)/t_c^{1/4} \xrightarrow{d} \left(\frac{2^{1/2}\pi r^{1/4}}{\sqrt{r-1}} (\xi + y\sqrt{r-1})^{-1/2} \mid \xi + y\sqrt{r-1} > 0 \right).$$

In particular, then $\tau(3t_c) = \Theta_p(t_c^{1/4})$.

Note that in the supercritical case in Proposition 10.4, the time $\tau(3t_c)$ is always smaller than $t_c^{1/4}$, but that it approaches the order $t_c^{1/4}$ when $a - a_c^*$ grows only a little faster than the critical value $a_c^{1/2}$. Hence, we can say that the worst possible number of generations to pass the bottleneck at t_c is of the order $t_c^{1/4}$.

10.2. *The doubly exponential growth.* We next consider the growth from size $3t_c$ up to $1/p$. We will show that in this range, the growth is doubly exponential. Again, we approximate $A(t)$ by deterministic functions.

Define for any $\delta \in \mathbb{R}$ [cf. (3.4)],

$$(10.7) \quad F_\delta(t) := n \frac{(tp)^r}{r!} (1 + \delta) = \left(\frac{t}{t_c}\right)^{r-1} \frac{t}{r} (1 + \delta).$$

LEMMA 10.6. *For every $\delta > 0$, there are positive constants ε and K such that w.h.p. $F_{-\delta}(t) \leq A(t) \leq F_\delta(t)$ for all $t \in [K(t_c + a), \varepsilon/p]$.*

PROOF. By (8.1) and (10.7), for $Kt_c \leq t \leq \varepsilon/p$ (with $\varepsilon \leq 1$), if n is large enough so $t_c \geq 1$,

$$\pi(t) = \frac{(tp)^r}{r!} (1 + O(\varepsilon + K^{-1})) = \frac{1}{n} F_0(t) (1 + O(\varepsilon + K^{-1})).$$

We may thus choose ε and K such that for all such t (and large n)

$$(10.8) \quad F_{-\delta/4}(t) \leq n\pi(t) \leq F_{\delta/4}(t).$$

For $t \geq K(t_c + a)$, (10.7) implies

$$F_0(t) = \left(\frac{t}{t_c}\right)^{r-1} \frac{t}{r} \geq K^r \frac{a}{r},$$

so choosing K large enough, we have $a \leq (\delta/4)F_0(t)$ for all $t \in [K(t_c + a), \varepsilon/n]$, and thus by (10.8)

$$F_{-\delta/4}(t) - a \leq \mathbb{E}S(t) = (n - a)\pi(t) \leq F_{\delta/4}(t) \leq F_{\delta/2}(t) - a.$$

Hence, by Chebyshev’s inequality, using (2.13) and (10.8),

$$(10.9) \quad \begin{aligned} &\mathbb{P}\{A(t) \notin [F_{-3\delta/4}(t), F_{3\delta/4}(t)]\} \\ &= \mathbb{P}\{S(t) \notin [F_{-3\delta/4}(t) - a, F_{3\delta/4}(t) - a]\} \\ &\leq \frac{n\pi(t)}{(\delta F_0(t)/4)^2} \leq \frac{F_{\delta/4}(t)}{(\delta F_0(t)/4)^2} = \frac{16(1 + \delta/4)}{\delta^2 F_0(t)}. \end{aligned}$$

Define $t_j := (1 + \delta/5)^{j/r} K(t_c + a)$. Then, (10.9) and (10.7) show that, assuming as we may $\delta \leq 1$,

$$\begin{aligned} \sum_{j \geq 0: t_j \leq \varepsilon/p} \mathbb{P}\{A(t_j) \notin [F_{-3\delta/4}(t_j), F_{3\delta/4}(t_j)]\} &\leq \sum_{j \geq 0} \frac{20}{\delta^2 F_0(t_j)} \\ &= \sum_{j \geq 0} \frac{20}{\delta^2 F_0(t_0)} (1 + \delta/5)^{-j} \\ &= \frac{100(1 + \delta/5)}{\delta^3 F_0(t_0)} \rightarrow 0, \end{aligned}$$

since, using (10.7) again and (3.5),

$$F_0(t_0) = F_0(K(t_c + a)) \geq F_0(t_c) = \frac{t_c}{r} \rightarrow \infty.$$

Consequently, w.h.p. $A(t_j) \in [F_{-3\delta/4}(t_j), F_{3\delta/4}(t_j)]$ for all $j \geq 0$ with $t_j \leq \varepsilon/p$. However, if $t_j \leq t \leq t_{j+1}$, then $F_0(t_j) \leq F_0(t) \leq F_0(t_{j+1}) = (1 + \delta/5)F_0(t_j)$, and it follows that, since both $A(t)$ and $F_0(t)$ are monotone, w.h.p.

$$(1 + \delta/5)^{-1} F_{-3\delta/4}(t) \leq A(t) \leq (1 + \delta/5) F_{3\delta/4}(t)$$

for all $t \in [K(t_c + a), (1 + \delta/5)^{-1/r} \varepsilon/p]$, which, provided δ is small and ε is replaced by $\varepsilon/2$, say, yields the result. \square

PROPOSITION 10.7. *Suppose that $r \geq 2$ and $n^{-1} \ll p \ll n^{-1/r}$. Then w.h.p., when $A^* \geq 3t_c$,*

$$\tau(1/p) - \tau(3t_c) = \frac{1}{\log r} \left(\log \log(np) - \log_+ \log \frac{a}{a_c} \right) + O(1).$$

PROOF. Choose a fixed $0 < \delta < 1$, and choose ε and K as in Lemma 10.6. (In this proof, we do not have to let $\delta \rightarrow 0$, so we can take $\delta = 1/2$, say.) First, $\tau(K(t_c + a)) - \tau(3t_c)$, the number of generations from $3t_c$ to $K(t_c + a)$, is w.h.p. $O(1)$. Indeed, after $\tau(3t_c)$ generations we have at least $\max(3t_c, a)$ active vertices, and in each of the following generations until well beyond $K(t_c + a)$, the number is w.h.p. multiplied by at least 1.3, say, by the proof of Lemma 8.2 or by Lemmas 9.1, 9.4 and 7.3. Similarly, $\tau(1/p) - \tau(\varepsilon/p) \leq 1$ w.h.p., arguing as in Case 3 of the proof of Lemma 8.2.

Consequently it suffices to consider $\tau(\varepsilon/p) - \tau(K(t_c + a))$. We define iterates $T_j^{F_\delta}$ as in Section 10.1 by $T_{j+1}^{F_\delta} := F_\delta(T_j^{F_\delta})$, $j \geq 0$, but now starting with $T_0^{F_\delta} := K(t_c + a)$. Further, let

$$(10.10) \quad N_\delta := \min\{j \geq 0 : T_j^{F_\delta} \geq \varepsilon/p\}.$$

By Lemma 10.6 we may assume that $F_{-\delta}(t) \leq A(t) \leq F_\delta(t)$ for all $t \in [K(t_c + a), \varepsilon/p]$, and then, by induction as in Lemma 10.2, $T_j^{F_{-\delta}} \leq T_{j+\tau(K(t_c+a))} \leq T_{j+1}^{F_\delta}$ for all $j \geq 0$ with $T_{j-1+\tau(K(t_c+a))} \leq \varepsilon/p$. Consequently, w.h.p.

$$(10.11) \quad N_{-\delta} \geq \tau(\varepsilon/p) - \tau(K(t_c + a)) \geq N_\delta - 1.$$

To find N_δ , rewrite (10.7) as

$$\frac{F_\delta(t)}{c_\delta t_c} = \left(\frac{t}{c_\delta t_c} \right)^r,$$

where $c_\delta := (r/(1 + \delta))^{1/(r-1)}$. Iterating we see that, for $j \geq 0$,

$$\frac{T_j^{F_\delta}}{c_\delta t_c} = \left(\frac{T_0^{F_\delta}}{c_\delta t_c} \right)^{r^j} = \left(\frac{K(t_c + a)}{c_\delta t_c} \right)^{r^j}$$

and thus

$$\log \left(\frac{T_j^{F_\delta}}{c_\delta t_c} \right) = r^j \log \left(\frac{K(t_c + a)}{c_\delta t_c} \right)$$

and

$$j \log r = \log \log \left(\frac{T_j^{F_\delta}}{c_\delta t_c} \right) - \log \log \left(\frac{K(t_c + a)}{c_\delta t_c} \right).$$

Consequently,

$$(10.12) \quad N_\delta = \left\lceil \left(\log \log \left(\frac{\varepsilon/p}{c_\delta t_c} \right) - \log \log \left(\frac{K(t_c + a)}{c_\delta t_c} \right) \right) / \log r \right\rceil.$$

In order to simplify this, note that, using (3.1),

$$(10.13) \quad \log\left(\frac{\varepsilon/p}{c_\delta t_c}\right) = \log\left(\frac{1}{pt_c}\right) + O(1) = \frac{1}{r-1} \log(np) + O(1)$$

and thus

$$(10.14) \quad \log \log\left(\frac{\varepsilon/p}{c_\delta t_c}\right) = \log \log(np) + O(1).$$

Further, we may assume that $a \geq a_c/2 \geq t_c/4$, since otherwise the process is sub-critical and $A^* < 3t_c$ w.h.p. by Theorem 3.1. Hence, $\log(K(t_c + a)) = \log a + O(1)$ and thus, since also $\log(c_\delta t_c) = \log a_c + O(1)$,

$$(10.15) \quad \log\left(\frac{K(t_c + a)}{c_\delta t_c}\right) = \log a - \log a_c + O(1) = \log \frac{a}{a_c} + O(1).$$

We may assume that $K \geq ec_\delta$, so $\log(K(t_c + a)/(c_\delta t_c)) \geq 1$, and then (10.15) yields

$$(10.16) \quad \log \log\left(\frac{K(t_c + a)}{c_\delta t_c}\right) = \log_+ \log \frac{a}{a_c} + O(1).$$

Finally, (10.12), (10.14) and (10.16) yield

$$N_\delta \log r = \log \log(np) - \log_+ \log \frac{a}{a_c} + O(1).$$

Note that the right-hand side depends on δ only in the error term $O(1)$. Hence, we have the same result for $N_{-\delta}$, and the result follows by (10.11) and the comments at the beginning of the proof. \square

10.3. *The final stage.* We finally consider the evolution after $1/p$ vertices have become active. We let, as in Section 8, $b^* := b_c \omega(n)$ where $\omega(n) \rightarrow \infty$ slowly; we assume that $b^* \ll 1/p$ [which is possible since $pb_c \rightarrow 0$ by (3.5)]. By Remark 8.3, $\tau(n - b^*) \leq \tau(1/p) + 3$ w.h.p., so it suffices to consider the evolution when less than b^* vertices remain.

Let $\mathcal{F}_t := \sigma\{I_i(s) : 1 \leq i \leq n, 1 \leq s \leq t\}$ be the σ -field describing the evolution up to time t .

LEMMA 10.8. *For any t and u with $0 \leq t \leq t + u \leq n$, the conditional distribution of $A(t + u) - A(t) = S(t + u) - S(t)$ given \mathcal{F}_t is $\text{Bin}(n - A(t), \pi(t; u))$, where*

$$(10.17) \quad \pi(t; u) := \frac{\pi(t + u) - \pi(t)}{1 - \pi(t)}.$$

If further $n - b^ \leq t \leq t + u \leq n$, then, uniformly in all such t and u ,*

$$(10.18) \quad \pi(t; u) = pu(1 + o(1)).$$

PROOF. Conditioned on \mathcal{F}_t , $A(t)$ is a given number, and of the $n - a$ summands in (2.8), $n - a - S(t) = n - A(t)$ are zero. For any of these terms, the probability that it changes from 0 at time t to 1 at time $t + u$ is, by (2.11),

$$\mathbb{P}(Y_i \leq t + u \mid Y_i > t) = \frac{\mathbb{P}(t < Y_i \leq t + u)}{\mathbb{P}(Y_i > t)} = \frac{\pi(t + u) - \pi(t)}{1 - \pi(t)} = \pi(t; u).$$

Hence, the conditional distribution of $S(t + u) - S(t)$ is $\text{Bin}(n - A(t), \pi(t; u))$.

To see the approximation (10.18), note first that for $n - b^* \leq t \leq n$, since we assume $pb^* \rightarrow 0$, we have $b^* \ll 1/p \ll n$ so $t \sim n$. Hence, using again $pb^* \rightarrow 0$ and recalling the notation b'_c from (3.7),

$$\begin{aligned} \pi(t + 1) - \pi(t) &= \mathbb{P}(Y_1 = t + 1) = \binom{t}{r - 1} p^r (1 - p)^{t+1-r} \\ (10.19) \quad &\sim \frac{n^{r-1}}{(r - 1)!} p^r (1 - p)^n = \frac{pb'_c}{n}. \end{aligned}$$

Furthermore [cf. (3.7)], still for $n - b^* \leq t \leq n$,

$$\begin{aligned} 1 - \pi(t) &= \mathbb{P}(\text{Bin}(t, p) \leq r - 1) \sim \mathbb{P}(\text{Bin}(t, p) = r - 1) \\ (10.20) \quad &\sim \frac{n^{r-1}}{(r - 1)!} p^{r-1} (1 - p)^n = \frac{b'_c}{n}. \end{aligned}$$

Consequently, $\pi(t + u) - \pi(t) = (1 + o(1))upb'_c/n$ and

$$\pi(t; u) = \frac{\pi(t + u) - \pi(t)}{1 - \pi(t)} = (1 + o(1)) \frac{upb'_c/n}{b'_c/n} = (1 + o(1))up. \quad \square$$

LEMMA 10.9. *Suppose that $r \geq 2$, $n^{-1} \ll p \ll n^{-1/r}$ and $a = o(n)$. If $b_c \rightarrow \infty$ and $n - b^* \leq t \leq n$, then $A(t) = n - b_c(1 + o_p(1))$; in particular, $n - A(t) < 2b_c$ w.h.p.*

PROOF. We have, using (2.12) and (10.20), since $b_c \rightarrow \infty$ implies $b'_c \sim b_c$,

$$\mathbb{E}(n - A(t)) = n - a - \mathbb{E} S(t) = (n - a)(1 - \pi(t)) \sim (n - a) \frac{b'_c}{n} \sim b_c$$

and similarly, using (2.13),

$$\text{Var}(n - A(t)) = \text{Var} S(t) \leq (n - a)(1 - \pi(t)) \sim b_c.$$

Thus, by Chebyshev's inequality, since $b_c \rightarrow \infty$,

$$n - A(t) = (1 + o(1))b_c + O_p(b_c^{1/2}) = (1 + o_p(1))b_c. \quad \square$$

PROPOSITION 10.10. *Suppose that $r \geq 2$, $n^{-1} \ll p \ll n^{-1/r}$ and $a = o(n)$. Then, when $A^* \geq 3t_c$,*

$$\tau - \tau(1/p) = (1 + o(1)) \frac{\log n}{np} + O_p(1).$$

In particular, if further $p \geq c \log(n)/n$ for some $n \geq 0$, then $\tau - \tau(1/p) = O_p(1)$. Furthermore, when $A^ = n$, w.h.p. $\tau - \tau(1/p) \leq 3$.*

PROOF. By Remark 8.3, after $\tau(1/p) + 3$ generations, the active size is $T_{\tau(1/p)+3} \geq n - b^*$ w.h.p.

If $b_c \rightarrow 0$, we can choose $b^* = 1/2$, so w.h.p. $T_{\tau(1/p)+3} = n$ and $\tau \leq \tau(1/p) + 3$. More generally, if $b_c = O(1)$, we have by (10.19),

$$\begin{aligned} \mathbb{E}(S(n) - S(n - b^*)) &\leq n(\pi(n) - \pi(n - b^*)) \sim nb^* \frac{pb'_c}{n} \\ &= pb^*b_c = O(pb^*) = o(1). \end{aligned}$$

Hence, w.h.p. $S(n) = S(n - b^*)$, which means that no further activations occur after $n - b^*$. Consequently, in this case too, w.h.p. $\tau = \tau(n - b^*) \leq \tau(1/p) + 3$. In particular, this proves that $\tau \leq \tau(1/p) + 3$ w.h.p. when $A^* = n$, since w.h.p. $A^* < n$ if $b_c \rightarrow \infty$ by Theorem 3.2.

Further, when $b_c = O(1)$, (3.10) implies that $np \geq \log n$ for large n , so $\log n/(np) \leq 1$, and the result holds in this case.

Now assume that $b_c \rightarrow \infty$. For convenience, we modify the counting of generations and start at $t = n - b^*$, regarding the active but unused vertices at $n - b^*$ as “generation 0.” (We may assume that b^* is an integer.) Thus define, recursively,

$$\begin{aligned} T'_0 &:= n - b^*, \\ T'_{j+1} &:= A(T'_j), \quad j \geq 0, \\ \Delta_j &:= T'_{j+1} - T'_j = A(T'_j) - T'_j, \\ \tau' &:= \max\{j \geq 0 : \Delta_j > 0\}. \end{aligned}$$

Since w.h.p. $T_{\tau(1/p)-1} \leq \max(1/p, a) < n - b^* \leq T_{\tau(1/p)+3}$, it follows by induction that $T_{\tau(1/p)-1+j} \leq T'_j \leq T_{\tau(1/p)+3+j}$, $j \geq 0$, and thus w.h.p.

$$(10.21) \quad \tau' + \tau(1/p) - 1 \leq \tau \leq \tau' + \tau(1/p) + 3.$$

Consequently, it suffices to estimate τ' .

By Lemma 10.8, conditioned on $\mathcal{F}_{T'_j}$ [i.e., on T'_j and the evolution up to T'_j , which in particular specifies $A(T'_j)$], for large n ,

$$\mathbb{E}(\Delta_{j+1} \mid \mathcal{F}_{T'_j}) = (n - A(T'_j))\pi(T'_j; \Delta_j) \leq (n - A(T'_0))2p\Delta_j$$

and thus, by induction, since $\Delta_0 \leq n - T'_0 = b^*$,

$$(10.22) \quad \mathbb{E}(\Delta_j \mid \mathcal{F}_{T'_0}) \leq (2(n - A(T'_0))p)^j \Delta_0 \leq (2(n - A(T'_0))p)^j b^*.$$

Further, Lemma 10.9 yields $n - A(T'_0) = n - A(n - b^*) < 2b_c$ w.h.p. Consequently, (10.22) implies, w.h.p. for all $j \geq 0$ (simultaneously),

$$(10.23) \quad \mathbb{E}(\Delta_j \mid \mathcal{F}_{T'_0}) \leq (4pb_c)^j b^*.$$

Recall that $pb_c \rightarrow 0$ by (3.5), so we may assume $4pb_c < 1$. If j is chosen such that $(4pb_c)^j b^* \rightarrow 0$, then (10.23) implies that w.h.p. $\Delta_j = 0$ and thus $\tau' < j$. Hence, for any $\omega' = \omega'(n) \rightarrow \infty$, w.h.p.

$$\tau' \leq \frac{\log b^*}{|\log(pb_c) + \log 4|} + \omega'(n),$$

which is another way of saying [33], Lemma 3,

$$(10.24) \quad \tau' \leq \frac{\log b^*}{|\log(pb_c) + \log 4|} + O_p(1) = \frac{\log b^*}{|\log(pb_c)|} (1 + o(1)) + O_p(1).$$

For a lower bound, fix ε with $0 < \varepsilon < 1$, and define the deterministic numbers Δ_j^- by

$$(10.25) \quad \Delta_j^- := (1 - \varepsilon)^{j+1} (pb_c)^j b^*.$$

Let $\omega'' := 1/(pb_c) \rightarrow \infty$. We claim that w.h.p.

$$(10.26) \quad \Delta_j \geq \Delta_j^- \quad \text{for all } j \geq 0 \text{ such that } \Delta_j^- \geq \omega''.$$

By our assumption $4pb_c < 1$, we have $\Delta_{j+1}^-/\Delta_j^- < 1/4$, so $\Delta_j^- \rightarrow 0$ geometrically fast.

By Lemma 10.9 and $b_c/b^* \rightarrow 0$, w.h.p.

$$\Delta_0 = A(T'_0) - (n - b^*) = A(T'_0) - n + b^* \geq b^* - 2b_c \geq (1 - \varepsilon)b^* = \Delta_0^-,$$

so (10.26) holds w.h.p. for $j = 0$.

Say that $j \geq 0$ is *good* if $\Delta_j \geq \Delta_j^-$ and *fat* if $A(T'_j) > n - (1 - \varepsilon/4)b_c$. Let $j \geq 0$. At time T'_j we have $A(T'_j) - T'_j = \Delta_j$ active but unused vertices. Further, by Lemma 10.8 we have, conditioned on $\mathcal{F}_{T'_j}$ (which specifies both T'_j and Δ_j),

$$\Delta_{j+1} = T'_{j+2} - T'_{j+1} = A(T'_j + \Delta_j) - A(T'_j) \in \text{Bin}(n - A(T'_j), \pi(T'_j; \Delta_j)).$$

By Lemma 10.8, $\pi(T'_j; \Delta_j) = p\Delta_j(1 + o(1)) \geq p\Delta_j(1 - \varepsilon/4)$ for n large, so if j is good but not fat,

$$\begin{aligned} \mathbb{E}(\Delta_{j+1} \mid \mathcal{F}_{T'_j}) &= (n - A(T'_j))\pi(T'_j; \Delta_j) \geq (1 - \varepsilon/4)^2 b_c p \Delta_j \\ &\geq (1 - \varepsilon/2) b_c p \Delta_j^- \geq (1 + \varepsilon/2) \Delta_{j+1}^- \end{aligned}$$

and Chebyshev’s inequality yields, since $x \mapsto x/(x - a)^2$ is decreasing for $x > a$,

$$\begin{aligned} \mathbb{P}(\Delta_{j+1} < \Delta_{j+1}^- \mid \mathcal{F}_{T'_j}) &\leq \frac{\text{Var}(\Delta_{j+1} \mid \mathcal{F}_{T'_j})}{(\mathbb{E}(\Delta_{j+1} \mid \mathcal{F}_{T'_j}) - \Delta_{j+1}^-)^2} \\ &\leq \frac{\mathbb{E}(\Delta_{j+1} \mid \mathcal{F}_{T'_j})}{(\mathbb{E}(\Delta_{j+1} \mid \mathcal{F}_{T'_j}) - \Delta_{j+1}^-)^2} \\ &\leq \frac{(1 + \varepsilon/2)\Delta_{j+1}^-}{(\varepsilon\Delta_{j+1}^-/2)^2} = O\left(\frac{1}{\Delta_{j+1}^-}\right). \end{aligned}$$

Say that j is *bad* if j is not good and that j *fails* if j is fat or bad. Then, by stopping at the first j that fails we see that

$$\begin{aligned} \mathbb{P}(\text{some } j \leq \omega'' \text{ fails}) &\leq \mathbb{P}(\text{some } j \leq \omega'' \text{ is fat}) + \mathbb{P}(0 \text{ is bad}) \\ &\quad + \sum_{j>0: \Delta_j^- \geq \omega''} \mathbb{P}(j \text{ is bad} \mid j - 1 \text{ is good and not fat}) \\ &\leq \mathbb{P}(A(n) > n - (1 - \varepsilon/4)b_c) + o(1) \\ &\quad + \sum_{j: \Delta_j^- \geq \omega''} o\left(\frac{1}{\Delta_j^-}\right) \\ &= o(1), \end{aligned}$$

since $A(n) < n - (1 - \varepsilon/4)b_c$ w.h.p. by Lemma 10.9 and the final sum is $O(1/\omega'') = o(1)$ because the terms $1/\Delta_j^-$ increase geometrically, so the sum is dominated by its largest (and last) term.

We have shown that w.h.p., if $\Delta_j^- \geq \omega''$, then $\Delta_j \geq \Delta_j^- > 0$ and thus $\tau' \geq j$. Hence, by (10.26) and (10.25), w.h.p.

$$(10.27) \quad \tau' \geq \left\lfloor \frac{\log((1 - \varepsilon)b^*/\omega'')}{|\log((1 - \varepsilon)pb_c)|} \right\rfloor = \frac{\log b^*}{|\log(pb_c)|}(1 + o(1)) + O(1).$$

Combining the upper bound (10.24) and the lower bound (10.27), we find

$$(10.28) \quad \tau' = \frac{\log b^*}{|\log(pb_c)|}(1 + o(1)) + O_p(1).$$

By (3.3), $\log(pb_c) = -(np - r \log(np) + O(1))$ and

$$\log n \geq \log b^* \geq \log b_c \geq \log n - pn - O(1).$$

Hence, finally (10.28) yields

$$\tau' = \frac{\log n + O(np)}{np - r \log(np) + O(1)}(1 + o(1)) + O_p(1) = \frac{\log n}{np}(1 + o(1)) + O_p(1).$$

The result now follows from (10.21). \square

11. Proofs of Theorems 5.2, 5.6, 5.8. We prove in this section Theorems 5.2, 5.6 and 5.8 related to the boundary cases. We consider first the case $p \sim c/n$.

PROOF OF LEMMA 5.1. By the implicit function theorem, at least locally, the root $x_0(\theta)$ is smooth except at points where

$$(11.1) \quad f(x, c, \theta) = \frac{\partial}{\partial x} f(x, c, \theta) = 0.$$

We begin by studying such *critical points*.

Let $g(y) := \mathbb{P}(\text{Po}(y) \leq r - 1) = 1 - \psi(y)$; cf. (3.13). Differentiations yield

$$(11.2) \quad g'(y) = -\mathbb{P}(\text{Po}(y) = r - 1) = -\frac{y^{r-1}}{(r - 1)!} e^{-y},$$

$$(11.3) \quad g''(y) = \left(\frac{r - 1}{y} - 1\right)g'(y) = \frac{r - 1 - y}{y}g'(y).$$

We have [see (5.2)] $f(x, c, \theta) = 1 - x - (1 - \theta)g(cx)$ and thus $\frac{\partial}{\partial x} f(x, c, \theta) = -1 - c(1 - \theta)g'(cx)$. Hence, (11.1) holds if and only if

$$\begin{cases} (1 - \theta)g(cx) = 1 - x, \\ c(1 - \theta)g'(cx) = -1, \end{cases}$$

which imply $g(cx) = -c(1 - x)g'(cx)$ and thus

$$(11.4) \quad c = cx - \frac{g(cx)}{g'(cx)}.$$

Let $h(y) := y - g(y)/g'(y)$, $y > 0$, so (11.4) says $c = h(cx)$. Then, by (11.3),

$$h'(y) = 1 - \frac{g'(y)}{g'(y)} + \frac{g(y)g''(y)}{g'(y)^2} = \frac{r - 1 - y}{y} \cdot \frac{g(y)}{g'(y)}.$$

Since $g(y) > 0$ and $g'(y) < 0$ for $y > 0$, h has a global minimum at $y = r - 1$, and the minimum value is

$$\min_{y>0} h(y) = h(r - 1) = r - 1 - \frac{g(r - 1)}{g'(r - 1)} = r + \frac{\mathbb{P}(\text{Po}(r - 1) \leq r - 2)}{\mathbb{P}(\text{Po}(r - 1) = r - 1)} = c_c.$$

Furthermore, $h(y) > y \rightarrow \infty$ as $y \rightarrow \infty$, and $h(y) \rightarrow \infty$ as $y \rightarrow 0$ too, because then $g(y) \rightarrow 1$ and $g'(y) \rightarrow 0$.

Consequently, if $0 \leq c < c_c$, then (11.4) has no solution $x > 0$, and thus there is no critical point. If $c = c_c$, there is exactly one $x > 0$ satisfying (11.4) [viz., $x = (r - 1)/c_c$], and if $c > c_c$, there are two. Since (11.4) implies $c > cx$, these roots are in $(0, 1)$.

To complete the proof, it is perhaps simplest to rewrite (5.4) as $\theta = \vartheta(x)$, with

$$(11.5) \quad \vartheta(x) := 1 - (1 - x)/g(cx).$$

Since $g(y) > 0$ for $y \geq 0$, ϑ is a smooth function on $[0, 1]$, with $\vartheta(0) = 0$ and $\vartheta(1) = 1$. Moreover, $f(x, c, \theta) = g(cx)(\theta - \vartheta(x))$, which implies that

$$f(x, c, \theta) = \frac{\partial}{\partial x} f(x, c, \theta) = 0 \iff \theta = \vartheta(x) \text{ and } \vartheta'(x) = 0.$$

Consequently, by the results above, if $c < c_c$, then $\vartheta' \neq 0$ so $\vartheta'(x) > 0$ for $x \geq 0$. In this case, ϑ is strictly increasing and thus a bijection $[0, 1] \rightarrow [0, 1]$, and x_0 is its inverse.

If $c = c_c$, then $\vartheta' = 0$ only at a single point, and it follows again that ϑ is a strictly increasing function and x_0 is its inverse.

If $c > c_c$, then $\vartheta'(x) = 0$ at two values x_1 and x_2 with $0 < x_1 < x_2 < 1$ and $cx_1 < r - 1 < cx_2$. It can be seen, for example, using (11.3), that $\vartheta''(x_1) < 0 < \vartheta''(x_2)$, and thus ϑ is *decreasing* on the interval $[x_1, x_2]$. The result follows, with $\theta_c = \vartheta(x_1)$, $\theta_c^- = \max(\vartheta(x_2), 0)$ and $x_0(\theta_c) = x_1$. [Note that $\vartheta(x_2) = \min_{x \in [0, 1]} \vartheta(x) < 0$ if c is large enough.] \square

REMARK 11.1. If $c > c_c$, then thus $x_0(\theta_c) = x_1$ is the smallest root of $\vartheta'(x) = 0$, or equivalently $x_1 = y_1/c$ where y_1 is the smallest root of $h(y) = c$; further, $\theta_c = \vartheta(x_1)$ while $x_0(\theta_c^+)$ is the other root of $\vartheta(x) = \vartheta(x_1)$.

If $c = c_c$, we have $y_1 = r - 1$ and thus $x_1 = (r - 1)/c_c$ and, using (11.5) and (5.5),

$$\begin{aligned} \theta_c(c_c) &= \vartheta\left(\frac{r-1}{c_c}\right) = 1 - \frac{1 - (r-1)/c_c}{g(r-1)} \\ &= 1 - \frac{1}{r\mathbb{P}(\text{Po}(r-1) = r-1) + \mathbb{P}(\text{Po}(r-1) \leq r-2)}. \end{aligned}$$

For $c > c_c$, the two roots $x_1(c)$ and $x_2(c)$ of $\vartheta'(x) = 0$ are smooth functions of c , and thus

$$\frac{d\theta_c}{dc} = \frac{d}{dc} \vartheta(x_1(c)) = \frac{\partial \vartheta}{\partial c}(x_1(c)) + \frac{\partial \vartheta}{\partial x}(x_1(c))x_1'(c) = \frac{\partial \vartheta}{\partial c}(x_1(c)) < 0,$$

where the last inequality follows from (11.5), and similarly $d\theta_c^-/dc < 0$. Hence, $\theta_c(c)$ and $\theta_c^-(c)$ are decreasing functions of c , as claimed in Remark 5.3.

LEMMA 11.2. *Suppose that $r \geq 2$, $p = O(1/n)$ and $tp = o(1)$. Then $S(t) = o_p(t)$.*

PROOF. We may assume $1 \leq t \leq 1/p$. [Note that $S(t) = 0$ for $t < r$.] Then $\pi(t) = O(t^r p^r) = o(tp)$ by (8.1), and thus the expected number of activated vertices is $\mathbb{E} S(t) = (n - a)\pi(t) = o(npt) = o(t)$. \square

PROOF OF THEOREM 5.2. First, in (i) and (ii), $ap \rightarrow \theta c = 0$. Let $\varepsilon > 0$. Taking $t = (1 + \varepsilon)a$ in Lemma 11.2, we find w.h.p. $S((1 + \varepsilon)a) < \varepsilon a$ and thus

$$A((1 + \varepsilon)a) = a + S((1 + \varepsilon)a) < (1 + \varepsilon)a,$$

whence $A^* = T < (1 + \varepsilon)a$. Consequently, $1 \leq A^*/a < 1 + \varepsilon$ w.h.p., proving (i) and (ii).

Next, by (2.9), Lemma 7.1 and (2.12), uniformly for all $t \geq 0$,

$$A(t) = a + S(t) = a + \mathbb{E} S(t) + o_p(n) = (n - a)\pi(t) + a + o_p(n)$$

and thus, using also (3.15),

$$n^{-1}A(t) = (1 - \theta)\pi(t) + \theta + o_p(1) = (1 - \theta)\tilde{\pi}(t) + \theta + o_p(1).$$

Substituting $t = xn$, we find by (3.13), since $tp = xc + o(x)$, uniformly in all $x \geq 0$,

$$\begin{aligned} n^{-1}A(xn) &= (1 - \theta)\mathbb{P}(\text{Po}(tp) \geq r) + \theta + o_p(1) \\ &= (1 - \theta)\mathbb{P}(\text{Po}(cx) \geq r) + \theta + o_p(1) \end{aligned}$$

and, recalling (5.1), still uniformly in $x \geq 0$,

$$(11.6) \quad n^{-1}(A(xn) - xn) = f(x, c, \theta) + o_p(1).$$

Let $\varepsilon > 0$. Since $f(x, c, \theta) > 0$ for $x \in [0, x_0(\theta))$, and thus by compactness $f(\cdot, c, \theta)$ is bounded from below on $[0, x_0(\theta) - \varepsilon]$, (11.6) implies that w.h.p. $A(xn) - xn > 0$ on $[0, x_0(\theta) - \varepsilon]$, and thus $T > (x_0(\theta) - \varepsilon)n$. Furthermore, both in (iii) and in (iv) with $\theta \neq \theta_c$, we have $\frac{\partial}{\partial x} f(x_0(\theta), c, \theta) \neq 0$ and thus if $\varepsilon > 0$ is small enough, $f(x_0(\theta) + \varepsilon, c, \theta) < 0$, so (11.6) implies that w.h.p. $A((x_0(\theta) + \varepsilon)n) < (x_0(\theta) + \varepsilon)n$ and thus $T < (x_0(\theta) + \varepsilon)n$. \square

The proof of Theorem 5.5 is very similar to the one of Theorem 3.6. We first give a more precise estimate of the process $S(t)$, which is the analog of Lemma 9.3 in the case $p = c/n$. However, in this case, we get a Brownian bridge because here we consider a large part of the distribution of Y_i .

LEMMA 11.3. Suppose $r \geq 2$, $p = c/n$ and $a \sim \theta n$ with $c > 0$ and $0 < \theta < 1$. Then

$$(11.7) \quad Z(x) := \frac{S(xn) - \mathbb{E} S(xn)}{\sqrt{(1 - \theta)n}} \xrightarrow{d} W_0(\psi(cx))$$

in $D[0, 1]$, where W_0 is a Brownian bridge and $\psi(y) := \mathbb{P}(\text{Po}(y) \geq r)$ as in (3.13).

PROOF. Let $\tilde{S}(u) := \sum_{i=1}^{n-a} \mathbf{1}\{U_i \leq u\}$, $0 \leq u \leq 1$, where $U_i \in U(0, 1)$ are i.i.d. By (2.8) and (2.11), we have $S(t) \stackrel{d}{=} \tilde{S}(\pi(t))$, jointly for all $t \geq 0$. Further,

$\frac{1}{n-a}\tilde{S}(u)$, $u \in [0, 1]$, is the empirical distribution function of U_1, \dots, U_{n-a} , and thus by [16], Theorem 16.4, in $D[0, 1]$,

$$\frac{\tilde{S}(u) - \mathbb{E}\tilde{S}(u)}{\sqrt{n-a}} \xrightarrow{d} W_0(u).$$

Furthermore, by (3.15) and (3.13),

$$\pi(xn) = \tilde{\pi}(xn) + O(1/n) = \psi(xnp) + O(1/n) = \psi(cx) + O(1/n),$$

uniformly for $x \geq 0$, and it follows, using the continuity of W_0 , that

$$\frac{S(xn) - \mathbb{E}S(xn)}{\sqrt{n-a}} \stackrel{d}{=} \frac{\tilde{S}(\pi(xn)) - \mathbb{E}\tilde{S}(\pi(xn))}{\sqrt{n-a}} \xrightarrow{d} W_0(\psi(cx))$$

in $D[0, 1]$, which proves the result since $n - a \sim (1 - \theta)n$. \square

PROOF OF THEOREM 5.5. It suffices to consider a such that $a \sim \theta_c n$. By (2.12), (3.15) and (3.13),

$$\begin{aligned} \mathbb{E}S(xn) &= (n-a)\pi(xn) = (n-a)\tilde{\pi}(xn) + O(1) \\ (11.8) \quad &= (n-a)\psi(cx) + O(1). \end{aligned}$$

By the Skorohod coupling theorem ([35], Theorem 4.30), we may assume that the processes for different n are coupled such that the limit (11.7) in Lemma 11.3 holds a.s., and not just in distribution. Since convergence in $D[0, 1]$ to a continuous function is equivalent to uniform convergence, this means that a.s. $Z(x) \rightarrow W_0(\psi(cx))$ uniformly for $x \in [0, 1]$. Hence, we have, using (11.8) and (5.1),

$$\begin{aligned} A(xn) - xn &= a + S(xn) - xn \\ &= a + \mathbb{E}S(xn) + \sqrt{(1-\theta_c)n}Z(x) - xn \\ (11.9) \quad &= a + (n-a)\psi(cx) + \sqrt{(1-\theta_c)n}Z(x) - xn + O(1) \\ &= (a - \theta_c n)(1 - \psi(cx)) + nf(x, c, \theta_c) + \sqrt{(1-\theta_c)n}Z(x) + O(1) \\ &= nf(x, c, \theta_c) + (a - \theta_c n)(1 - \psi(cx)) + \sqrt{(1-\theta_c)n}W_0(\psi(cx)) \\ &\quad + o_p(n^{1/2}), \end{aligned}$$

uniformly for $x \in [0, 1]$.

We first use (11.9) to derive the simple estimate

$$(11.10) \quad A(xn) - xn = nf(x, c, \theta_c) + o_p(n),$$

uniformly for $x \in [0, 1]$. By Lemma 5.1, $f(x, c, \theta_c) = 0$ for $x = x_0$ or $x = x_1$, with $f(x, c, \theta_c) > 0$ for $x \in [0, x_0) \cup (x_0, x_1)$ and $f(x, c, \theta_c) < 0$ for $x \in (x_1, 1]$.

Hence, for any fixed small $\varepsilon > 0$, (11.10) implies that w.h.p. $A(xn) - xn > 0$ for $x \in [0, x_0 - \varepsilon] \cup [x_0 + \varepsilon, x_1 - \varepsilon]$ and $A(xn) - xn < 0$ for $x \in [x_1 + \varepsilon, 1]$, and hence $T \in [x_0 - \varepsilon, x_0 + \varepsilon] \cup [x_1 - \varepsilon, x_1 + \varepsilon]$. It follows by a standard argument that there exists a sequence $\varepsilon_n \searrow 0$ such that w.h.p.

$$A^* = T \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n] \cup [x_1 - \varepsilon_n, x_1 + \varepsilon_n].$$

Moreover, w.h.p. $T \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]$ if and only if $\inf_{[x_0 - \varepsilon_n, x_0 + \varepsilon_n]} (A(xn) - xn) < 0$. (We may also assume that ε_n is so small that $\varepsilon_n < x_0$ and $2\varepsilon_n < x_1 - x_0$.)

For $x \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]$, we have by (11.9) again, and the continuity of ψ and W_0 ,

$$(11.11) \quad \begin{aligned} A(xn) - xn &= nf(x, c, \theta_c) + (a - \theta_c n)(1 - \psi(cx_0)) + o(1) \\ &\quad + \sqrt{(1 - \theta_c)n} W_0(\psi(cx_0)) + o_p(n^{1/2}). \end{aligned}$$

Further, $f(x_0, c, \theta_c) = 0$ and $f(x, c, \theta_c) \geq 0$ for $x \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]$, and thus (11.11) yields

$$(11.12) \quad \begin{aligned} \inf_{x \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]} (A(xn) - xn) &= (a - \theta_c n)(1 - \psi(cx_0)) + o(1) \\ &\quad + \sqrt{(1 - \theta_c)n} W_0(\psi(cx_0)) + o_p(n^{1/2}). \end{aligned}$$

The cases (i) and (ii) are easily derived. We thus focus on (iii). We then have, from (11.12),

$$\begin{aligned} n^{-1/2} \inf_{x \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]} (A(xn) - xn) \\ = y(1 - \psi(cx_0)) + \sqrt{1 - \theta_c} W_0(\psi(cx_0)) + o_p(1) \end{aligned}$$

and thus, since $(1 - \psi(cx_0))^{-1} \sqrt{1 - \theta_c} W_0(\psi(cx_0)) \in N(0, \sigma^2)$, where $\sigma^2 = (1 - \theta_c)\psi(cx_0)/(1 - \psi(cx_0)) > 0$,

$$\begin{aligned} \mathbb{P}\left(\inf_{x \in [x_0 - \varepsilon_n, x_0 + \varepsilon_n]} (A(xn) - xn) < 0\right) \\ = \mathbb{P}(y(1 - \psi(cx_0)) + \sqrt{1 - \theta_c} W_0(\psi(cx_0)) < 0) + o_p(1) \\ = 1 - \Phi(y/\sigma) + o_p(1). \end{aligned}$$

The result follows. \square

To prove Theorem 5.6 ($p \sim cn^{-1/r}$), we first show using the previous results that if we can activate $\omega(n) \rightarrow \infty$ vertices, then the activation spreads w.h.p. to the entire graph. It remains to show that starting with a finite number of active vertices, the process activates $\omega(n)$ vertices with a probability bounded away from 0 and 1. This will be done using a branching process argument.

LEMMA 11.4. *Suppose that $p \geq cn^{-1/r}$ for some $c > 0$. If $\omega(n) \rightarrow \infty$, then w.h.p. $A(t) > t$ for all t with $\omega(n) \leq t \leq n - 1$.*

PROOF. This is easy to prove directly, but we prefer to view it as a corollary of our estimates for smaller p . Thus, let $p' := \omega(n)^{-1/2r} n^{-1/r}$. We may assume $\omega(n) \leq n$ and then $n^{-1} \ll p' \ll n^{-1/r}$, so $p' < p$, at least for large n , and we may assume that $G_{n,p'} \subseteq G_{n,p}$. We may consider bootstrap percolation on $G_{n,p'}$ and $G_{n,p}$ simultaneously, with the same initial set A_0 of size a ; we use the description in Section 2, starting with families of i.i.d. random indicators $I'_i(s) \in \text{Be}(p')$ and $I_i(s) \in \text{Be}(p)$ where we may assume $I'_i(s) \leq I_i(s)$. Then, using $'$ to denote variables for $G_{n,p'}$, $S'(t) \leq S(t)$ and $A'(t) \leq A(t)$.

We apply Lemma 8.2 to $G_{n,p'}$. The critical time for $G_{n,p'}$ is [see (3.1)]

$$t'_c = O((n(p')^r)^{-1/(r-1)}) = \omega(n)^{1/2(r-1)} = o(\omega(n)).$$

Further, $p' \geq n^{-3/2r} \geq n^{-3/4}$ so, by (3.3), $b'_c \rightarrow 0$, and we may choose $b^{*'} with $b^{*'} \rightarrow 0$. Hence, Lemma 8.2 shows that w.h.p. $A(t) \geq A'(t) > t$ for $t \in [3t'_c, n - b^{*'}]$, and the result follows since, for large n , $3t'_c \leq \omega(n)$ and $n - b^{*' > n - 1$. $\square$$

PROOF OF THEOREM 5.6. For (ii), we apply Lemma 11.4 (if necessary with a smaller c). Taking $\omega(n) = a$, we see that w.h.p. $A(t) > t$ for all $t \in [a, n - 1]$. Since also $A(t) \geq a$, we have $A(t) > t$ for all $t \leq n - 1$, and thus $A^* = T = n$.

For (i) suppose $r \geq 2$, $p \sim cn^{-1/r}$ and let $a \geq r$ be some constant. The probability that a vertex is activated at a given time k is by (2.7)

$$(11.13) \quad \mathbb{P}(Y_1 = k) = \binom{k-1}{r-1} p^r (1-p)^{k-r} \sim \binom{k-1}{r-1} \frac{c^r}{n}.$$

For any fixed K , the random variables

$$X_k := A(k) - A(k-1) = S(k) - S(k-1) = \sum_{i \notin A(0)} \mathbf{1}\{Y_i = k\},$$

$k = 1, \dots, K$, form together with

$$X_{K+1} := n - a - A(K) = \sum_{i \notin A(0)} \mathbf{1}\{Y_i > K\}$$

a random vector with the multinomial distribution $\text{Mul}(n - a, (p_k)_{k=1}^{K+1})$ with $p_k = \mathbb{P}(Y_1 = k)$, $k \leq K$, and $p_{K+1} = \mathbb{P}(Y_1 \geq K + 1)$. By (11.13), $(n - a)p_k \rightarrow \binom{k-1}{r-1} c^r$ for $k \leq K$, and it follows that X_k for $k \leq K$ have a joint Poisson limit,

$$(11.14) \quad (X_k)_{k=1}^K \xrightarrow{d} (\xi_k)_{k=1}^K \quad \text{with } \xi_k \in \text{Po}\left(\binom{k-1}{r-1} c^r\right) \quad \text{independent.}$$

Using the notation of Remark 5.7 we thus obtain

$$A(k) \xrightarrow{d} a + \sum_{j=1}^k \xi_j = a + k + \tilde{S}_k \quad \text{for } k = 1, \dots, t \quad \text{jointly}$$

and thus $\mathbb{P}(T = k) \rightarrow \mathbb{P}(\tilde{T} = k)$ for $k \leq K$ and $\mathbb{P}(T > K) \rightarrow \mathbb{P}(\tilde{T} > K)$.

Since K is arbitrary, we have shown $\mathbb{P}(A^* = k) = \mathbb{P}(T = k) \rightarrow \mathbb{P}(\tilde{T} = k) = \zeta(a, c, k)$ for every finite $k \geq 1$. Furthermore, $\mathbb{P}(T > K) - \mathbb{P}(\tilde{T} > K) \rightarrow 0$ for any fixed K , and a standard argument shows that there exists a sequence $K_n \rightarrow \infty$ such that $\mathbb{P}(T > K_n) - \mathbb{P}(\tilde{T} > K_n) \rightarrow 0$, and thus $\mathbb{P}(T > K_n) \rightarrow \mathbb{P}(\tilde{T} = \infty)$. On the other hand, Lemma 11.4 with $\omega(n) = K_n$ shows that $\mathbb{P}(K_n \leq T < n) \rightarrow 0$. Consequently, $\mathbb{P}(T = n) = \mathbb{P}(T > K_n) + o(1) \rightarrow \mathbb{P}(\tilde{T} = \infty) = \zeta(a, c)$.

It is clear that $\zeta(a, c, k) = \mathbb{P}(\tilde{T} = k) > 0$ for every $k \geq a$. To see that also $\zeta(a, c) = \mathbb{P}(\tilde{T} = \infty) > 0$, note that, see (11.14), $\mathbb{E} \xi_k = \binom{k-1}{r-1} c^r \rightarrow \infty$ as $k \rightarrow \infty$. Hence, there is some K_0 such that $\mathbb{E} \xi_{K_0} > 1$. Since ξ_k stochastically dominates ξ_{K_0} for $k \geq K_0$, it follows that if the process reaches K_0 without stopping, the continuation dominates (up to a change of time) a Galton–Watson branching process with offspring distribution ξ_{K_0} , which is supercritical and thus has a positive probability of living forever. Hence, $\mathbb{P}(\tilde{T} = \infty) > 0$. \square

PROOF OF THEOREM 5.8. It suffices to consider $a = r$. Thus assume $a = r$, and consider the vertices activated in the first generation, that is, at time $t = r$. There are $S(r) \in \text{Bin}(n - r, p^r)$ such vertices. [Note that, see (2.11), $\pi(r) = \mathbb{P}(\text{Bin}(r, p) = r) = p^r$.] Consequently, $\mathbb{E} S(r) = (n - r)p^r \rightarrow \infty$. Let $\omega(n) = \mathbb{E} S(r)/2$, so $\omega(n) \rightarrow \infty$. It follows from Chebyshev’s inequality (or Chernoff’s) that w.h.p. $S(r) > \omega(n)$. Hence, w.h.p. for all $t \in [r, \omega(n)]$, $A(t) \geq A(r) > S(r) > \omega(n) \geq t$. Together with the trivial $A(t) \geq a = r > t$ for $t < r$ and Lemma 11.4, this shows that w.h.p. $A(t) > t$ for all $t \leq n - 1$, and thus $A^* = T = n$. \square

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