Explanatory Notes 2 for Bayesian Inference **Rejection sampling** Takis Konstantopoulos, Spring 2009

The proof of correctness of the rejection sampling algorithm (SZ notes, page 19) requires further elaboration.

Let us recall the algorithm. We wish to simulate a random variable with given density f. We assume that we know how to simulate a random variable from another density g which has the property that there exists a constant M > 0 such that

$$f(x) \le Mg(x)$$
, for all x (in the support of f). (1)

Define

$$p(x) = \frac{f(x)}{Mg(x)}.$$

Let Z_1, Z_2, \ldots be i.i.d. random variables with common density g. Let ξ_1, ξ_2, \ldots be i.i.d. coin tosses, conditionally on the Z_1, Z_2, \ldots In other words,

$$P(\xi_n = 1 | Z_1, Z_2, \ldots) = P(\xi_1 = 1 | Z_n) = p(Z_n)$$

$$P(\xi_n = 0 | Z_1, Z_2, \ldots) = P(\xi_0 = 1 | Z_n) = 1 - p(Z_n).$$

Let

$$N = \min\{n \ge 1 : \xi_n = 1\}.$$

The claim is:

Claim 1. Z_N has density f.

Proof. First observe that

$$P(\xi_n = 1) = E[P(\xi_n = 1 | Z_n)] = E[p(Z_n)]$$

= $\int p(z)g(z)dz = \int \frac{1}{M}f(z)dz = \frac{1}{M}$,

where the integral is taken over the support of g which, by assumption (1),

includes the support of f. We then have

$$\begin{split} P(Z_N \in dy) &= \sum_{n=1}^{\infty} P(Z_N \in dy, N = n) \\ &= \sum_{n=1}^{\infty} P(Z_n \in dy, N = n) \\ &= \sum_{n=1}^{\infty} P(Z_n \in dy, \xi_1 = \dots = \xi_{n-1} = 0, \xi_n = 1) \\ &= \sum_{n=1}^{\infty} P(Z_n \in dy) P(\xi_1 = \dots = \xi_{n-1} = 0, \xi_n = 1 | Z_n = y) \\ &= \sum_{n=1}^{\infty} g(y) dy \ P(\xi_1 = 0) \dots P(\xi_{n-1} = 0) \ P(\xi_n = 1 | Z_n = y) \\ &= \sum_{n=1}^{\infty} g(y) dy \ (1 - M^{-1})^{n-1} \ p(y) \\ &= \sum_{n=1}^{\infty} f(y) dy \ \frac{1}{M} \ (1 - M^{-1})^{n-1} \\ &= f(y) dy \frac{1}{M} \sum_{n=1}^{\infty} (1 - M^{-1})^{n-1} = f(y) dy \frac{1}{M} M = f(y) dy, \end{split}$$

as required.

Apology. If you don't like the notation $P(\dots \in dy)$ then replace it by $P(\dots \in A)$ for some set A and rewrite the calculation.