

Notes on HW1  
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**Hw1/Ex2**

This problem is not well-defined. Let us examine a more general situation to see what we can do to provide a remedy.

The urn contains a random subset of distinct colours. Not only the elements of the set are random, but also its size.

Say there are  $n$  possible distinct colours. Call them  $1, 2, \dots, n$ . Define  $\xi$  to be one of the  $n!$  permutations of these colours.

Now let  $N$  be a random integer with values in  $\{1, \dots, n\}$ . We are TOLD that  $P(N = i) = 1/n$ , for  $i = 1, \dots, n$ , i.e. we have complete ignorance of what  $N$  is.

What we do is this: we define

$$Y = (\xi_1, \dots, \xi_N),$$

so, for example, if  $N = 2$ , then  $Y = (\xi_1, \xi_2)$ .

Thus, to define  $Y$ , we need two random variables: one is  $\xi$ , the other is  $N$ . To find the law of  $Y$  we need the JOINT law of  $\xi$  and  $N$ . The problem only tells us the law of  $N$ . Now, it is REASONABLE to assume (but this is an ASSUMPTION!) that

$$P(\xi_1 = \alpha_1, \dots, \xi_i = \alpha_i | N = i) = \frac{1}{n} \frac{1}{n-1} \cdots \frac{1}{n-i+1},$$

where  $\alpha_1, \dots, \alpha_i$  are distinct elements of  $\{1, \dots, n\}$ . This is tantamount to assuming that  $\xi$  and  $N$  are independent. (Show this!)

Say colour 1 is the red colour. We are asked to compute a conditional probability of the form

$$P(N = 1 | 1 \in Y),$$

where “ $1 \in Y$ ” is an abbreviation for the statement “ $\xi_1 = 1$  or  $\xi_2 = 1$  or  $\dots$  or  $\xi_N = 1$ ”. Having defined the joint law of  $N$  and  $\xi$  properly, we have no fear to make a mistake like Alice. So here we go:

$$P(N = 1 | 1 \in Y) = \frac{P(N = 1, 1 \in Y)}{P(1 \in Y)} = \frac{P(N = 1, 1 \in Y)}{\sum_{i=1}^n P(N = i, 1 \in Y)}$$

But

$$\begin{aligned}
 P(N = i, 1 \in Y) &= P(1 \in Y | N = i)P(N = i) \\
 &= P(\xi_1 = 1 \text{ or } \cdots \text{ or } \xi_i = 1 | N = i) \frac{1}{n} \\
 &= P(\xi_1 = 1 \text{ or } \cdots \text{ or } \xi_i = 1) \frac{1}{n} \quad [\text{by independence between } N \text{ and } \xi] \\
 &= \frac{i}{n} \frac{1}{n} \quad [\text{because the events between the or's are pairwise disjoint}]
 \end{aligned}$$

Since  $1 + 2 + \dots + n = n(n-1)/2$ , we have that  $\sum_{i=1}^n P(N = i, 1 \in Y) = n(n-1)/2n^2$ . So

$$P(N = 1 | 1 \in Y) = \frac{1/n^2}{n(n-1)/2n^2} = \frac{2}{n(n-1)}.$$

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If we have a choice with replacement, we need to change our probability space. Let  $1, \dots, n$  represent the different colours. We let  $\xi_1, \dots, \xi_n$  be i.i.d. colours, i.e., each  $\xi_i$  is uniformly distributed in  $\{1, \dots, n\}$  and the  $\xi_1, \dots, \xi_n$  are mutually independent. (This assumption is equivalent to the assumption that the random vector  $(\xi_1, \dots, \xi_n)$  is uniformly distributed in the set  $\{1, \dots, n\}^n$ . (WHY?))

Let  $N$  be a random variable with values in  $\{1, \dots, N\}$ . We are told that  $P(N = i) = 1/n$ , for all  $i$ .

As before, we look at the random variable  $Y = (\xi_1, \dots, \xi_N)$ , and, as before, we are going to make the ASSUMPTION that  $\xi$  and  $N$  are independent. Hence

$$P(1 \in Y | N = i) = P(\xi_1 = 1 \text{ or } \cdots \xi_i = 1),$$

but now the events are not disjoint:  $\{\xi_1 = \xi_2 = 1\} \neq \emptyset$ . However, by independence

$$P(\xi_1 \neq 1, \dots, \xi_i \neq 1) = ((n-1)/n)^i.$$

So

$$P(1 \in Y | N = i) = 1 - (1 - (1/n))^i.$$

And so

$$\begin{aligned}
 P(N = 1 | 1 \in Y) &= \frac{P(1 \in Y | N = 1)P(N = 1)}{\sum_{i=1}^n P(1 \in Y | N = i)P(N = i)} \\
 &= \frac{1/n}{n - \sum_{i=1}^n (1 - (1/n))^i} = \frac{1/n}{n - n(1 - (1/n))^{n+1}} = \frac{1}{n(1 - (1/n))^{n+1}}
 \end{aligned}$$

For large  $n$  this is about  $1/en \approx 0.37/n$ , which is much larger than the previous answer (about  $2/n^2$ ). This is reasonable!

### **HW1/Ex3**

One way to model the situation “there are  $n$  balls, some black, some white, but the proportion of the black balls is  $\beta$ ” (for given  $0 < \beta < 1$ ) is as follows.

We may take the point of view that each of the balls have been painted black with probability  $\beta$ , independently of one another. This means that the total number of black balls in the urn equals the total number of successes (success=black) in  $n$  independent coin tosses, whence the number of black balls has law  $\text{BINOMIAL}(n, \beta)$ .